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# Multi-modal analysis of aerosol robotic network size distributions for remote sensing applications: dominant aerosol type cases

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## Abstract

To date, size distributions obtained from the aerosol robotic network have been fit with bi-lognormals defined by six secondary microphysical parameters: the volume concentration, effective radius, and the variance of fine and coarse particle modes. However, since the total integrated volume concentration is easily calculated and can be used as an accurate constraint, the problem of fitting the size distribution can be reduced to that of deducing a single free parameter – the mode separation point. We present a method for determining the mode separation point for equivalent-volume bi-lognormal distributions based on optimisation of the root mean squared error and the coefficient of determination. The extracted secondary parameters are compared with those provided by AERONET's Level 2.0 Version 2 inversion algorithm for a set of benchmark dominant aerosol types including: desert dust, biomass burning aerosol, urban sulphate and sea salt. The total volume concentration constraint is then also lifted by performing multi-modal fits to the size distribution using nested Gaussian mixture models and a method is presented for automating the selection of the optimal number of modes using a stopping condition based on Fisher statistics and via the application of statistical hypothesis testing. It is found that the method for optimizing the location of the mode separation point is independent of the shape of the AVSD, does not require the existence of a local minimum in the size interval  $0.439 \mu\text{m} \leq r \leq 0.992 \mu\text{m}$ , and shows some potential for optimizing the bi-lognormal fitting procedure used by AERONET particularly in the case of desert dust aerosol. The AVSD of impure marine aerosol is found to require 3 modes. In this particular case, bi-lognormals fail to recover key features of the AVSD. Fitting the AVSD more generally with multi-modal models allows automatic detection of a statistically-significant number of aerosol modes, is applicable to a very diverse range of aerosol types, and gives access to the secondary microphysical parameters of additional modes currently not available from bi-lognormal fitting methods.

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# 1 Introduction

## Highlights

- A method for optimizing bi-lognormal fits to the size distribution with a single free parameter – the mode separation point
- Sensitivity analysis of the dependence of secondary microphysical parameters on the mode separation point and on aerosol optical depth
- A method for multi-modal analysis of the size distribution using Gaussian mixture models and access to the microphysical parameters of higher modes
- A test of the feasibility of the methods for fitting size distributions of dominant aerosol types of diverse morphology

Satellite retrievals of aerosol optical depth (AOD) and related parameters typically require the use of prescribed models of aerosol size and composition. In particular, the aerosol volume size distribution (AVSD) and the spectral complex refractive index are needed to compute properties such as the scattering phase function, the single scattering albedo and the extinction coefficient, which are in turn used to calculate quantities such as the total AOD from the columnar abundance. In general, the information content of measurements from current satellite radiometers is insufficient to unambiguously retrieve all these parameters particularly when the (spectral and directional) behavior of the surface reflectance is unknown (Hasekamp and Landgraf, 2007). For this reason, aerosol retrieval algorithms employed by most of these sensors are required to make assumptions about microphysical properties. The consequence is that these assumptions then contribute to differences in retrieved AOD – even in the idealized case of a black (non-reflecting) surface (Kokhanovsky et al., 2010).

So, while the ability of satellite retrieval algorithms to represent the radiative-behaviour of real aerosols is still in question (most recently raised in the context of

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pure marine aerosol models by Sayer et al., 2012), satellite retrievals are usually validated (e.g. Remer et al., 2005) against co-located and synchronous retrievals provided by ground-based sun photometer and sky radiometer systems like those in the aerosol robotic network (AERONET). The main reason for this is that remote sensing of the AVSD in particular is exceedingly difficult and no one sensor system is capable yet of providing totally unambiguous information (King et al., 2009). Moreover, the aerosol model inter-comparison initiative – AeroCom has carried out analysis of aerosol simulations from various global chemical transport and climate models and found that, even on the scale of yearly averages, aerosol life cycles and particle sizes span a large range of values (Textor et al., 2006, 2007) and the total number of aerosol modes, mode composition and control parameters vary considerably both between models and their final simulation results (Zhang et al., 2010). As a result, there is heavy reliance on the AVSDs provided by AERONET.

The advanced mathematical inversion algorithm developed to provide AERONET retrievals (Dubovik and King, 2000) from direct (sun) and diffuse (sky) measurements of radiation, returns aerosol optical parameters such as the spectral AOD, single scattering albedo and phase function, as well as important microphysical parameters like the AVSD, the complex refractive index, and the percentage of spherical particles (see Dubovik et al., 2002, 2006). Since the AVSD plays a pivotal role in the relation of the radiation field to the microphysics of aerosol particles (Hansen and Travis, 1974) as well as for the determination of aerosol type and composition (Dubovik et al., 2011), this paper focuses on the development of a method for assessing whether or not the AVSD for a couple of characteristic cases can be fit using multiple aerosol modes with a procedure that can be automated.

The AERONET Level 2.0 Version 2 inversion code inverts sky radiances simultaneously at wavelengths available in the CIMEL instrument (most frequently at 440, 675, 870 and 1020 nm) for the complete solar almucantar scenario or principal plane scenario together with measurements of AOD at the operational wavelengths. In particular, the algorithm returns the AVSD  $dV(r)/d\ln r$  in 22 equidistant logarithmic radial size

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bins spanning the range of particle radii  $0.05 \leq r \leq 15 \mu\text{m}$ , normalized to the value of the total volume concentration of aerosol in  $\mu\text{m}^3 \mu\text{m}^{-2}$ . The AERONET inversion code approximates the AVSD using trapezium rule integration (Dubovik and King, 2000) and, while the option of allowing the use of lognormal-shaped bins was included in Dubovik et al. (2006), it has only recently been found that sufficiently accurate modeling of POLDER/PARASOL observations requires such an optimization of the shape of each radial size bin (Dubovik et al., 2011). In particular, lognormal-shaped bins were found to provide notable improvements over the trapezoidal approximation in terms of smoothness and suggested some advantage in modeling the AVSD as a superposition of  $n$ -lognormals with the modal volume concentrations  $V_i$ , geometric mean radii  $r_i$  and standard deviations  $\sigma_i$  as fixed parameters (Dubovik et al., 2011),

$$\frac{dV(r)}{d \ln r} = \sum_{i=1}^n \frac{V_i}{\sigma_i \sqrt{2\pi}} e^{-\frac{1}{2} \left( \frac{\ln r - \ln r_i}{\sigma_i} \right)^2} \quad (1)$$

This finding is the main motivation for this paper. For several decades now, the AVSD of tropospheric aerosols has been known to contain several distinct modes, each most commonly being modeled by a lognormal function (Whitby, 1978; Remer et al., 1997, 1998; Remer and Kaufman, 1998; O'Neill et al., 2000). The statistical properties of the lognormal and bi-lognormal distribution are well-known (O'Neill et al., 2000) and are applied in this paper to test the feasibility of modeling the AVSD of distinct aerosol cases with super-positions of several ( $n \geq 2$ ) lognormals. The multi-modal method presented here, it is hoped, will add a new layer of detail to existing studies without the need for too much additional mathematical complexity. Furthermore, since many available radiative-transfer codes are now able to take as input lognormal distribution parameters (Sayer et al., 2012), the results of this paper can be readily applied and implemented. In Appendix A, the equations used to calculate secondary microphysical parameters such as  $V_i$ ,  $r_i$  and  $\sigma_i$  are presented.

The paper is organized as follows. The approach adopted for aerosol typing and the selection of sites impacted by dominant aerosol types is presented in Sect. 2.

Section 3 then briefly outlines two new methods for optimizing bi-lognormal fits and for fitting the AVSD with multiple modes. In Sect. 4, the results of applying the two new methods to a cohort of AVSDs representative of different aerosol types are presented and compared with AERONET, and the major impacts and feasibility of these new approaches are noted and analyzed. Finally, we conclude in Sect. 5 with a summary of our findings and an assessment of the potential offered by these new methods for analyzing AVSDs provided by AERONET or other remote sensing instruments.

## 2 Data selection

In this paper we apply new methodologies (developed in Sect. 3) to a set of dominant aerosol type AVSDs. While portals like NASA's AERONET Data Synergy Tool ([http://aeronet.gsfc.nasa.gov/cgi-bin/bamgommas\\_interactive](http://aeronet.gsfc.nasa.gov/cgi-bin/bamgommas_interactive)) and the Multi-sensor Aerosol Products Sampling System (MAPSS: <http://giovanni.gsfc.nasa.gov/mapss>) provide a framework for multi-sensor aerosol validation, inter-comparison, and joint analysis, a search for dominant aerosol type cases and high load aerosol events must still be done manually or with reference to field campaigns published in the literature. Here, we describe the approach we adopted to isolate candidate AERONET sites as well as those days which are most dominated by desert dust, biomass burning products, urban sulphate, and marine sea salt. We will see below that "dominant" sea salt is the most problematic case for bi-lognormal fitting methods owing to the fact that the data is drawn from an island site where the marine aerosol is mixed to a high degree (in the proportion 60% : 40%) with other aerosol species.

The Georgia Institute of Technology–Goddard Global Ozone Chemistry Aerosol Radiation and Transport (GOCART) model (Chin et al., 2000, 2002 and Ginoux et al., 2001) used by NASA's GEOS-5, simulates the AOD for major types of tropospheric aerosols. In particular, it provides 3 hourly measurements of the total extinction AOD as well as the contribution to the total extinction AOD of sulphate (SU), black carbon (BC), organic carbon (OC), desert (mineral) dust (DU) and sea salt (SS). It therefore

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provides a model-driven aerosol classification. This is complementary to the way the AERONET's *Spectral Deconvolution Algorithm Product* (O'Neill et al., 2003) uses a more generalized set of microphysical assumptions to estimate the contributions of fine and coarse particles to the AOD at visible wavelengths. GOCART data spanning the years 2001–2005 (inclusive) is obtainable from the AERONET data synergy portal and was downloaded for the first 155 AERONET sites (75 % of all *Level 2.0 Version 2 Inversion Product* records  $N$ ) ranked by the number of daily-averaged data available. Since GOCART provides eight 3 hourly measurements per day, these were averaged to produce daily-averages. The ratio of the contribution of individual aerosol types to the total extinction AOD was then calculated (as percentages) and used to sort the ranked sites by dominant (“nearly-pure”) aerosol type. This approach provides a simple and straight-forward method for site selection. During this process, an additional column was added to the data provided by GOCART – so as to monitor a combination of aerosol: the combined percentage of organic and black carbon (OC + BC). The reason for this is that, although SU accompanies the burning of biomass products, the combination OC + BC was found to better distinguish biomass burning sites from urban sites (which can also have high levels of SU). While no site of course has 100 % “pure” aerosol of a single type, this approach enables one to rank sites by *dominant* aerosol type. For example, for the biomass burning products OC + BC, Mongu was selected since: (i) it has the longest AERONET data record ( $N = 1573$  days) during the period 2001–2005, and (ii) it has a very high OC + BC presence (71.3 %), second only to Alta Floresta (77.78 %). Analogous criteria were used to select sites dominated by dust, urban-industrial SU and sea salt aerosol. As a result, the following sites that are representative of the dominant aerosol types were selected:

- Urban-industrial pollution: GSFC-Washington, US (76.840° W, 38.992° N, elevation = 87 m)
- Biomass burning: Mongu, Zambia (23.151° E, 15.254° S, elevation = 1107 m)
- Dust: Banizoumbou, Niger (2.665° E, 13.541° N, elevation = 250 m)





re-calculated the integral. This was then repeated until the integral converged. The rationale for interpolating the AVSD is twofold:

1. in order to decrease the radial step size and hence improve the validity of the sensitivity analysis applied to the position of the mode separation point  $r_s$  described in Sect. 3.2
2. in order to avoid spikes in the errors propagated using the Gaussian mixture method (GMM) described in Sect. 3.3 that are caused by jaggedness resulting from straight line connections across 22 discrete bins.

For the case of peak marine (sea salt) AVSD at Lanai, Fig. 1 shows how successive doubling of the number of interpolation points leads to a suitable smoothing of the AVSD without introducing spurious features.

Regarding the accuracy of the AERONET retrieval (the grey band in Fig. 1), the overall uncertainty in AOD data (under cloud-free conditions) is  $\pm 0.01$  for wavelengths longer than 440 nm, and  $\pm 0.02$  for shorter wavelengths (Dubovik et al., 2000), and the error in aerosol AVSD is estimated to be  $< 10\%$  for particle radii between 0.1  $\mu\text{m}$  and 7  $\mu\text{m}$  (see Dubovik et al., 2000). While this is true near the maxima of the distribution, the errors can be as large as 35% for the lowest AVSD values in this particle range (Dubovik et al., 2002). Furthermore, at the edges of the AVSD ( $r < 0.1 \mu\text{m}$  and  $r > 7 \mu\text{m}$ ) the accuracy of the retrieval drops significantly because of the low sensitivity of aerosol scattering at 440, 670, 870 and 1020 nm to particles of these sizes (Dubovik et al., 2002). Correspondingly, the retrieval errors rise sharply to 80–100% at the edges but do not significantly affect the derivation of the secondary microphysical parameters because typically the value of the AVSD is very low there (Dubovik et al., 2002). To provide a conservative uncertainty context for the results presented here, AERONET AVSDs are overlaid with an error band that is 10% at peaks in the interval  $1 \leq r \leq 7 \mu\text{m}$ , 35% at local minima in this range, and 100% in edge regions when  $r < 0.1 \mu\text{m}$  and  $r > 7 \mu\text{m}$ . Between these thresholds, the error is interpolated on an equidistant, logarithmically-spaced grid to ensure a smooth transition. Note that the

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overall trend with increasing number of points is as follows:  $V$ ,  $V_f$  and  $V_c$  were all found to converge with increasing  $N$  but did not converge asymptotically to the values given by the AERONET Level 2.0 Version 2 inversion retrieval. In particular when comparing with the quoted AERONET values, the interpolated value of  $V$  is slightly lower (by 1.19 %),  $V_f$  is moderately lower (16.20 %) and  $V_c$  is slightly higher (1.81 %) – but all well within the prescribed AERONET uncertainty band.

In Fig. 3 (in Sect. 3.2), the AERONET-retrieved AVSD together with the associated bi-lognormal fit reconstructed from the AERONET-quoted values of the secondary microphysical parameters is shown. Note that 3 goodness of fit statistics are also provided: the mean of the residuals  $b$ , the standard error of the fit  $s$ , and the degrees of freedom-adjusted  $R^2$  which measure the difference between the interpolated AERONET AVSD and the bi-lognormal re-constructed using the secondary parameters in Eq. (1) for  $n = 2$  modes. All data-model comparisons in this work are accompanied by these measures of bias, location and spread ( $b$ ,  $s$  and  $R^2$  respectively). Figure 3 clearly shows that in the case of maritime (sea salt) aerosol, the fit to the fine mode is good but the fit to the coarse region is inappropriate – both from the perspective of fitting the “double hump” with a single broader peak, and from the perspective of the overall amplitude in this region. Given that the double hump occurs in the coarse region, it is possible that it is caused by the existence of a mixture of 2 coarse populations arising perhaps from: (a) 2 different aerosol types, (b) fresh aerosol with an aged component, or (c) some combination of these. In any case, it appears that the mode separation point  $r_s$  for this AVSD leads to a bi-lognormal fit that is good for the fine mode but poor for the coarse mode. The requirement for the total integrated volume concentration to remain constant means that, while it is feasible that the double-hump coarse mode region can be better fit by changing the mode separation point, the fit to the fine mode peak will necessarily have to worsen. This is suggestive of a need for handling the problem in a different way. With this in mind, in the next section, we present a method for unambiguously and automatically calculating the optimal location of  $r_s$  by optimizing the statistical measures  $s$  and  $R^2$ .

### 3.2 Optimised Equivalent-Volume (OEV) bi-lognormal fitting

Since the AERONET code uses  $r_s$  to separate fine and coarse modes and then obtains the spectral AOD extinction and absorption and asymmetry factors for these modes, the location of  $r_s$  is central. Furthermore, the fact that such spectral parameters are being used to validate satellite retrievals of fine and coarse modes means that the role played by  $r_s$  is becoming more prominent. It is therefore important to test the assumption that bi-lognormals provide the best fit to the AVSD and, if so, to assess the impact of uncertainty in the value of  $r_s$  on derived microphysical parameters. For this purpose, we stepped  $r_s$  through the set of 2200 (interpolated) equidistant logarithmically-spaced radial bins (excluding the end points). Then, using the interpolated volume concentration as a constraint, in each step, a bi-lognormal was fit to the AVSD and the secondary (derived) microphysical parameters:  $V_f$ ,  $V_c$ ,  $r_f$ ,  $r_c$ ,  $\sigma_f$  and  $\sigma_c$ , and goodness of fit measures:  $s$  and  $R^2$ , were calculated and tabulated. While the 2198 fits obtained (minus the end-points) fill a continuum, we show in Fig. 2 the result of applying this procedure to the AERONET AVSD radial bins in the range 0.1–7  $\mu\text{m}$  (the “10 % error range”).

While Fig. 2 shows no discernable bias  $b = 0.000$  (to 3 decimal places) and a small and stable standard error  $s \approx 0.001 \mu\text{m}$ ,  $R^2$  is much more sensitive to changes in  $r_s$  and reveals a peak value of 0.893 at  $r_s = 0.286 \mu\text{m}$ . This suggests a method for automating the detection of the optimal value of  $r_s$  related to  $\max(R^2)$ . Despite appearing constant in the legend of Fig. 2, a unique trough was found to exist in the curve of  $s$  – suggesting an optimal separation related to  $\min(s)$  when  $r_s = 0.290 \mu\text{m}$  ( $\approx 9\%$  smaller than that obtained with  $R^2$ ). These two estimates of  $r_s$  are significantly lower than the mode separation point  $r_s = 0.439 \mu\text{m}$  quoted by AERONET at this site on this day. Since we could not find a clear reason for favouring the  $\max(R^2)$  method over the  $\min(s)$  method, we decided to define the optimal  $r_s$  as the mean of the values obtained from the  $\min(s)$  method and the  $\max(R^2)$  method, i.e.  $r_s = (0.290 \mu\text{m} + 0.315 \mu\text{m})/2 = 0.303 \mu\text{m}$ . Figure 3 compares the results of the standard AERONET bi-lognormal fit with that obtained by optimizing the separation point using the method described above.

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The OEV method developed above, based on optimisation of statistical measures of goodness of fit, obtains a new mode separation point that marginally improves the bi-lognormal fit to the AVSD ( $R^2 = 0.894$  as compared with  $R^2 = 0.885$ ). However, while the improvement is only of the order of 1 % in terms of  $R^2$ , there is a significant *qualitative* improvement. Visually, the fit to the fine mode with the new OEV method is much better, particularly on the smaller radius side of this marine aerosol AVSD. Its peak, while within the AERONET error in this region, appears to be slightly over-estimated in amplitude but is well-located. The fit to the coarse mode is also better in terms of the peak amplitude. It is also nearer to the raw data on the rising edge of the coarse mode peak region at smaller radii. However, neither the AERONET bi-lognormal fit nor the new OEV fit are able to fit the double-peak in the coarse mode region. In Sect. 3.3, we present a second fitting method based on Gaussian mixture models to investigate whether or not the inclusion of additional modes can account for such features in the AVSD.

With regard to the dependence of the secondary microphysical parameters on  $r_s$ , we also performed a sensitivity analysis over the range  $0.1 \leq r_s \leq 7 \mu\text{m}$  for the AVSD at Lanai, Hawaii on the 21 January 2002 interpolated with 2200 points as described in Sect. S1.1 of the Supplement. In the range of mode separation points  $0.286 \mu\text{m} \leq r_s \leq 0.567 \mu\text{m}$  (i.e. up to the edge of the coarse region  $\approx 0.6 \mu\text{m}$ ) the magnitude of the relative errors of the geometric mean radii  $r_f$  and  $r_c$ , geometric standard deviations  $\sigma_f$  and  $\sigma_c$  and mode volumes  $V_f$  and  $V_c$  were found not to exceed 30 % as can be seen in Table S1. However, steep gradients were observed outside this range. For example, in the range of mode separation points  $0.439 \leq r_s \leq 0.992 \mu\text{m}$  used by the AERONET inversion code, the fine mode parameters  $r_f$ ,  $\sigma_f$  and  $V_f$  reached large negative relative errors especially at the higher radius end where  $r \approx 1 \mu\text{m}$  (see Table S1). This suggests that the AERONET bi-lognormal fit is strongly under-predicting their values. The sensitivity analysis shows that apparently small differences in the value of the deduced separation point  $r_s$ , can be seen to translate into large differences in the deduction of secondary microphysical parameters – and hence the shape of the reconstructed







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answer at this point is: how can detection of the 3-modal “optimum” be automated? For example, a detection algorithm based on seeking the maximum value of  $R^2$  is likely to flag up the 6-modal fit as the optimal fit to the above AVSD. In particular, care should be taken to ensure that additional modes are *physical* and not just artefacts of the fitting procedure. What is needed therefore is a stopping condition to find the optimal mixture. One way to do this is to define a statistic and then to perform a hypothesis test to assess whether or not adding an extra mode leads to a statistically-significant improvement in the fit. Here, we adopt the protocol outlined by Harel (2009) and work with the square root of the degrees of freedom-adjusted  $R^2$  as a proxy for the Pearson product-moment correlation coefficient  $\rho$ . This is based on the assumption that the nesting procedure (i.e. adding more modes and therefore model parameters) does not cause much divergence to occur between the coefficient of determination  $R_d^2$  and the degrees of freedom-adjusted  $R^2$  (defined in Eqs. B4 and B6 in Appendix B). This was verified for all of the dominant aerosol type cases studied here and the results of the calculations are presented in Sect. S2 of the Supplement. The percentage relative error (RE) between  $R^2$  and  $R_d^2$  was found to be very small for GMMs containing 1–6 modes – reaching a maximum value of RE = 0.060 %. Propagating this error into the square root of  $R^2$  (the proxy for Pearson’s  $\rho$ ), we found that this had an effect only on the 4th decimal place and did not impinge on the results of the hypothesis testing procedure (see below) at the 95 % level. Having justified the use of the square root of  $R^2$  as a proxy for Pearson’s  $\rho$ , we then proceeded to construct confidence intervals on  $\rho$  using the Fisher transform (Fisher, 1921):

$$F(\rho) = \frac{1}{2} \ln \frac{1 + \rho}{1 - \rho} \quad (6)$$

where  $F(\rho)$  is a transformed value of  $\rho$  that follows approximately a normal distribution with standard error  $SE = 1/\sqrt{N-3}$  for a sample of  $N$  points (Fisher, 1921). If we note that the 0.975 quantile of the normal distribution has a  $z$ -score of 1.96, then the upper and lower 95 % confidence limits are simply:  $F(\rho) \pm 1.96/\sqrt{N-3}$ . We calculated these

limits for the value of  $\rho$  obtained for each GMM (1–6 modes). Two values of  $F(\rho)$  (and hence  $R^2$ ) show a significant statistical difference when the lower confidence limit of the larger  $F(\rho)$  value does not overlap the upper confidence limit of the smaller  $F(\rho)$  value. In the event of an overlap, the Welch t-statistic for unequal variances (Welch, 1947),

$$t = \left| \frac{F(\rho_1) - F(\rho_2)}{\sqrt{\frac{1}{N_1-3} + \frac{1}{N_2-3}}} \right| \quad (7)$$

reports a significant statistical difference when  $t > 1.96$ . In this way, a test was performed as modes were successively added to the Gaussian mixture. The optimal GMM fit occurs when adding a new mode does not lead to a significant statistical difference in  $F(\rho)$  (or  $t$  in the case of over-lapping values) at the 95 % level of confidence. The calculation is presented in Table 2 for the automatic identification of the optimal (3-mode) mixture pertinent to the case of maximum marine sea salt illustrating this section.

Figure 5 shows the resulting 3-modal GMM fit ( $b = 0.00$ ,  $s = 0.000$ ,  $R^2 = 0.998$ ) to the AVSD. Comparing with the AERONET bi-lognormal fit ( $b = 0.00$ ,  $s = 0.001$ ,  $R^2 = 0.885$ ) and the OEV fit ( $b = 0.00$ ,  $s = 0.001$ ,  $R^2 = 0.894$ ) shown in Fig. 3, the result of fitting with the GMM is clearly both quantitatively and qualitatively better.

## 4 Results

The case of fitting the AVSD of dominant marine aerosol at Lanai with a bi-lognormal shows that things are not so simple but that fitting problems could be overcome with the GMM method. This motivates a study of other geo-locations where the aerosol composition is also clearly defined so as to assess under what conditions, bi-lognormal fits are appropriate or not. With this in mind, in Table S3 in the Supplement accompanying this paper, we collect together the results of fitting the AVSD for each of the 4 dominant aerosol type cases with the methods introduced in Sects. 3.2 and 3.3.

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OEV method is  $r_s = 0.858 \mu\text{m}$ . With this mode separation point, Table S1 shows that the size of the relative error (using AERONET values as a reference) is very high for the fine volume ( $-115.9\%$ ) and the fine radius ( $-87.0\%$ ) – i.e. the value quoted by AERONET ( $r_f = 0.195 \mu\text{m}$ ) is  $87.0\%$  lower than that found using the optimized fit bi-lognormal fit ( $r_f = 0.365 \mu\text{m}$ ). In contrast, in the case of peak biomass burning at Mongu, the tabulated entry closest to the optimal OEV value is  $r_s = 0.567 \mu\text{m}$ . With this mode separation point, the relative errors of the microphysical parameters are in good agreement with those derived by AERONET and are in the narrow range:  $-3.3\%$  (for  $\sigma_c$ ) to  $+2.6\%$  (for  $\sigma_f$ ). The same is true for peak urban SU at GFSC-Washington, where the tabulated entry closest to the OEV optimum is  $r_s = 0.528 \mu\text{m}$ . The similarity of this value to that quoted by AERONET for this day ( $r_s = 0.756 \mu\text{m}$ ) also translates into small relative errors, spanning the narrow range:  $-9.0\%$  (for  $\sigma_c$ ) to  $+5.1\%$  (for  $r_c$ ). The situation takes a turn for the worse in the case of peak marine aerosol. While both the AERONET and OEV methods point to similar mode separation points ( $r_s = 0.885 \mu\text{m}$  and  $r_s = 0.894 \mu\text{m}$  respectively), referencing the closest tabulated entry of  $r_s = 0.885 \mu\text{m}$  in Table S1, shows that the fine mode parameters:  $r_f$ ,  $\sigma_f$  and  $V_f$  are strong under-estimated ( $-53.69\%$  to  $-68.73\%$ ) and the coarse mode parameters  $r_c$ ,  $\sigma_c$  and  $V_c$  are being over-estimated with a magnitude of  $\approx 15\%$ . Dust and marine aerosol microphysical parameters are highly sensitive to the location of the mode separation point  $r_s$ .

Turning to the results of the GMM method, the fits to the AVSD for each of the 4 dominant aerosol type cases show a significant improvement over both those obtained by AERONET and using the OEV method. In Fig. 6 below, the interpolated AERONET AVSD retrieval is overlaid with the re-constructed AERONET bi-lognormal fit, the OEV bi-lognormal fit and the GMM optimal fit obtained by hypothesis testing.

Figure 6a shows that the GMM method flags up only peak dust aerosol as an  $n = 2$  bi-lognormal. In contrast, it fits peak biomass burning, urban SU and marine (sea salt) AVSDs in Fig. 6b–d with  $n = 3$  modes. While the qualitative appropriateness of these detections may be uncertain (see below) the results are statistically-significant in the





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simultaneously a double-peak in the coarse mode region could not be attributed to miscalculation of the position of the mode separation point – suggesting strongly that 2 modes are not sufficient to fit this type of aerosol. Overall, the OEV method produced an improvement over the AERONET method in all 4 dominant aerosol type cases studied. It also provided a way to make explicit the sensitivity of secondary microphysical parameters to the location of the mode separation point  $r_s$ . It was found that displacing  $r_s$  by even a single bin could lead to relative errors of the order of several tens of percent on the derived parameters. While this is precisely what gives the OEV method its capacity for improving bi-lognormal fits, it highlights that some care should be taken when using the values of microphysical parameters provided by AERONET – particularly in the case of dust and marine aerosol.

In relation to the GMM method, in all cases, it was possible to fit dominant aerosol type cases with a very high goodness of fit which is almost indistinguishable from the raw AERONET AVSD. For peak dust, the best fit was found to be bi-modal whereas for peak biomass burning (BC + OC), urban SU and marine (sea salt), the best fit was tri-modal. The use of iterated nonlinear least squares to obtain the microphysical coefficients was very efficient – although it was necessary to interpolate the AVSD with a 100-fold increase in the number of points (from 22 bins to 2200 bins) so as to avoid numerical instability (i.e. so that the propagated errors of the fit were stable at the 95 % level of confidence). Using the square root of  $R^2$  as a proxy for the Pearson product-moment correlation coefficient  $\rho$  and applying the Fisher  $z$  transform allowed us to perform a test for a statistically-significant improvement in the fit with the addition of each additional mode. The null hypothesis at the 95 % confidence level provided a stopping condition that enabled automatic identification of the optimal number of modes in the mixture. The GMM method, in addition to providing a better overall fit, provides important details concerning the amplitude, location and width of each mode contributing to the mixture and hence allows for determination of secondary microphysical parameters in the case of  $n > 2$  modes – something not currently possible with methods based on a single mode separation point. As such, the multi-modal content of the fit to the





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With regard to estimation of the accuracy of the methods developed here, the mean bias  $b$ , standard error of the fit  $s$ , and the degrees of freedom-adjusted coefficient of determination  $R^2$ , were found to be very useful statistics for assessing the goodness of fit of the OEV method and the GMM method with respect to AERONET AVSDs. Consideration of the rate of change of  $s$  and  $R^2$  with respect to changes in  $r_s$  was what gave the OEV method its capacity to automate the detection of the optimal separation point. In the case of the GMM method, the calculation of  $R^2$  for consecutive mixtures ( $n$  modes vs.  $n + 1$  modes) in conjunction with Fisher statistics, allowed for the development of a stopping condition to automatically detect the optimal aerosol mixture that best fits the AVSD. Note that, while the estimated errors on AERONET-retrieved AVSD are modeled, they serve only as a visual point of reference since they are still yet to be verified. Having said this, application of nonlinear least squares fitting and standard error propagation allowed 95 % confidence bounds to be placed on the multi-modal fits to the interpolated AERONET AVSD in the GMM method.

It is hoped that the methods presented here will help contribute to the vast body of knowledge already provided by AERONET. AERONET retrievals are now being used for the accurate calculation of atmospheric broadband fluxes and aerosol radiative-forcing, and have been shown to agree very reasonably with available coincident ground-based flux observations in desert regions (Derimian et al., 2008) and also globally (Garcia et al., 2008). Furthermore, new retrieval algorithms are being developed to extend the capability of AERONET and to transfer knowledge to new remote sensing domains. For example, Dubovik et al. (2011) developed an inversion procedure for spectral multi-angle, polarimetric, satellite observations from POLDER/PARASOL, and it is hoped that the new methods introduced here will help contribute additional information content as this exciting field evolves.

## Appendix A

### Calculation of secondary microphysical parameters

From retrieved AERONET AVSDs, the volume concentration  $V$  occupied by particles spanning the range of sizes  $r = [r_1, r_2]$  is easily calculated by integrating the  $dV/d\ln r$  over the complete range of values of  $\ln r$ ,

$$C_V = \int_{r_1}^{r_2} \frac{dV(r)}{d\ln r} d\ln r \quad (\text{A1})$$

In principle, the aerosol number size distribution (ANSD)  $dN(r)/d\ln r$  or  $dN(r)/dr$  could equally well be used instead of the AVSD (e.g. see King et al., 1978). The conversion between the volume and number distribution parameters is also straight-forward (see for example Appendix A of Sayer et al., 2012). In particular, the spread  $\sigma$  remains the same for both AVSD and ANSD (King et al., 1978). However, it has been found that the AVSD is preferable to the ANSD as it is more accurate when inverting optical data that is highly sensitive to aerosol particle size (Dubovik et al., 2011). The AERONET inversion code approximates the AVSD using trapezium rule integration (Dubovik and King, 2000) and, while the option of allowing the use of lognormal-shaped bins was included in the calculations of Dubovik et al. (2006), it has only recently been found that accurate modeling of POLDER/PARASOL observations can only be achieved by optimizing the shape of each radial size bin in this way (Dubovik et al., 2011). For an overview of the properties of lognormal distributions in the physical sciences we refer the reader to Limpert et al. (2001). By using a mode separation point  $r = r_s$ , the fine mode fraction  $\eta$  – a key parameter in aerosol forcing estimates (Kaufman et al., 2002) – can then be calculated as follows:

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$$\eta = \frac{V_f}{V_f + V_c} = \frac{\int_{r_1}^{r_s} \frac{dV(r)}{d \ln r} d \ln r}{\int_{r_1}^{r_s} \frac{dV(r)}{d \ln r} d \ln r + \int_{r_s}^{r_2} \frac{dV(r)}{d \ln r} d \ln r} \quad (\text{A2})$$

The fine mode fraction reflects the contribution of the fine mode to the total volume concentration. For desert (mineral) dust in the Sahara and the Arabian Peninsula it is low ( $\eta \approx 25\%$ ), for multi-year averages of biomass burning in Africa and South America, and regional pollution in the eastern US, south-east Asian and Europe, the fine mode contribution is high and spans the range 92–95%, while for maritime aerosol over the Pacific it is more moderate  $\approx 67\%$  (Kaufman et al., 2002). Other important secondary microphysical parameters are statistical measures of central location and dispersion used to characterize individual aerosol modes in the AVSD. The logarithmic volume geometric mean radius (mean logarithm of radius) is a measure of the typical size of aerosol particles and is given by,

$$\ln r_V = \frac{\int_{r_1}^{r_2} \ln r \frac{dV(r)}{d \ln r} d \ln r}{\int_{r_1}^{r_2} \frac{dV(r)}{d \ln r} d \ln r} \quad (\text{A3})$$

The geometric mean radius is obtained by exponentiating the result. In addition, the geometric standard deviation is a measure of the spread (“width”) of the particle mode(s) and is given by,

$$\sigma_V = \sqrt{\frac{\int_{r_1}^{r_2} (\ln r - \ln r_V)^2 \frac{dV(r)}{d \ln r} d \ln r}{\int_{r_1}^{r_2} \frac{dV(r)}{d \ln r} d \ln r}} \quad (\text{A4})$$

## Appendix B

### Comparative statistics measures

Comparative statistics approaches necessarily involve the calculation of the differences (or residuals) between the model data ( $\hat{y}_i$ ) and the target AERONET AVSD data ( $y_i$ ).

5 An initial picture is presented by the bias,  $b$  (or mean of the residuals) which, for  $N$  pairs of data points is given by,

$$b = \frac{1}{N} \sum_{i=1}^N y_i - \hat{y}_i \quad (\text{B1})$$

This statistic is used to assess whether or not models systematically under-predict or over-predict. As a measure of the average difference, we avoid statistics such as the mean relative error and the  $\chi^2$  statistic since, while dependent on residuals, they are fractional quantities and contain  $y_i$  in the denominator. This can inflate their values in the tails of the AVSD where the values of  $y_i$  are extremely small compared to regions, for example, dominated by modal (“fine” and “coarse”) peaks. Instead, statistics involving sums of squares of the residuals were selected that are strongly sensitive to outliers – and hence better able to discriminate between good and bad fits. In particular, we calculated the sum of squares of the residual errors (SSE),

$$\text{SSE} = \sum_{i=1}^N (y_i - \hat{y}_i)^2 \quad (\text{B2})$$

20 From this, the standard error of the fit,  $s$  was calculated,

$$s = \sqrt{\frac{\text{SSE}}{N - p - 1}} \quad (\text{B3})$$

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where  $N$  is the number of points and  $p$  is the number of independent model parameters. This is our choice of location (average) measure and is the unbiased sample estimator version of the traditional root mean square of the errors (RMSE). This measure is sensitive to outliers (due to its dependence on SSE) and is also an interval scale quantity – i.e. it has the same measurement units as  $y_i$ . In order to assess the dispersion (or spread) in the residuals of the fits, we decided to use a regression statistic known as the coefficient of determination ( $R_d^2$ ),

$$R_d^2 = 1 - \frac{\text{SSE}}{\text{SST}} \quad (\text{B4})$$

where SST is the total sum of squares of the difference between the target AVSD data and its *mean* value  $\bar{y}_i$ :

$$\text{SST} = \sum_{i=1}^N (y_i - \bar{y}_i)^2 \quad (\text{B5})$$

$R_d^2$  measures how well a model reproduces data in terms of the amount of the total variance it explains (Steel and Torrie, 1960) and ranges from 0 to 1 such that  $R_d^2 = 0.95$  is taken to mean, for example, that the model fit explains 95% of the total variance in the data. Care must be taken when using this statistic since models involving more modes have more model parameters  $p$  and their improved fit is reflected in a smaller value of SSE. This results in a correspondingly gradual increase in the value of  $R_d^2$  with the number of parameters. To compensate for this effect, we therefore use, in this paper, the *degrees of freedom-adjusted*  $R^2$  statistic which penalizes the value of  $R_d^2$  as extra parameters are included in the model:

$$R^2 = 1 - \frac{\text{SSE}}{\text{SST}} \left( \frac{N-1}{N-p-1} \right) \quad (\text{B6})$$



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This method was scripted in MATLAB using its in-built object-oriented scripting language, and required initial constraints to be placed on the values of the coefficients. For reproducibility of results obtained here, we provide the interested reader with the parameters used in the fitting procedure. The lower bounds were set to 0.0005 since the amplitudes, locations and spreads must be non-zero and positive for each GMM. The upper bounds were set to 3,  $\ln(15)$  and 3 for the amplitudes, log-locations and spreads respectively. The fit at each step was then obtained by minimizing the least absolute (total) residual (LAR). This “Trust region” method also calculates the Jacobian of  $f(x, \beta)$  to determine whether or not the fit is improving (based on the direction and magnitude of the previous adjustment). The minimum and maximum change in the coefficients for this finite difference Jacobian was set to  $10^{-8}$ . Regarding convergence criteria, we set the maximum number of model evaluations in each iteration to the default value of 600 and the maximum number of overall iterations to its default value of 400. The stopping condition on the minimum value of the LAR was set to  $10^{-6}$  which is 1/100th of the minimum volume concentration in our dataset. This entire recipe was then repeated 8 times – being applied to GMMs containing 1 to 8 modes in succession.

Confidence bounds for the optimal GMM (obtained with the fitting procedure that uses the stopping condition outlined in Sect. 3.3) was calculated by the standard approach of propagating errors. This could be achieved because a closed form exists for the optimal GMM as given by Eq. (1) for  $n$ -modes. For each independent mode,

$$f_i = a_i e^{-\left(\frac{\ln r - b_i}{c_i}\right)^2} \quad (\text{C2})$$

the standard error  $SE_i$  is given by,

$$SE_i = \sqrt{\left(\frac{\partial f_i}{\partial a_i} \times SE(a_i)\right)^2 + \left(\frac{\partial f_i}{\partial b_i} \times SE(b_i)\right)^2 + \left(\frac{\partial f_i}{\partial c_i} \times SE(c_i)\right)^2} \quad (\text{C3})$$

in terms of the partial derivatives,

$$\frac{\partial f_i}{\partial a_i} = a_i e^{-\left(\frac{\ln r - b_i}{c_i}\right)^2} \quad (\text{C4})$$

$$\frac{\partial f_i}{\partial b_i} = \frac{2a_i(\ln r - b_i)}{c_i^2} e^{-\left(\frac{\ln r - b_i}{c_i}\right)^2} \quad (\text{C5})$$

$$\frac{\partial f_i}{\partial c_i} = \frac{2a_i(\ln r - b_i)^2}{c_i^3} e^{-\left(\frac{\ln r - b_i}{c_i}\right)^2} \quad (\text{C6})$$

The upper and lower 95 % confidence bounds for the overall GMM fit are then obtained by noting that the standard errors of the modes also combine as root mean squares and are centred on the sum of the modes, i.e.:

$$\text{upper bound} = \sum_{i=1}^n a_i e^{-\left(\frac{\ln r - b_i}{c_i}\right)^2} + 1.96 \sqrt{\sum_{i=1..n} (\text{SE}_i)^2} \quad (\text{C7})$$

$$\text{lower bound} = \sum_{i=1}^n a_i e^{-\left(\frac{\ln r - b_i}{c_i}\right)^2} - 1.96 \sqrt{\sum_{i=1..n} (\text{SE}_i)^2} \quad (\text{C8})$$

**Supplementary material related to this article is available online at <http://www.atmos-meas-tech-discuss.net/6/10571/2013/amtd-6-10571-2013-supplement.pdf>.**

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**Table 1.** Dataset comprising the 4 dominant aerosol type cases studied in this work. Peak percentages are highlighted. Note that SS at Lanai is the least “dominant” aerosol type case with marine aerosol being mixed with other aerosols in the proportion  $\approx 60\% : 40\%$ .

AERONET site	Peak date	SU	OC	BC	DU	SS	OC + BC
Banizoumbou	16 Mar 2005	1.02 %	0.74 %	0.31 %	97.91 %	0.03 %	1.04 %
Mongu	14 Aug 2003	5.61 %	77.36 %	16.76 %	0.22 %	0.05 %	94.12 %
GSFC-Washington	17 Aug 2005	87.53 %	8.31 %	2.72 %	1.38 %	0.05 %	11.04 %
Lanai	21 Feb 2002	28.92 %	5.31 %	2.31 %	3.32 %	60.14 %	7.61 %

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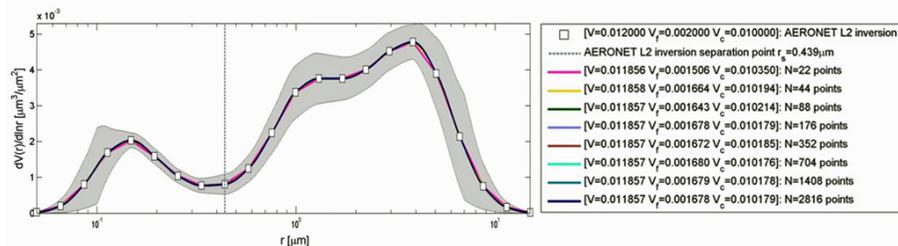
**Table 2.** Statistical testing of  $R^2$  during application of the GMM fit to the interpolated AVSD of dominant marine (sea salt) aerosol at Lanai, Hawaii on the 21 January, 2002. In the case of a single mode, statistical testing is not performed. In the case of both  $n = 2$  and  $n = 3$  modes, the lower confidence limit  $Cl_2(l)$  of the larger-valued  $F(\rho_2)$  is less than the upper confidence limit  $Cl_1(u)$  of the lower-valued  $F(\rho_1)$ , and the t-Welch statistic being  $> 1.96$  shows that there is a statistically-significant improvement in  $R^2$ . In the case of  $n = 4$  modes, two things should be noted. Firstly, that  $F(\rho_2) < F(\rho_1)$  (i.e. a reduction in the improvement in the goodness of fit). Secondly, since the lower confidence limit  $Cl_1(l)$  of the larger-valued  $F(\rho_1)$  is greater than the upper confidence limit  $Cl_2(u)$  of the lower-valued  $F(\rho_2)$  then this reduction is statistically-significant, i.e. the addition of the 4th mode worsens the fit and the optimal number of modes is therefore  $n = 3$ .

$n$ Modes	$R^2(n)$	$R^2(n+1)$	$F(\rho_1)$	$F(\rho_2)$	$Cl_1(l)$	$Cl_1(u)$	$Cl_2(l)$	$Cl_2(u)$	$t$ Welch
1	0.777								
2	0.777	0.819	1.38	1.50	1.34	1.54	1.46	1.54	3.87
3	0.819	0.998	1.50	3.80	1.46	3.84	3.76	3.84	76.26
4	0.998	0.993	3.80	3.17	3.76	3.21	3.13	3.21	20.80

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**Fig. 1.** Interpolation of the AVSD with between 22 and 2816 equidistant logarithmically-spaced points (7 doublings) for dominant marine (sea salt) aerosol at Lanai on the 21 January 2002. Only the interpolations with 22 points (pink) and 44 points (yellow) show visible deviations from the interpolation having the maximum 2816 points. Also shown is the value of the mode separation point  $r_s = 0.439 \mu\text{m}$  provided by the AERONET retrieval on this day. The grey band is the uncertainty on the AERONET AVSD obtained by following the approach described in the next paragraph.

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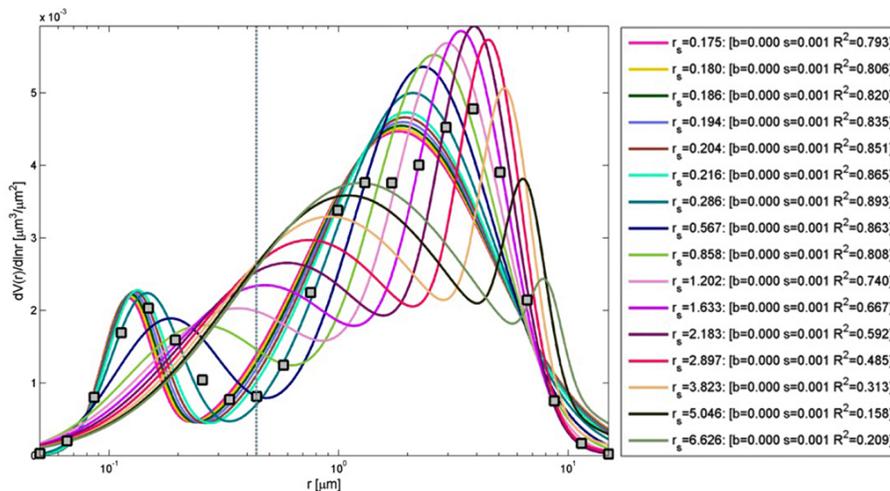
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**Fig. 2.** Sensitivity analysis of the equivalent volume bi-lognormal fit to the AERONET AVSD data with varying mode separation point  $r_s$  for dominant marine (sea salt) aerosol at Lanai, Hawaii on the 21 January 2002. The grey squares are the values of the AERONET AVSD.

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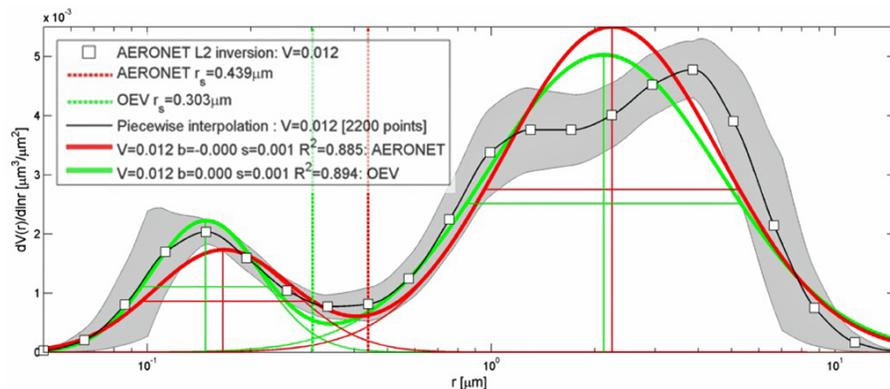
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**Fig. 3.** Comparison of the interpolated AVSD for dominant marine (sea salt) aerosol at Lanai, Hawaii on the 21 January 2002 with the AERONET bi-lognormal fit and the optimized equivalent volume (OEV) fit. The grey band is the uncertainty on the AERONET AVSD.

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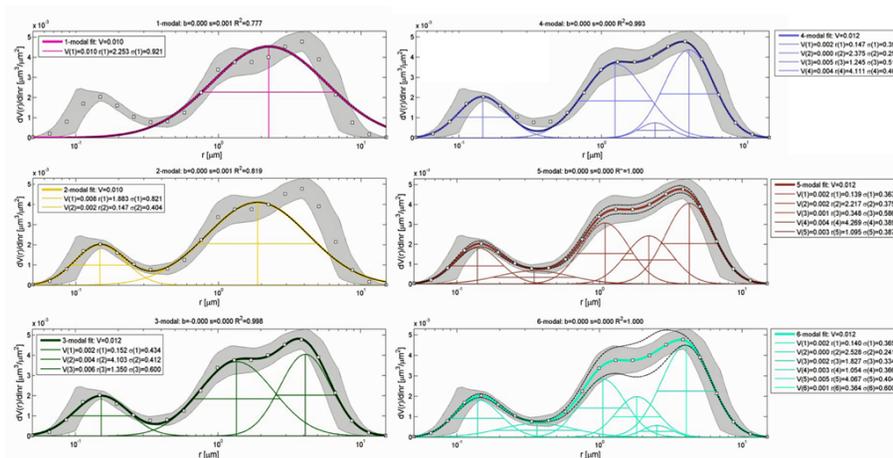
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**Fig. 4.** Gaussian mixture model (GMM) 1–6 modal fits of the interpolated AVSD for dominant marine (sea salt) aerosol at Lanai, Hawaii on the 21 January 2002. The grey shaded region is the error on the AERONET data and the black dotted lines (most visible in the 5 and 6-modal plots) are the 95 % confidence level curves on the fit.

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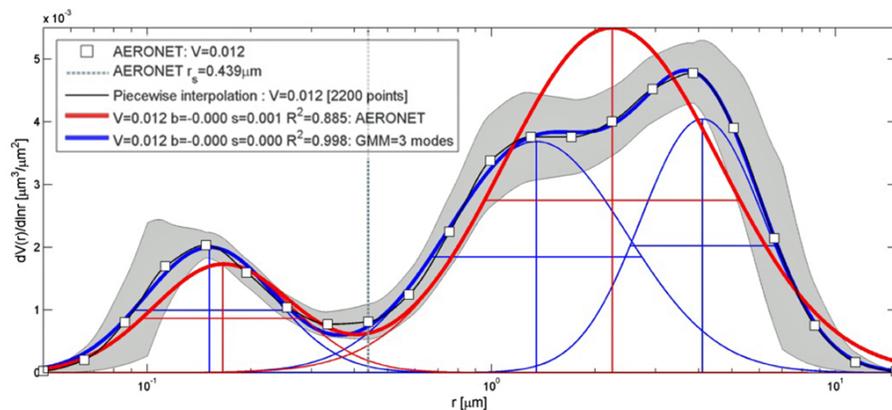
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**Fig. 5.** Comparison of the interpolated AVSD for dominant marine (sea salt) aerosol at Lanai, Hawaii on the 21 January 2002 with the AERONET bi-lognormal fit and the optimal tri-modal Gaussian mixture model (GMM) fit. The grey band is the uncertainty on the AERONET AVSD.

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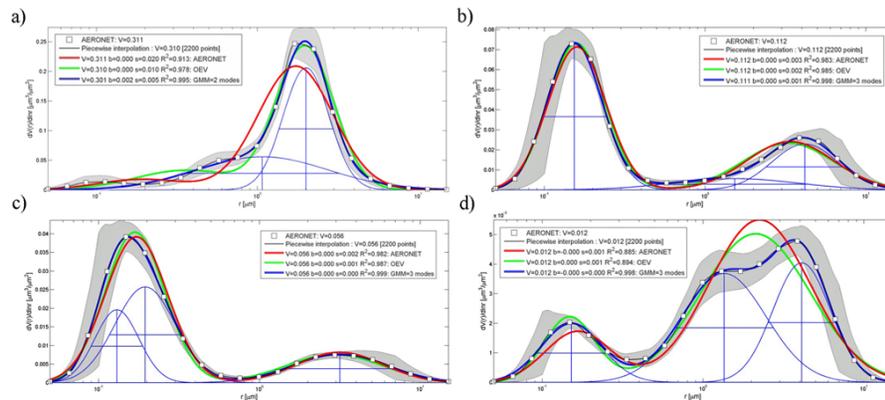
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**Fig. 6.** Comparison of the interpolated AVSD, the AERONET bi-lognormal fit, the OEV bi-lognormal fit and the GMM optimal fit for **(a)** dominant dust aerosol at Banizoumbou, Niger on the 16 March 2005, **(b)** dominant biomass burning aerosol at Mongu, Zambia on the 14 August 2003, **(c)** dominant urban SU aerosol at Washington-GSFC, US on the 17 August 2005, and **(d)** dominant marine (sea salt) aerosol at Lanai, Hawaii on the 21 February 2002. The grey band is the uncertainty on the AERONET AVSD.

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