

Reply to Anonymous Referee #1

This paper describes the production of a harmonised data set for HCHO column abundances from 21 FTIR stations located across the globe. First of all, I must commend the authors for pulling this off. I cannot imagine it was an easy task. Bringing together the HCHO measurements from these different stations/groups is an important achievement, and it will be a valuable resource for modelling studies and for satellite validation purposes. It is a great step forward. I urge the authors to create an online repository where the data can be download easily by others.

Overall I recommend the paper for publication, the science and methods are sound, the results important, and it is well-written. I only have minor comments that need clarifying.

We warmly thank the referee for his/her kind supporting words. Concerning an online repository where the data can be downloaded, this is currently under discussion within the InfraRed Working Group (IRWG) community, and a final decision will be taken at the next IRWG meeting in June. The data will be very likely downloaded in the public NDACC repository. In the meantime, the whole data set is provided on request by myself (corinne.vigouroux@aeronomie.be), or station by station by the individual PIs. I take this opportunity to remind that if any FTIR data is used in a publication, even when the data is released in the public NDACC database, the appropriate PI has to be contacted. If the use of FTIR data is a significant contribution to the publication, a co-authorship should be offered. Otherwise, in agreement with the PI, acknowledgements can be sufficient.

Minor comments:

Please ensure all figures are large in the final version of the manuscript – they are very small and difficult to resolve; it is frustrating. Maybe they could be enlarged by breaking them into separate figures.

The figures have been enlarged in the AMT version.

I noticed that in Table 1 the observing period is very long for some of these stations (e.g., Ny-Alesund). Could the authors possibly comment on the instrument stability and performance over such long periods, and if it affects the HCHO retrievals at all? HCHO is difficult to retrieve is not?

Usually the instrument stability is checked by making regular cell measurements, which allows to verify the good alignment of the instrument (Hase et al. 2009; see also Sect. 2.1 in the present AMTD version). The instruments are re-aligned when a misalignment is detected.

Other type of degradation, e.g. of the mirrors, can easily be seen in the decrease in signal to noise ratio (SNR) of the spectra. This decrease in SNR has a direct impact on the precision of HCHO columns (dominant random error source). Therefore, the precision can indeed vary during the time-series period, but the information is anyway provided in the data sets which include random uncertainties associated to individual measurements. The mirrors need to be regularly replaced to avoid a too low SNR.

The data sets are also quality controlled after the retrieval process. A too low SNR would lead to a bad root-mean-square (RMS) of the fit, and a threshold on this RMS is used at each station to reject the bad quality spectra.

The spectra that do not pass the quality assurance (instrument alignment, RMS,...) are removed from the data sets, leading to small gaps in the time-series as seen in Fig. 5.

Page 7, line 25. The a priori HCHO profile. The approach used here seems sensible, but how sensitive are the retrieval total columns to the a priori – especially as the DOFs is low.

The effect of the a priori profile (and Svar matrix), is calculated in the smoothing error, which has a random but also a systematic component. The systematic smoothing error component, more closely related to the a priori profile itself, was found to be non-significant in our study (1-2% in most cases, therefore dominated by the other systematic sources, which range from 12 to 26%), as was (too shortly, indeed) written p.11, l.10-11.

This small systematic uncertainty was obtained using the following equation:

$$(I-A) (x_a - \langle x \rangle) (x_a - \langle x \rangle)^T (I-A)^T,$$

which accounts for the bias of x_a , i.e. which accounts for the fact that x_a might be different than the expected real $\langle x \rangle$, following von Clarmann (2014).

The $x_a - \langle x \rangle$ is obviously not known (otherwise, $\langle x \rangle$ would be chosen as the correct a priori in the retrievals). Therefore, we had chosen in our AMTD version to use the diagonal elements of the Svar used in Eq. 4 (for the random smoothing error component), as an estimation of $x_a - \langle x \rangle$.

However, Referee#3 also asked for more discussion about this systematic uncertainty. We have therefore added the equation above to the new manuscript. We have also decided to use larger values than the ones from Svar for the revised manuscript: we considered the systematic smoothing error that would occur if the a priori profile is differing from the real $\langle x \rangle$ by -50%, -20%, -10%, +10%, +8%, +5% for the ground-4km; 4-8km; 8-13km; 13-25km; 25-40km; 40-120km layers. The values have to vary with altitude to induce a different a priori profile shape: if 50% is used at all altitudes, the a priori profile is then different from $\langle x \rangle$ by a simple scaling factor, and the systematic smoothing error is close to zero. Using the above values, we obtain systematic smoothing errors from 1 to 9% (median value of 3.4%), which is still small compared to other systematic error sources.

These values rely on the fact that the model WACCM a priori profile shapes are not too far from the reality, which should be the case: due to the known short lifetime of HCHO and its production at or near the surface, we expect that the mean profile peaks at the ground.

This is, as for the random smoothing part, only an estimate of the smoothing systematic error. As discussed in von Clarmann (2014), one would prefer even to not give these smoothing errors at all. We prefer to give them in our paper to provide to the reader as least an idea of the impact of the smoothing in the precision and accuracy of our FTIR HCHO measurements. But these smoothing errors are not provided in the .hdf files that are delivered to the public. When making model or instrument comparisons, the appropriate use of the averaging kernel and a priori profile information (provided in the .hdf files), following Rodgers et al. (2003), allows the user to take implicitly into account the smoothing uncertainty.

This means that, for satellite or model comparison, if Rodgers et al. (2003) is used, there cannot be some different systematic biases at different stations due to different $x_a - \langle x \rangle$.

Figure 3.

The use of atm16 is clearly necessary; HITRAN 2012 needs some corrections...

Indeed. Spectroscopy is often an issue for atmospheric retrievals. We can only wish that more funding is provided for improved spectroscopic measurements.

Table 3: The ‘DIFF30’ is a useful metric, it is given in %, but relative to what? Please be explicit. I’m actually surprised its values are so low (<25%) which is encouraging. Can you also indicate which sites are PROFFIT.

In the AMTD manuscript, the DIFF30 was given relative to the mean of the daily means for historical reasons (in previous work, the metric used was the standard deviation within a day). This explains why the % values did not correspond to the absolute values divided by the given mean (individual) TC (3rd column of Table 3). However, it is better to give this DIFF30 in percent relative to the mean of individual columns. Therefore, even if the numbers are very close, we have corrected the DIFF30 numbers in Table 3, and have specified in the legend to what they are relative. We have also changed the definition of our mean systematic error (7th column in AMTD): in AMTD version we did: mean(individual Syst / individual TC), and we now do: mean (individual Syst) / mean(individual TC). We then avoid too large effect of outliers or negative small columns.

Indeed, the DIFF30 values, which are, given the lifetime of HCHO of a few hours, an “empirical” measurement of the precision in our FTIR measurements, are quite low (median value of 9%) for a species with such very weak absorptions. We are also very pleased to reach such a good precision. The accuracy is less good (calculated as 14%), and this accuracy should be ideally also empirically evaluated by comparisons with correlative measurements.

The PROFFIT sites were already indicated with an * in Table 3 in AMTD. For AMT, we have changed this by ***, to be more clearly seen.

Page 13, lines 14-16. Some units are missing

Corrected for AMT version.

Page 13, last line. Variability faster than 30 mins. Is there any evidence for this in the literature (e.g., from models, campaigns).

We did not find any evidence for this tentative explanation from our side. We removed this sentence for AMT version, since it appears indeed too speculative.

Page 13, line 34: Typo – capital needed at “.this matrix...”

Page 15, line 7, Typo. “1E13” should be x1013

Corrected.

Figure 5: I can understand why this has been plotted but I think it would also probably be good to show the individual measurements for a single (common) year, rather than over the

entire time record at each location. That way you can look at the day to day variability more closely – maybe put such a figure in the supplementary material.

We followed the Referee's suggestion. Instead of putting this plot of one single year in the supplementary material, we give it in Figure 5 in the new manuscript (chosen common year: 2016; except StDenis: 2011). and the complete time-series are given in supplementary material (Figure S1).

Figure 6. The variability in the HCHO observations poses some interesting questions. There is a lot of science in here.

Indeed. We hope that this data set will be soon exploited in modeling studies to explore the reasons for these different observed diurnal cycles.

Some discussion and a few comparisons of FTIR and model diurnal cycles have been provided to answer Referee #3 's concerns about the diurnal cycle (but the model diurnal cycles are still not included in the revised manuscript).

Page 19. The 45% yield reduction – this maybe indeed correct – but can you provide a little more explanation/justification.

We have added the following sentence in the manuscript: "This fraction of 45% is higher, but of the same order, as the estimated overall impact of deposition on the average HCHO yield from isoprene oxidation (28%), based on IMAGES model calculations. The higher fraction for monoterpenes is intended to account for the impact of the more complex chemistry and larger number of oxygenated intermediates involved in their oxidation, compared to isoprene."

Page 23: High mountain sites – at such locations it might be wise to quantify the difference between the station elevation and the model elevation for the 2x2.5 degree grid cells (maybe add information to Table 4). Is there any correlation between this difference and the bias?

Actually, the model takes the altitude of the station into account. We have added the sentence: "The model column is calculated from the calculated formaldehyde profile, between the altitude of the station and the model uppermost level (approximately 20 km), and from the a priori FTIR profile, above that level."

Therefore, the overestimation of the model is related to the coarse resolution ($2 \times 2.5^\circ$), when the mountain site is not in clean area (e.g. Zugspitze, or Altzomoni which is in the same pixel than Mexico City), while the model performs well at mountain sites located in clean areas (e.g. Izaña or Maïdo for which the bias is the same as at StDenis located at the same island but close to sea level).

Reply to Anonymous Referee #3

This study presents the retrieval settings of formaldehyde from ground-based FTIR solar spectra, which has been harmonized to allow for consistent retrievals at various stations, under various conditions (remote area, polluted sites, high-altitude sites...). An error budget is presented for each station. The formaldehyde times-series are then presented along with a preliminary investigation of trend and diurnal cycles. Finally, the consistency of the FTIR products is evaluated via comparison with formaldehyde columns simulated by a chemistry-transport model.

Developing harmonized formaldehyde retrieval settings through the NDACC (and future affiliated stations) is quite challenging because of the weak absorptions of HCHO in the infrared and the many interferences. This work is therefore valuable in the framework of validation efforts of model simulations and of current and future satellite instruments. The topic developed here fits perfectly the scope of AMT. The paper is globally well written and the structure is clear. Nonetheless, some results/figures are not adequately presented, which somehow impedes a proper evaluation of the results (see major comments here below). Therefore, I recommend publication of this study after addressing the comments listed hereafter.

We thank the referee for his/her careful review and his/her constructive comments. We have replied below and changed the manuscript accordingly for the AMT version.

Major comments:

Additional effort is needed to present the results more synthetically and to make some figures easier to read. Fig. 5-8 are particularly difficult to read because of the numerous small panels. It is really unfortunate because these figures present the main results of the study. I assume that the large number of subplots makes them difficult to display, but I really think that such figures deserve a reshape.

We definitively agree that these figures are difficult to read in the AMTD version. We have enlarged them for the AMT version.

In particular:

- The various x-axes in Fig. 5 and 7 don't help the reader. Some seasonal cycles appear completely squeezed because these panels encompass >15 years, while others represent 2-3 years only on a panel of the same size. Please display the time series on an x-axis common to all the stations.

For Fig. 5: This point has been also raised by Referee #1. Following his/her suggestion, we have therefore chosen to show a plot with a common year to all stations (2016), except for St-Denis (for which, due to lack of measurements in 2016, we have chosen 2011). The complete time-series as in Fig. 5 will be given in supplemental material (Fig. S1).

For Fig. 7: the x-axes are already common for all stations in the AMTD version (Jan 2003 – Dec 2016). It is Jul 2002 – Jul 2017 in the AMT version.

For Fig. 5 and 7, I also suggest to gather some time-series within the same panels, e.g., following the subdivision in the text (i.e. clean, intermediate, polluted sites), using different colour lines. The interest to gather time series within the same panels would also be to help the reader to appraise the large panel of HCHO columns covered by the FTIR sites.

We understand the point of view of the referee, and we had hesitated to use a such division in the AMTD version. But, due to Fig. 8 and the discussion with the model comparisons (that mainly follows location/latitude because location is the main reason for different model behaviors), and to increase the facility to a reader only interested in one station to find it at the same place in all Figures and in all Tables, we have decided to keep a latitudinal order.

However, we find the suggestion of the referee to use different colors very helpful to stress the different concentration levels. Therefore, we use three different colors in the new Fig. 5 to 8. (blue, orange and red, for clean, intermediate and polluted sites).

- For Fig. 6, perhaps it is not needed to display all the diurnal cycles in the manuscript. A solution would be to keep here only a few representative of those to support the discussion. The others can be moved to supplementary material.

We followed the Referee's suggestion. Figure 6 gives the diurnal cycles for only 10 of the stations in the new manuscript, and the other ones are given in the supplementary material (Fig. S2).

In Fig. 7-8, it is very hard to distinguish the raw model data from the smoothed ones.

We have enlarged Fig. 7 for a better visibility. However, the main reason why it is hard to distinguish the raw model data to the smoothed one, it that in general the effect of the smoothing is rather small. This is in line with the small reported smoothing uncertainty, and with the fact that the FTIR a priori profile shapes (from the model WACCM) are close to the IMAGES profile shapes.

I also find the discussion on the basis of Table 4 quite “raw”. The authors made huge efforts to harmonize the retrievals and to produce a consistent pattern of HCHO measurements worldwide. There could be, along with Table 4, a map including in colour background the mean HCHO from the model over 2003-2016, and the mean FTIR HCHO in dots filled following to the same colour bar, at the location of each station. In one glance, the reader would have a good overview of the pattern of FTIR measurements as well as of the overall consistency with the model.

Indeed, the discussion of the model comparison is quite “raw” in this paper which aims primarily at presenting the harmonized retrieval strategy and the overall network of HCHO FTIR data sets. The model is only used to show the consistency in terms of bias / seasonal cycles. That is also why

this paper is published in the AMT journal. More discussion on comparisons with models will be the subject of future publications.

However, this is indeed a good suggestion to provide such a map. We prefer to provide it as supplementary material (as Fig. S3; IMAGES climatology for 2005-2015, values in 10^{15} molec/cm²), and we give it also below for discussion (Fig. 1), because the FTIR and model data cannot be quantitatively compared with such a map: the model data are plotted for the 2005-2015 period, while the FTIR data usually cover different time-period (e.g. Porto Velho has only values in 2016-2017). Furthermore, in such a map the model is calculated for the model surface of course, while the mountain sites have a high elevation altitude. For this reason, we prefer to show high altitude sites (>2km) as black crosses. The map shows indeed very clearly the different levels of HCHO covered by the FTIR stations. The agreement with the model looks good (very clean Arctic sites, intermediate European sites, large gradient levels in the Southern Hemisphere from the clean site Lauder, then StDenis, followed by the higher levels at Wollongong, and the strong maximum in Porto Velho). However, this comparison can unfortunately only be qualitative, because of the differences in model and data temporal sampling used for this map.

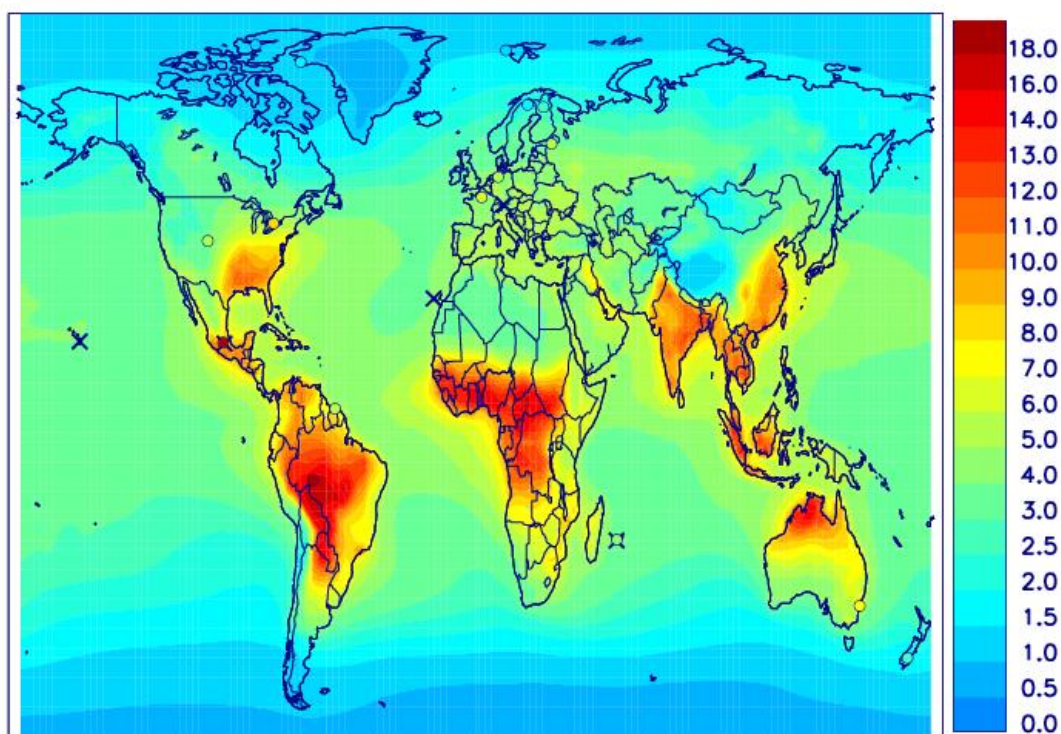


Figure 1: Climatological daytime HCHO columns (2005-2015, 8-17 h local time) calculated by the IMAGES model (in 10^{15} molec/cm²). The long-term averaged HCHO columns at the FTIR stations are shown as filled circles using the same color code. The high-altitude stations (for which the comparison with the model is severely biased due to surface altitude difference) are denoted by crosses.

A single HCHO a priori profile is used at each station for the retrievals. This assumes that not only the shape, but also the HCHO concentration simulated by the model, are quite reliable. What is the impact of another HCHO a priori profile on the retrieved columns? e.g., an a priori from another model that would be significantly different, or again an a priori that would be derived from other measurements (like ACE-FTS)? For a very weak absorber like HCHO, with little retrieved vertical information, I expect the impact to be, if not substantial, at least not negligible. For example, it is clear from the very similar shapes of the a priori and retrieved profiles in Fig. 4, that the retrievals are dependent on the a priori. It is important to discuss this point and to add this component to the error budget.

Indeed, with 1 to 1.5 DOFS only, the retrieved profiles follow the shape of the a priori profiles.

Please, see the detailed reply to Referee #1 about the effect of the a priori profile.

We did not use another model or ACE-FTS data to evaluate the possible difference in a priori, because of the high number of stations there, and the relatively small impact of this smoothing systematic uncertainty. In addition, ACE is not providing profiles down to the ground where the largest difference is expected. Instead, we have chosen a common bias in a priori profiles for all the stations (see reply to Referee #1).

We have re-evaluated it from 1-2% (when S_{var} was used in the AMTD version) to a median value of 3.4%. We have added an equation and some additional text in the new manuscript, and 2 columns in Table 3 (systematic smoothing error and total systematic error).

The error budget is established for each station on the basis of a single measurement. Why one measurement only per station? I find this very reductive, especially that it is not even said whether this single measurement is representative for the whole data set (DOFS, residuals, total column...). Hence, one can easily casts doubts about such error budget. This should be made ideally with a representative subset of FTIR measurements, covering different seasons, different zenith angle, etc.

Sorry if this was not clear in the AMTD version, but the error budget is made for each single measurement, and the mean of all individual errors is reported on Table 3. We think that the misunderstanding comes from the sentence p. 11, l.11-15 (and from the legend of Table 3). When we wrote "... on one individual HCHO FTIR total column measurement.", we meant that the reported uncertainties are valid for a single measurement (and not e.g. for the monthly means that are used for model comparisons). In l.12, it is written "...the mean of the random uncertainty...". But we agree that this is not clear enough, so we have modified the sentence and the legend.

Page 14, Lines 9-16: Since you know that models usually underestimate the natural variability of HCHO, and since you know the impact of such underestimation on the smoothing error estimation, why wouldn't you increase the variability of the model to be more representative for the real variability? Knowing the difficulty for the models to simulate highly-variable reactive gases like HCHO, a model variability multiplied by e.g., 2, would still be conservative.

The estimation of the smoothing error is a delicate subject. It is supposed to give meaningful values when a real variability matrix S_a is available, which is usually not the case. For this reason, the FTIR data sets that are archived in the NDACC database (for official target species such as ozone, CO, CH₄,...) do not include the smoothing uncertainty. The user is asked, if he needs this information, to find by himself the most accurate information available at the time he is using the data (which may not be the same as at the time the data sets are archived in NDACC, if more / better correlative measurements / modeling studies are made available in between). Then he can calculate the smoothing error according the AK provided in the files, and using Eq. 4 of our manuscript.

Another reason for not providing the smoothing error is the fact that if the data are used for validation, then with the appropriate use of the FTIR AK prescribed in Rodgers and Connor (2003), this smoothing uncertainty component is discarded.

Here, we decided to give an estimation of the smoothing error based on our present knowledge of model calculations (here WACCM). This smoothing error can be improved in the future, when model variability will be improved or when more correlative profile measurements will become available. The S_a matrix is currently calculated from several decades of model output, and the dominant variability (also seen in the measurements) come from the seasonal cycles. If the IMAGES model shows indeed usually less variability than the FTIR measurements (Table 4, last column), it seems that it is worse in some cases (Mauna Loa, Izaña) than in other ones (Ny-Alesund, Toronto). So it would seem arbitrary to choose to multiply by e.g. 2 at all stations.

We therefore prefer to keep our estimation as it is in the AMTD version, knowing that it is not perfect.

The example given for Saint-Denis (comparing Vigouroux et al., 2009, where aircraft measurements were used for constructing the S_a matrix, and the present work) looks maybe too extreme (from 14% to 2% for the respective smoothing errors). The S_a matrix used in Vigouroux et al. 2009 had as large values as 70%, which is probably overestimated given the low FTIR variability observed there. We have added a few words in the manuscript about this.

However, if a reader would like to know the smoothing error that would be due to a doubled variability S_{var} matrix (i.e., all – diagonal and non-diagonal – elements multiplied by $2^2=4$) the result is straight-forward: the smoothing uncertainty would be doubled.

Page 15, Lines 15-27: The diurnal cycle is sometimes very weak. Furthermore, the midday observations (low zenith angle) probing less atmosphere, we can expect larger uncertainties associated with such measurements. Hence, owing to these larger midday uncertainties, are the diurnal cycles still significant? Or couldn't some of these diurnal cycles (e.g., the midday minimum found at some sites) be just the effect of larger uncertainties and less sensitivity associated with the low zenith angle?

The averaging kernels (AK) for the midday measurements are similar to the morning and afternoon ones, especially below 15km, where most of the HCHO lies. Only the very high solar zenith angles

give (at some of the stations), significant enhanced degrees of freedom for signal (DOFS), and this enhanced sensitivity is not located at the altitude where most of the HCHO lies. To illustrate this, we give below a plot of the AK (Fig. 2), for a station with a minimum diurnal cycle observed at midday (Sodankyla) and with a station where a maximum is found at midday (StDenis), for local time (LT) = 12 (left panels), and local time at high solar zenith angle, LT=19 and 17 for Sodankyla and StDenis, respectively (right panels). The DOFS (therefore the sensitivity) is very similar in the ground-15 km partial column, and very close to 1 at both stations also at midday (0.98 and 0.96 at Sodankyla and StDenis, respectively). We do not show the plots for e.g. a local time of 15, because they are actually the same as at midday, and the DOFS (total and for the ground-15km partial columns) are exactly the same (0.98 and 0.96 at Sodankyla and StDenis, respectively).

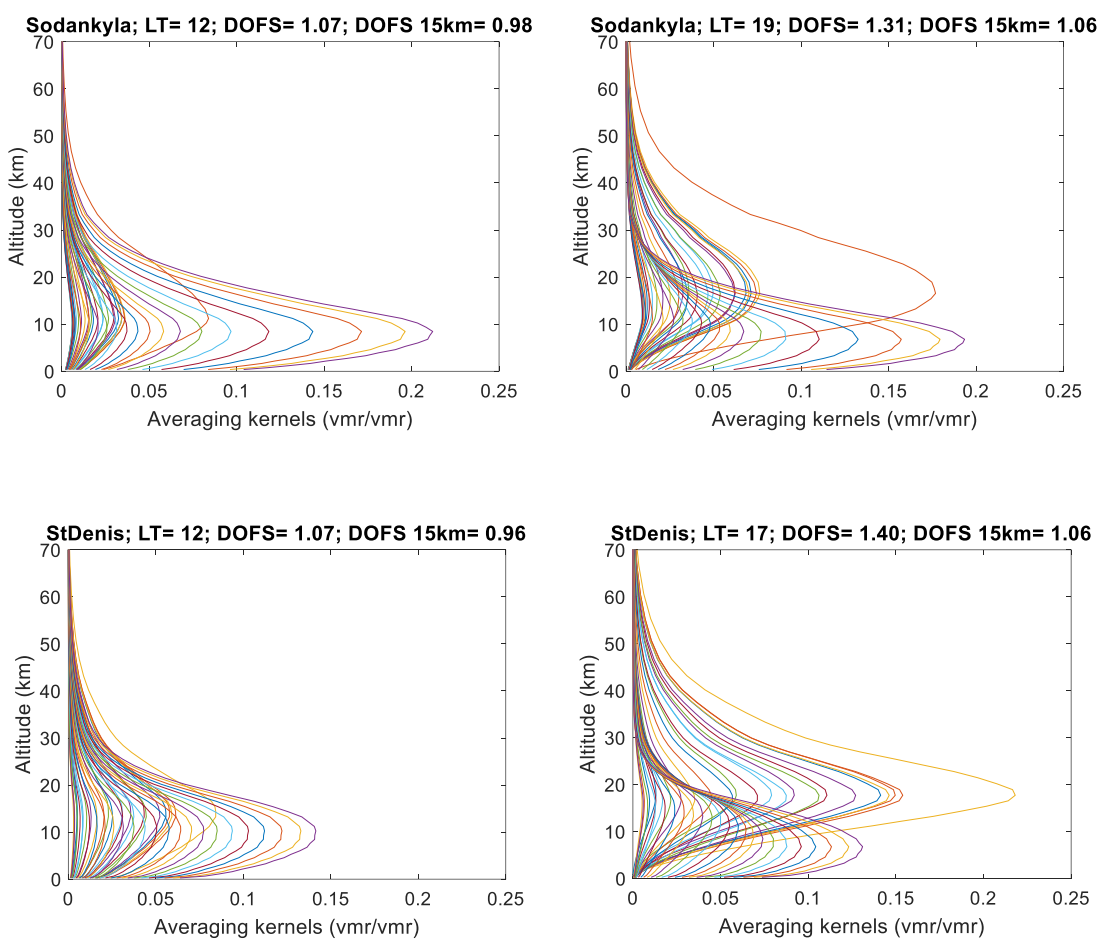


Figure 2: Averaging kernels at Sodankyla (upper panels) and StDenis (lower panels) at Local Time (LT)=12 (left panels) and 17 (right panels). The DOFS for total columns and partial column (ground-15km) are also provided.

Also, the uncertainty budget is not significantly larger for the midday measurements.

Furthermore, the diurnal cycles are really different from station to station, with sometimes indeed a midday minimum, but sometimes a midday maximum, and sometimes a continuous increase from the morning to the late afternoon. Since the technique and the retrievals settings are harmonized among the stations, the Referee's suggested dependence on zenith angle should be reflected in all stations, which is not the case.

We therefore think that the observed diurnal cycles are true.

Maybe, some references to literature is missing in our AMTD version to strengthen our results. The problem is the sparse data providing diurnal cycles, and often at different locations than our stations. And as we just discussed above, the diurnal cycles seem very site dependent. However, in e.g. De Smedt et al. (2015), diurnal cycles from a few MAXDOAS stations show indeed various behaviors: very weak diurnal cycle at OHP (Southern France) in Winter and Spring; wide minimum around midday at Beijing and Xianghe in Spring and Autumn, and constant increase in Summer (as observed for Bremen, Toronto and Paris). This reference to MAX-DOAS observations have been added in the revised manuscript.

We found a paper providing diurnal cycles at Zugspitze (Leuchner et al., 2016). This study is using surface measurements so is not fully comparable to our measurements (therefore, we did not add this reference in the revised manuscript). A rather weak diurnal cycle is found in winter as in our FTIR measurements, but in Spring and Autumn, a maximum is found around midday, in opposition of our minima in these seasons. In summer, a larger diurnal cycle is found, more centered in the afternoon which is also observed from FTIR measurements.

The diurnal cycles observed at the close station Jungfraujoch by FTIR measurements (Franco et al., 2016) are showing, for all months of the year, a midday maximum, which is very different from the Zugspitze FTIR diurnal cycles. The IMAGES model shows diurnal cycles in phase agreement with our FTIR measurements at Zugspitze except for the Summer (Fig. 3 below, upper panels). We also give the FTIR and IMAGES diurnal cycles at Sodankyla, StDenis and Maïdo to illustrate that the model also show different diurnal cycles at different locations and seasons, even very close ones (St-Denis/Maïdo). In some cases, a minimum is indeed calculated at midday, in StDenis the maximum is around midday – 1pm. The model also reproduces quite well more variable diurnal cycles as in Maïdo. We have added the Jungfraujoch diurnal cycles in the discussion, suggesting that more investigation is needed to understand the observed different diurnal cycles.

We did not include the model diurnal cycles in the AMTD paper, because it is a paper focusing on the harmonization of the FTIR data, and there is so much science that can be exploited from these data sets that it deserves several separate papers.

Note also, that the model is not always in agreement with the FTIR diurnal cycles, e.g. the strong maximum at Mexico City is not reproduced., the model providing very weak cycles there (not shown).

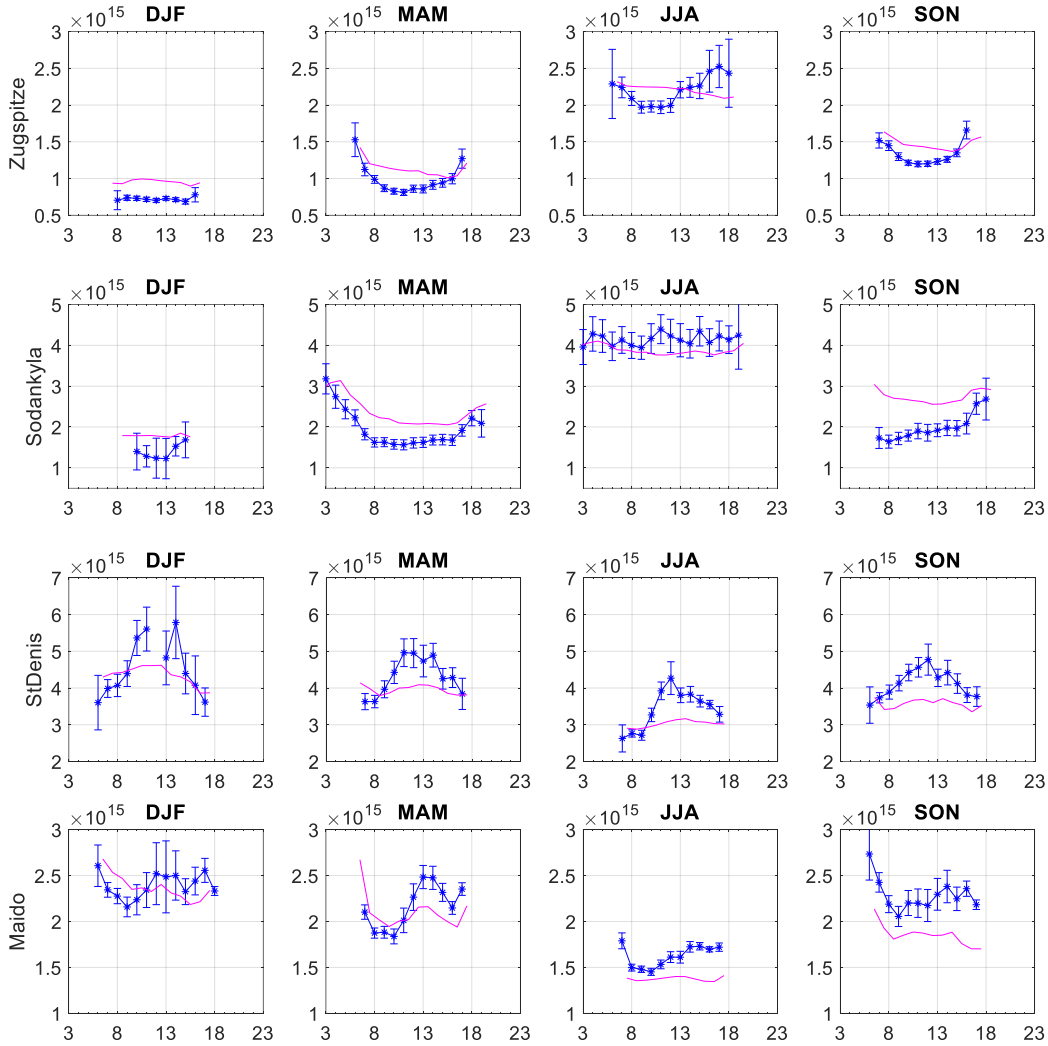


Figure 3: FTIR and model IMAGES diurnal cycles at four stations, for the four seasons.

Section 3.3: The investigation of the trends are very preliminary, and there is no discussion of the results. If, as quoted in this section, a more comprehensive investigation is beyond the scope, I don't see the interest of this section as it is currently. Or I recommend to add a bit of discussion, e.g., how do significant trends compare to other trends in the literature from, e.g.: De Smedt et al. (2010), De Smedt et al. (2015), Franco et al. (2016), Jones et al. (2009).

We thank the referee for the interesting suggestion. We have added such a discussion in the AMT version. The signs of the FTIR observed trends look indeed in good agreement with the previous studies De Smedt et al. (2015) and Franco et al. (2016).

We did not compare with Jones et al. (2009) because the period of concern in this study (1992-2005) is too different from ours (2001-2016). Since the De Smedt et al. (2015) study is going further than De Smedt et al. (2010), we only compare to the last version of the satellite work.

There have already been comparisons between previous HCHO columns from the FTIR and from UV-Vis instruments (satellites and MAX-DOAS), showing in overall a good agreement. However, it is obvious from this study that the HCHO retrievals from the FTIR are very sensitive to the retrieval choices (spectroscopic database, micro-windows, a priori...). Biases up to 50 % are even mentioned between different retrieval approaches. This means that the harmonized retrievals presented here can potentially improve or deteriorate significantly the comparisons with the UV-Vis instruments. I think this point needs to be discussed, or at least mentioned in the conclusion.

Concerning the 3 studies where FTIR have been compared to MAX-DOAS and/or satellite instruments:

- At Reunion Island (Vigouroux et al., 2009): we have compared both FTIR data sets (this “new” study and the “old” 2009 data set), and the bias is only of -1.4% (new-old / old). The comparisons FTIR-MAX-DOAS showed a bias of -8.4%, so the comparisons with the new FTIR data would still be within the systematic error budget on the differences (10%).
- At Lauder: there is indeed a high bias (-49%) between new and old data sets (new-old / old). The new data set would then be in worse agreement with the GOME data set used in Jones et al. (2009), where FTIR data showed a small low bias compared to the satellite. But this needs to be re-evaluated with the reprocessed QA4ECV satellite data (De Smedt et al., 2018), because the satellite products have changed a lot.

Furthermore, the new data set is in much closer agreement with the simulation of four different models that were all of them 50% lower than the old Lauder and Wollongong data sets (Zeng et al., 2015). The Wollongong and Lauder data in Zeng et al. (2015) were using the same retrieval strategies (described in Jones et al., 2009). This high bias between the two strategies is very likely due to the 2869.65-2870.0 cm^{-1} window used in Jones et al. (2009).

- At Jungfraujoch: the strategy used in Franco et al. (2015) does not include the problematic window of Jones et al. (2009). Therefore, we do not expect such a large bias depending on strategies. It requires some time to make new analysis with different strategies (reason why Jungfraujoch is not included in the present study), so this cannot be evaluated for the present review. What we have done up to now to evaluate the impact of the settings is to run our retrievals at Maïdo with the micro-windows (mws) and the spectroscopy used in Franco et al. (2015). If both mws and spectroscopy are changed, then the bias is only of -4% (new-old / old). If the same bias is assumed at Jungfraujoch (which is really “simplified” since the 2 sites have different atmospheric conditions), then the FTIR data would go closer to the model (-8% instead of -12%), therefore in closer agreement as well with MAX-DOAS. However, if only the mws are changed (i.e. if we use the mws of Franco et al. 2015, and the spectroscopy atm16), then a bias of -25% is observed at Maïdo. The small bias of 4% is therefore due to a compensation of opposite effects of mws and spectroscopy. This illustrates again why this is so important to build such harmonized

network where the parameters inducing systematic uncertainty sources are consistent, removing internal bias within the network.

We will add some discussion about this in the AMT version, for Reunion Island and Lauder/Wollongong. We will not do it for Jungfraujoch since the few tests mentioned above was made in Reunion Island atmospheric conditions, so the results are not directly applicable to Jungfraujoch. Once Jungfraujoch will join this harmonization effort, new comparisons with MAX-DOAS will be able to confirm (or not) the discussion above.

Zeng, G., Williams, J. E., Fisher, J. A., Emmons, L. K., Jones, N. B., Morgenstern, O., Robinson, J., Smale, D., Paton-Walsh, C., and Griffith, D. W. T.: Multi-model simulation of CO and HCHO in the Southern Hemisphere: comparison with observations and impact of biogenic emissions, *Atmos. Chem. Phys.*, 15, 7217-7245, <https://doi.org/10.5194/acp-15-7217-2015>, 2015.

Minor comments

Page 1, Line 3: “accurate and precise”. In light of the error budget and the large biases depending on the retrieval choices (“as large as 50%”, Line 5), this statement should be dampened.

This is still a good accuracy when compared to satellite products, and with improved spectroscopy we believe that the accuracy is much better than 50%. We found a median systematic uncertainty of 13% and a maximum of 27%. However, we followed the referee’s suggestion since in the past studies this might be not true. Therefore, we changed to “several independent studies have shown that the FTIR measurements can provide formaldehyde total columns with a good precision”.

Page 1, Line 8: stations. Most of them

Done.

Page 1, Line 11: Change “;” to “,”

Done.

Page 2, Line 1: Unclear. Is it for the systematic or the random uncertainties?

Changed: “Depending on the station, the systematic and random uncertainties of an individual HCHO total column measurement lie between 12 and 27%, and between 1 and 11 $\times 10^{14}$ molec/cm², respectively. The median values among all stations are 13% and 2.9 $\times 10^{14}$ molec/cm², for the systematic and random uncertainties, respectively.”

Page 2, Line 11: NO_x is not defined yet

Done.

Page 2, Line 16: of only a few hours

Done.

Page 2, Line 17: and to test

Done.

Page 2, Line 32: at a few locations

Done.

Page 3, Line 3: Change “geographical” by “spatial”

We prefer to keep “geographical”, because the variability is more related to “geography” (type of land; ocean;...) rather than on spatial distance.

Page 3, Line 3: A lot of efforts are

Native English colleague advises us either “A lot of effort is” or “Lots of effort are”. Therefore, we kept the sentence as it is in AMTD version.

Page 3, Line 7: stations that will also be part

Done.

Page 3, Lines 13-16: Isn’t it because HCHO is so challenging to retrieve that it is not (yet?) a standard product from the NDACC FTIR?

Indeed, the two reasons are somehow linked. But official NDACC targets also include some challenging species, such as HCN and C₂H₆ (small absorptions) or CH₄ (spectroscopic issues). Historical choices have been made in favor of some gases among others.

Page 3, Line 27: monthly mean time-series

Done.

Page 4: Fig. 1 would deserve to be a bit larger due to e.g., the concentration of stations in Europe

Done.

Page 4, Line 14: (1995) and/or Hase et

Done.

Page 4, Line 15: pressure- and temperature-dependent

Done.

Page 5, Lines 5-6: It is said elsewhere in the manuscript that the use of different retrieval parameters can substantially affect the retrieved columns. In particular, the use of either HITRAN 2004, 2008 or 2012 for the HCHO spectroscopic parameters leads to very large differences in the retrieved columns. I would have expected the authors to better motivate their choice of HITRAN 2012, especially that eventually some lines of interfering species needed to be empirically adjusted in this spectroscopic database.

HITRAN 2012 has been chosen because it includes the latest improved HCHO parameters (broadening coefficients, Jacquemart et al., 2010), which complements the release in HITRAN

2008 of new line intensities from the same group (Perrin et al., 2009). We will add this information, and the references, in the manuscript.

The motivation is not coming from the interfering species. Especially, this is true that for some interfering species, the oldest HITRAN versions can be better than HITRAN 2012 (e.g. for CH₄). By the way, the empirical adjustment made in atm16 is, in some cases, simply to use an oldest database (e.g. for H₂O, or, in the 2781 cm⁻¹ window's case, CH₄). The use of atm16 ensures us that for each species the best spectroscopy is used. The work is then done by one of us (G. Toon) and the IRWG community can make use of it, without redoing all the databases comparisons and/or adjustments. Furthermore, this spectroscopy is publicly available (the link is provided in the manuscript), so the users can have easily access to it.

For the Fig.3, it was chosen to plot our spectroscopy atm16 in comparison with HITRAN 2012, in order to stress some remaining problems in HITRAN 2012 for the interfering gases. This can be useful for the spectroscopic community which is often not aware of such specific problems.

Page 6, Line 3: that is distributed

Done.

Page 7, Line 8: Rodgers (2000)), is

Done.

Page 7, Lines 6-9: Is the little gain in information the only reason why you keep these two windows?

No: the sentence says “which contain less absorption for interfering gases”, which is an advantage of these 2 micro-windows. But the gain in information is small, therefore if one decides not to use them, there would be little difference in the retrievals.

Page 7, Lines 10-11: Is this individual spectrum representative for the whole time-series? How do its DOFS and its fitting residuals compare to the other observations?

This individual spectrum has been chosen because it is typical for Maïdo. It corresponds to a column of 25×10^{14} molec/cm², while the mean is 20×10^{14} molec/cm². The DOFS for this spectrum is 1.1 (mean of 1.2), and the root-mean-squares (RMS) of residuals is 0.11 (mean of 0.12).

We will add the information on the DOFS and RMS of this specific spectrum in the manuscript.

Page 7, Line 26: WACCM v4 (Garcia et al., 2007).

The reference is for WACCM v3 (we are not aware of a reference for v4). We propose to change with “the v4 of WACCM (Garcia et al., 2007)”.

Page 7, Line 29: not to use

Done.

Page 7, Line 33: Sussmann et al. (2011), and

Done.

**Page 8, Fig. 2 caption: total column of
(the same in Fig. 3 caption)**

And further, same line: Change “The figures in the lower panel are” to “The lower panels are

Done.

Page 10, Line 31: DOFS, in Table 3

Same line: provide more than

Done.

Page 10, Line 33: (upper panels)

Done.

Page 11, Fig. 4 caption: Upper panels: averaging

And further: Lower panels:

Done.

Page 11, Line 5: associated with

Page 11, Line 6: (lower panels)

Done.

Page 11, Fig. 4: Do you have an explanation for the contribution from the high-altitude averaging kernels (in green)? It looks a bit odd to have a contribution from such layers to the HCHO retrievals. Still Fig. 4: From the shape and the high values of the total column averaging kernel, don't you “overfit” the HCHO retrievals?

The shape of the total column averaging kernel cannot come from an overfitting of HCHO: Kiruna (and other stations who only scaled HCHO) obtains a similar behavior: smaller than 1 below 3-5 km, and larger than 2 above 10-15km. This is also not due to our specific retrieval strategy since this behavior is seen in Jones et al. (2009) and Viatte et al. (2014) as well. In Franco et al (2015) the total column averaging kernel (AK) is not provided.

Therefore, we believe that this is due to the specific spectroscopic parameters and the atmospheric information than can be obtained for the spectra (the localization of the altitude where the variability takes place cannot be well defined, and an underestimation of the variability below 5km is compensated by an overestimation above).

On the other hand, concerning the green high-altitude AKs: these ones appear significant only at stations where the DOFS are larger than 1. In these upper panels, the mean AKs are provided, and if we look at individual retrievals the green ones appear (or are larger) in the case of high solar zenith angles (see Figures above, provided in the diurnal cycle discussion). It is not straightforward to verify if these are real or due to overfitting. The regularization strength has been let to the appreciation of the PIs, using the L-curve method, knowing that the measurement noise and local conditions are site dependent. This is not a very strict method, especially for weak absorption gases such as HCHO, so it might indeed be that the highest DOFS (1.4-1.6) are overestimated. However, the regularization strength (in the reasonable range as provided in our study) has little influence on the retrieved total columns, as tested at Maïdo using a scaling of HCHO: only 2% of bias was observed compared to the present DOFS of 1.2.

We note here that, at the very high solar zenith angles the AKs at Porto Velho were found to have unrealistic values (strong negative oscillations), which was not detected in the AMTD version, since the AKs were good for low to medium/high solar zenith angles. The final retrieval data sets have then be more constrained, leading to a mean DOFS of 1.1 instead of 1.3, with a mean bias effect on total columns of only 1.3\%.

Page 12, Table 3, as well as in the manuscript: Providing the random uncertainties in total column only is really misleading, especially in the discussions. Each time, it forces the reader to look at Table 3 and to calculate the percentage before knowing whether it is significant for the station that is considered. I recommend to provide all the uncertainties (also) in percentage of the total column (you already do it for the systematic uncertainties).

The random uncertainties have been provided in absolute values, since this is the relevant quantity in terms of precision and detection limit. (e.g. satellite requirements for ground-based validation are given in absolute values). Also, similar values are expected among the stations in absolute units, and a percentage value would make a station with clean levels of HCHO appear less precise. The systematic uncertainties are given in percentage because they are expected to be similar among the stations in percentage (not in absolute values).

Therefore, we prefer to keep the absolute values in the discussion for random uncertainties as for the DIFF30 which are calculated to evaluate these random errors. In principle there is no need to make the conversion of random errors in % to follow the discussion. However, since previous studies provided their random error budget in %, we will follow the suggestion of the referee and add for easier comparison the % values for the Total Random error. For the detailed Random and Smoothing errors in %, the reader can simply divide by the mean TC given in the 3rd column.

Page 13, Lines 25-29: Why such exceptions for these SFIT4 stations, and not for others? Isn't the error budget supposed to be fully harmonized among all the SFIT4 stations?

The formalism for error calculation (Rodgers 2000) is harmonized. But, we could have some underestimation/overestimation at some stations due to the Sb matrices, e.g. the Sb matrix for temperature might be realistic at one station and optimistic/pessimistic at another one. Also, the Svar matrix used for the smoothing random error has been taken from the WACCM variability at each station, and the model might provide more realistic variability for some stations than for other ones (as IMAGES does). We see for Table 3 that the smoothing error can contribute significantly to the total random error in some cases (St-Petersburg, Toronto, Boulder, Porto Velho,...) while it has little impact at Lauder and Ny-Alesund (from WACCM).

Page 13, Line 29: might have

Page 13, Line 33: stratosphere. This matrix

Page 13, Line 34: while for the PROFFIT users, these values

Page 14, Lines 3-4: rephrase as “considering the random uncertainty in Table 3 (4th column) is sufficient” to avoid misleading

Page 14, Line 24: HCHO line intensity.

Page 14, Line 26: the PROFFIT channelling source (from 7 to 17 %), which also has a systematic component. We see from Table 3

Page 15, Line 5: few 1 x 10¹³ molec/cm²

Page 15, Lines 15-16: Bad sentence. “To reconcile the different... (afternoon), it is crucial to have ground-based...

Page 15, Line 21: mid-latitude cities

All done.

Page 18, Lines 1 and 5: The use of “time step” is here misleading with the computational time step of the model. I suggest to use “with outputs every 6 hours/ 20 minutes” instead.

We keep “time step” because we indeed mean the computational time step of the model. The outputs are daily averages. This last information is included in the new manuscript.

Page 18, Line 11: delete one “the”

Page 18, Line 22: in Bauwens et al. (2016)

Page 19, Line 1: justified by the

All done.

Page 19, Lines 3-6: How do you deal with the model surface that is below the altitude of the station (which should be the case for most of the mountain sites), especially where there is a substantial altitude difference?

The IMAGES model provides profiles (not total columns), so the columns from the model are calculated from the profiles starting at the altitude of the station (removing the profiles levels between the surface and the altitude of the station). A sentence has been added in the manuscript.

Page 19, Lines 7-9: Do you mean that you re-scale the model outputs at the time of the FTIR measurements? Or do you use the nearest model output to each FTIR data?

The global model output is given once per day. But an offline calculation of the diurnal cycle (at each model pixel) is used to re-scale the model data at the time (hour) of each FTIR measurement. In the first paragraph of Section 4.1, we replaced the sentence "The effect of diurnal variations is accounted for..." by the more complete description: "The model calculates daily averaged concentrations of chemical compounds. The effect of diurnal variations is accounted for through correction factors on the photolysis and kinetic rates obtained from a full diurnal cycle simulation using a time step of 20 minutes. The same model simulation also stores on files the diurnal shapes of formaldehyde columns required for the comparison with FTIR data."

The last sentence of Section 4.1 is replaced by: "Also, the local time of each observation is taken into account by re-scaling the daily averaged concentration using the formaldehyde diurnal shape factors calculated by the model with a time step of 20 minutes."

Page 19, Line 24: “within a month” is redundant

Page 19, Line 25: “The median of IMAGES and FTIR differences” or “The median of IMAGES and FTIR biases”

Page 19, Line 30: change “;” to “,”

Done.

Page 21, Fig. 8 caption: “in coincidence with”. Do you mean the same day, within 20 minutes of each FTIR data?

Yes, the same day. Then, the daily output of the model is scaled at the time (hour) of the FTIR measurement using the diurnal cycle calculated independently using a time set of 20 minutes. (so NOT within 20 minutes of each FTIR data, but same hour).

Page 23, Line 13: Is Boulder (~1600 m asl) really a urban site?

Boulder is a city of about 100 000 habitants. As for Wollongong, the emission sources should be both from biogenic and anthropogenic origin. We will keep Boulder in the “Mid-latitude cities” section, but remove the term “urban” in l.13.

Page 23, Lines 29-30: At such a remote site, the dominant source of HCHO should be CH4 oxidation. I don’t think that other sources from continental areas can be significantly at play here.

The referee is right that this is too speculative. We will change the sentence by “The reasons of the pronounced observed variability are unclear at present.”

Page 24, Lines 5-9: Could it be also due to the FTIR technique, which measures in clear-sky conditions only? The FTIR would sample only air masses free of huge emissions and hence would underestimate the gas abundance in this region.

We don’t believe that the bias is due to FTIR measurements sampling. The sampling of the measurements is actually very high during the biomass burning season due to the low cloudiness during the dry season. Furthermore, the model overestimates the FTIR measurements also during the background season, as seen as well as at Paramaribo. As written in the AMTD version, the overestimation of IMAGES over Amazonia was already found in Bauwens et al. (2016).

Page 25, Lines 14-15: We do not aimed at evaluating the model, but at showing that the FTIR

Done.

I have here an open question, which I think is relevant for a data set that is designed to be used for intense model and satellite validation efforts. Will this data set be made publicly available? And if yes, will there be an effort to fully harmonize the archives, the file format, and the way the data are saved? There are currently inconsistencies between FTIR data sets from different stations (especially for the AKs). Such inconsistencies sometimes refrain external users to use NDACC FTIR data, while such data sets deserve to be easily accessible and as user-friendly as possible for non-community users.

Concerning this data set to be made publicly available: as also replied to Referee #1, this is currently under discussion within the InfraRed Working Group (IRWG) community, and a final decision will be taken at the next IRWG meeting in June. The data will be very likely downloaded

in the public NDACC repository. In the meantime, the data are provided on request by myself (corinne.vigouroux@aeronomie.be).

Concerning harmonization of the archive: we believe that these last few years, lots of effort have been made to harmonize the file format (geoms hdf files). A few inconsistencies remain, e.g. descending or ascending vertical grid, profiles given in the middle of the layers or at the altitude levels. This is quite easy to be solved by the users, but however, it is certainly important to improve this as soon as possible. Furthermore, it is possible that a few stations are still not in line with the current IRWG format requirements. This is indeed crucial for the NDACC database to receive some feed-back from the users, when the inconsistencies still persist. This question will be raised at the next IRWG meeting. We encourage the users to contact the chairmen for reporting inconsistencies: <https://www2.acom.ucar.edu/irwg/contacts>.

NDACC harmonized formaldehyde time-series from 21 FTIR stations covering a wide range of column abundances

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Abstract. Among the more than twenty ground-based FTIR (Fourier Transform infrared) stations currently operating around the globe, only a few have provided formaldehyde (HCHO) total columns time-series until now. Although several independent studies have shown that the FTIR measurements can provide formaldehyde total columns **with a good precision**, the spatial coverage has not been optimal for providing good diagnostics for satellite or model validation. Furthermore, these past studies used different retrieval settings, and biases as large as 50% can be observed in the HCHO total columns depending on these retrieval choices, which is also a weakness for validation studies combining data from different ground-based stations. For the present work, the HCHO retrieval settings have been optimized based on experience gained from the past studies and have been applied consistently at the 21 participating stations. Most of them are either part of the Network for the Detection of Atmospheric Composition Change (NDACC), or under consideration for membership. We provide the harmonized settings and a characterization of the HCHO FTIR products. Depending on the station, the total systematic and random uncertainties of an

individual HCHO total column measurement lie between 12 and 27%, and between 1 and 11×10^{14} molec/cm², respectively. The median values among all stations are 13% and 2.9×10^{14} molec/cm², for the total systematic and random uncertainties, respectively.

This unprecedented harmonized formaldehyde data set from 21 ground-based FTIR stations is presented and its comparison to a global chemistry transport model shows its consistency, in absolute values as well as in seasonal cycles. The network covers very different concentration levels of formaldehyde, from very clean levels at the limit of detection (few 10^{13} molec/cm²) to highly polluted levels (7×10^{16} molec/cm²). Because the measurements can be made at any time during daylight, the diurnal cycle can be observed and is found to be significant at many stations. These HCHO time-series, some of them starting in the 1990's, are crucial for past and present satellite validation, and will be extended in the coming years for the next generation of satellite missions.

1 Introduction

Through reactions with hydroxyl radical (OH) and NO_x (NO+NO₂), the volatile organic compounds (VOCs) exert a strong influence on the oxidizing capacity of the atmosphere. These reactions produce ozone and secondary organic aerosols, which affect air quality and global climate. Given their short lifetimes (from a few minutes to a few hours for the most reactive ones, Kesselmeier and Staudt (1999)) and their different sources depending on geographical locations, it is very difficult to derive a global atmospheric burden for most of the VOCs from current measurements. The observation of formaldehyde (HCHO), which is an intermediate product of the degradation of many non-methane VOCs (NMVOCs) and has a lifetime of only a few hours, allows to constrain the NMVOCs emissions and to test our understanding of the complex and still uncertain degradation mechanisms of these NMVOCs (Stavrakou et al., 2009). The use of satellite HCHO measurements in combination with tropospheric chemistry transport models to derive NMVOCs emissions has been the subject of several past studies (e.g. Palmer et al. (2003); Millet et al. (2008); Stavrakou et al. (2009); Fortems-Cheiney et al. (2012); Barkley et al. (2013); Marais et al. (2014)). The past and present HCHO satellite data sets include those from GOME (1996-2003), SCIAMACHY (2003-2012), OMI (2004-), GOME2A (2006-), OMPS (2011-), GOME2B (2012-), and very recently TROPOMI (2017-). The NMVOCs emissions derived from the top-down approaches using these satellite data sets rely on the accuracy of the measurements. An indirect way to test these accuracies is to compare the emission budgets obtained using two different satellite data sets as in Barkley et al. (2013) for SCIAMACHY and OMI, or in Stavrakou et al. (2015) for OMI and GOME2. While the global emission budgets are in general consistent (Stavrakou et al., 2015), there are large differences on the top-down estimates on a regional scale, e.g. differences up to nearly 50% are observed over Amazonia between SCIAMACHY and OMI (Barkley et al., 2013), and up to nearly 25% between GOME2 and OMI (Stavrakou et al., 2015). To conclude unambiguously whether these differences are due to biases in the satellite products (due to retrieval settings, vertical sensitivities, horizontal resolution,...) or to the diurnal cycle of formaldehyde (SCIAMACHY and GOME-2 measuring in the morning and OMI in the afternoon) requires validation with independent and accurate ground-based measurements (Barkley et al., 2013; De Smedt et al., 2015; Stavrakou et al., 2015).

At present, validation studies of HCHO satellite products have been performed at a few locations only, mainly using aircraft data (Martin et al., 2004; Barkley et al., 2013; Zhu et al., 2016), the MAX-DOAS (Multi-AXis Differential Optical Absorption Spectroscopy) technique (Wittrock et al., 2006; De Smedt et al., 2015) and the FTIR (Fourier Transform Infra-Red) technique (Jones et al., 2009; Vigouroux et al., 2009; De Smedt et al., 2015). This is not sufficient to provide a good picture of the satel-
5 lites' accuracy, especially given the high geographical variability of formaldehyde. A lot of effort is therefore currently under-
way to increase the number of ground-based stations providing HCHO data, using the DOAS or the FTIR technique, initiated
in view of the TROPOMI Cal/Val activities. This paper presents the work accomplished in this direction using FTIR measure-
ments at most of the NDACC (Network for the Detection of Atmospheric Composition Change) stations, and including some
more recent observing stations, that will also be part of the NDACC in the near future.

10 Up to now, time-series of HCHO total columns have been studied at six FTIR stations only, among the more than 20 FTIR
sites currently in operation: Ny-Alesund (Notholt et al., 1997), Wollongong (Paton-Walsh et al., 2005), Lauder (Jones et al.,
2009), Reunion Island (Vigouroux et al., 2009), Eureka (Viatte et al., 2014), and Jungfraujoch (Franco et al., 2015). We note
that HCHO has also been measured by the JPL MkIV instrument (Toon, 1991) at various ground-based sites since 1985 (see
<http://mark4sun.jpl.nasa.gov/ground.html>), although these data are not used in this work due to their very different acquisition,
15 and analysis procedures. The main reasons for having so few FTIR HCHO data available are: 1) that it is challenging to find
robust retrieval settings for this species that has weak absorption signatures in the infrared, which are in addition surrounded
by strong lines from interfering gases; 2) that HCHO is not part of the NDACC FTIR target species (which are O₃, HNO₃, HF,
HCl, CO, CH₄, N₂O, ClONO₂, HCN, and C₂H₆, publicly available at <http://www.ndsc.ncep.noaa.gov/clickmap/>). In the above
cited studies, different retrieval settings are used, although the retrieved HCHO total columns can be very sensitive to some of
20 them: e.g. a positive bias of 30% or even 50% is found at Reunion Island if the settings of Franco et al. (2015) or Jones et al.
(2009) are used, respectively, instead of those from Vigouroux et al. (2009). Although these high biases are consistent with the
uncertainty budgets, it is important, to facilitate the interpretation of a satellite or model validation, to harmonize the settings
among the stations. Therefore, in the present work, we have set up common retrieval settings that can be used at any ground-
based site, even under very humid conditions or low HCHO concentrations. These settings will be described in Sect. 2 together
25 with a characterization of the retrieved HCHO products, i.e., their averaging kernels and uncertainty budget. The complete
time-series of HCHO total columns obtained at the 21 participating stations are shown in Sect. 3, as well as the diurnal cycles
and a short assessment of the long-term trends. We then use the chemistry-transport model IMAGES (Stavrakou et al., 2015),
which provides data for the 2003-2016 period, to show the consistency of our harmonized FTIR data sets: comparisons between
FTIR and IMAGES monthly mean time-series and seasonal cycle at the 21 stations are presented in Sect. 4.

30 **2 Ground-based FTIR HCHO data: description and characterization**

2.1 FTIR HCHO monitoring

Table 1 lists the ground-based FTIR stations included in this study, while Fig. 1 shows their geographical distribution. These
stations perform regular solar absorption measurements, under clear-sky conditions, using either the high-resolution spec-

FTIR stations with harmonized HCHO data sets

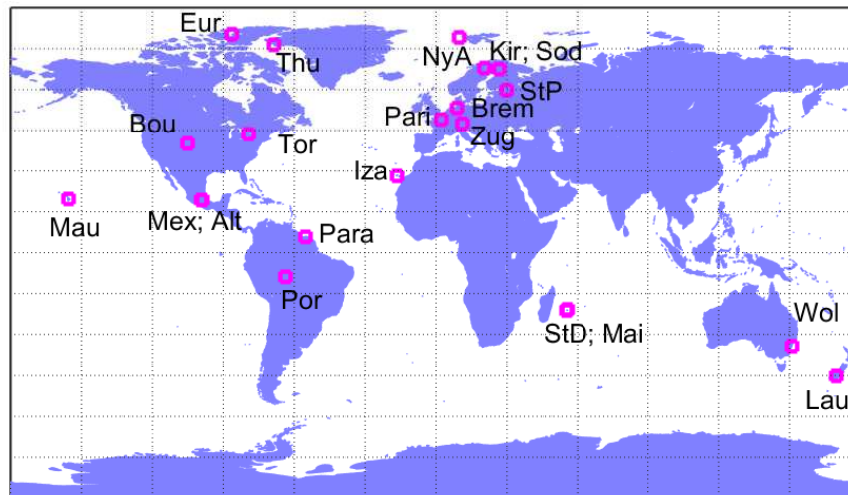


Figure 1. Location of the FTIR stations providing HCHO total columns.

trometers Bruker 120 M, 125 M, 120 HR, and/or 125 HR which can achieve a spectral resolution of 0.0035 cm^{-1} or better, or the Bomem DA8 which can achieve a spectral resolution of 0.004 cm^{-1} . The only lower spectral resolution spectrometer (0.06 cm^{-1}) used in this study is the Bruker Vertex at Mexico City. This instrument is not accepted by the NDACC FTIR standards at present, therefore Mexico City is the only site in this study that will not be part of NDACC.

5 The formaldehyde spectral signatures used in ground-based infrared measurements lie in the $3.6\ \mu\text{m}$ region and belong to the ν_1 and ν_5 bands. This implies that for HCHO, a CaF_2 or KBr beamsplitter and a nitrogen-cooled InSb detector are used, together with an optical filter which usually covers the $2400\text{--}3310\text{ cm}^{-1}$ region (so-called NDSC-3 filter, see e.g. Senten et al. (2008)). At St. Petersburg a broader filter is used ($1700\text{--}3400\text{ cm}^{-1}$). The spectral resolution can be reduced in order to increase the signal to noise ratio (SNR). In practice, the spectra used in the present study have a resolution between 0.0035 and 0.009
10 cm^{-1} , except for Mexico city (0.075 cm^{-1}).

HBr or N_2O cell measurements are regularly performed to verify the alignment of the instruments. The instrument line shape (ILS) can be obtained by analyzing these cell measurements using the LINEFIT program (Hase et al., 1999). This ILS can impact the shape of gas absorption lines, and its determination by LINEFIT can be used as an input parameter in the forward model of the retrieval codes (Sect. 2.2).

15 2.2 Harmonized retrieval strategy

We refer to e.g. Pougatchev et al. (1995) and/or Hase et al. (2004) for more details on the FTIR retrieval principles. Total columns of atmospheric gases, but also volume mixing ratio vertical profiles are obtained from their pressure- and temperature-dependent absorption lines. As seen in Table 1, two retrieval algorithms are used in the NDACC FTIR community: PROFITT9 (Hase et al., 2006), and SFIT2 (Pougatchev et al., 1995) which has been updated to SFIT4 09.4.4. It has been demonstrated

Table 1. Characteristics of the FTIR stations contributing to the present work: location and altitude (in km a.s.l.), time-period used in the present study, instrument type, retrieval code, team.

Station	Latitude	Longitude	Altitude	Time-period	Instrument	Code	Team
Eureka	80.05° N	86.42° W	0.61	2006–2016	Bruker 125 HR	SFIT4	U. of Toronto
Ny-Alesund	78.92° N	11.92° E	0.02	1993–2017	Bruker 120/5 HR	SFIT4	U. of Bremen
Thule	76.52° N	68.77° W	0.22	1999–2016	Bruker 120 M	SFIT4	NCAR
Kiruna	67.84° N	20.40° E	0.42	2005–2016	Bruker 120/5 HR	PROFFIT	KIT / IMK–ASF
Sodankyla	67.37° N	26.63° E	0.19	2012–2017	Bruker 125 HR	SFIT4	FMI & BIRA
St. Petersburg	59.88° N	29.83° E	0.02	2009–2017	Bruker 125 HR	SFIT4	SPb State U.
Bremen	53.10° N	8.85° E	0.03	2004–2017	Bruker 125 HR	SFIT4	U. of Bremen
Paris	48.97° N	2.37° E	0.06	2011–2016	Bruker 125 HR	PROFFIT	LERMA
Zugspitze	47.42° N	10.98° E	2.96	1995–2017	Bruker 120/5 HR	PROFFIT	KIT / IMK–IFU
Toronto	43.60° N	79.36° W	0.17	2002–2016	Bomem DA8	SFIT4	U. of Toronto
Boulder	40.04° N	105.24° W	1.61	2010–2016	Bruker 120 HR	SFIT4	NCAR
Izaña	28.30° N	16.48° W	2.37	2005–2016	Bruker 125 HR		
Mauna Loa	19.54° N	155.57° W	3.40	1995–2016	Bruker 120/5 M	SFIT4	NCAR
Mexico City	19.33° N	99.18° W	2.26	2013–2016	Bruker Vertex 80	PROFFIT	UNAM
Altzomoni	19.12° N	98.66° W	3.98	2012–2016	Bruker 120/5 HR	PROFFIT	UNAM
Paramaribo	5.81° N	55.21° W	0.03	2004–2016	Bruker 120/5 M	SFIT4	U. of Bremen
Porto Velho	8.77° S	63.87° W	0.09	2016–2017	Bruker 125 M	SFIT4	BIRA
Saint-Denis	20.90° S	55.48° E	0.08	2004–2011	Bruker 120 M	SFIT4	BIRA
				2011–2013	Bruker 125HR		
Maïdo	21.08° S	55.38° E	2.16	2013–2017	Bruker 125 HR	SFIT4	BIRA
Wollongong	34.41° S	150.88° E	0.03	1996–2007	Bomem DA8	SFIT4	U. of Wollongong
				2007–2016	Bruker 125 HR		
Lauder	45.04° S	169.68° E	0.37	2001–2016	Bruker 120 HR	SFIT4	NIWA

in Hase et al. (2004) that the profiles and total column amounts retrieved from these two different algorithms under identical conditions are in excellent agreement.

We summarize in Table 2 the forward model and retrieval parameters that have been harmonized. The forward model uses pressure and temperature profiles from NCEP (National Centers for Environmental Prediction) for each site, except that the temporal resolution can vary depending on the retrieval team from daily means, 6-hourly ones, or even hourly interpolated ones.

The dominant source of systematic uncertainty being the spectroscopic parameters, it is crucial that all stations use the same spectroscopic database. We use the compilation from G. Toon (JPL), the so-called atm16 linelist, which is available

Table 2. Summary of the HCHO harmonized forward model and retrieval parameters. The micro-windows limits are given in cm^{-1} .

Pressure and temperature profiles	NCEP
Spectroscopic database	atm16 (=HITRAN 2012 for HCHO)
Solar lines	SFIT4.09.4.4
Micro-windows	MW #1: 2763.42 - 2764.17 MW #2: 2765.65 - 2766.01 MW #3: 2778.15 - 2779.1 MW #4: 2780.65 - 2782.0
Deweighted spectral sections	2780.967 - 2780.993 (O_3) 2781.42 - 2781.48 (CH_4)
Retrieved species	HCHO, HDO, CH_4 , O_3 , N_2O , solar lines optional: CO_2 , H_2O
<i>a priori</i> profiles (except HDO and H_2O)	WACCM v4
Regularization	Tikhonov L1

at <http://mark4sun.jpl.nasa.gov/toon/linelist/linelist.html>. In this atm16 linelist, the HCHO and N_2O lines correspond to the HITRAN 2012 database (Rothman et al., 2013). This HITRAN 2012 database includes the latest improved HCHO parameters (broadening coefficients, Jacquemart et al. (2010)), which complements the release in HITRAN 2008 (Rothman et al., 2009) of new line intensities from the same group (Perrin et al., 2010). The water vapor and its isotopologues lines in atm16 are from Toth 2003¹; some lines of the other strong absorbing gases in the vicinity of HCHO (O_3 and CH_4) have been empirically adjusted or replaced with older HITRAN versions in atm16 when obvious problems were found in the HITRAN 2012 database. For the CO solar lines, we use the linelist updated from Hase et al. (2010) that is distributed in the NDACC community (SFIT4 package v09.4.4).

To avoid any bias between the stations due to different spectroscopic parameters, it is also mandatory to harmonize the spectral micro-windows (MW) containing the HCHO signatures. The challenge of the HCHO retrievals is that this species

¹<http://mark4sun.jpl.nasa.gov/data/spec/H2O/RAToH2O.tar>

has very weak absorption signatures in the infrared (below 1%), and it is therefore very important to minimize the impact of the interfering gases having more intense signatures, either by avoiding micro-windows with strong interfering lines when feasible, or by including them only in case they are very well fitted (e.g. no large residuals remain due to bad spectroscopic or incorrect ILS parameters). In the past studies, while the micro-window spectral widths differ, some common HCHO signatures were used: the two more intense signatures at about 2778.5 cm^{-1} and 2781.0 cm^{-1} were used in all previous studies (Notholt et al., 1997; Paton-Walsh et al., 2005; Viatte et al., 2014; Jones et al., 2009; Vigouroux et al., 2009), except in Franco et al. (2015) who discarded the 2781.0 cm^{-1} signature because of the bad residuals due to poorly fitted CH_4 lines (from HITRAN 2008, Rothman et al. (2009)). In Vigouroux et al. (2009), in which HITRAN 2004 (Rothman et al., 2005) was used, the micro-windows containing these two stronger signatures were quite narrow ($2778.20 - 2778.59$; $2780.80 - 2781.15\text{ cm}^{-1}$), in order to minimize residuals due to neighboring CH_4 lines. With the empirically improved CH_4 spectroscopy in atm16, we can use larger windows (see Table 2 and Fig. 2), with the advantages of fixing more the background and the interfering species, leading to an improved precision and accuracy in the HCHO total columns. We keep the two narrow micro-windows used in Vigouroux et al. (2009) and Franco et al. (2015) at about 2763.5 and 2765.8 cm^{-1} , which contain less absorption from interfering gases, but the gain in information, the so-called degrees of freedom for signal (DOFS, see Rodgers (2000)), is relatively small (0.1-0.2, compared to about 1.0 to 1.5 from the two main windows).

We give in Fig. 2 an example of a spectrum calculated from the retrieval using a spectrum recorded on the 12-02-2014 at Maïdo and corresponding to a retrieved HCHO total column of $2.48 \times 10^{15}\text{ molec/cm}^2$, a DOFS of 1.1, and a root-mean-square (RMS) of 0.11, which compares well to the mean obtained for all measurements at Maïdo of $2.00 \times 10^{15}\text{ molec/cm}^2$, 1.2, and 0.12, for columns, DOFS and RMS, respectively. The corresponding residuals (calculated - observed spectra) are shown in Fig. 3, when the spectroscopic parameters are taken from HITRAN 2012 and with the atm16 empirical linelist. We can see the improvement in MW #1 obtained simply by changing the line position of an O_3 line (2763.8598 cm^{-1} instead of 2763.8588 cm^{-1}). The spectroscopic parameters in MW #2 are the same in both cases, the little improvement seen in this MW is due to the better fitting of the other MWs, that allows better calculated profiles for all gases. The CH_4 line in MW #3 is poorly fitted using the HITRAN 2012 linelist, and the improvement in the atm16 is due to a change in several spectroscopic parameters (line position, line intensity,...). The two more intense CH_4 lines in MW #4 have also been improved by using the atm16 linelist. However, to further improve the fits, one CH_4 line and one O_3 line were empirically deweighted (see Table 2). The comparison of these two linelists shows the crucial need for good spectroscopic parameters in order to obtain precise amounts of atmospheric gases. As seen in Fig. 3 (right panel), the residuals are not perfect and there is still room for further improvement in forward model parameters. The atm linelist created by G. Toon (JPL) is updated each four years when HITRAN provides a new release, so that when the HITRAN linelist is improved and provides either similar or better residuals than the atm linelist, the empirical parameters of atm are changed by the preferred official database.

In SFIT4 and PROFFIT retrieval codes, based on optimal estimation, *a priori* information (profile and regularization matrix) needs to be provided. In this work, the *a priori* HCHO profile, as well as all interfering species except water vapor and its isotopologues, were provided for each station from the v4 of the model WACCM (Garcia et al., 2007). A single profile for each species is used in the time-series retrievals and corresponds to the mean of the model profiles calculated at each station

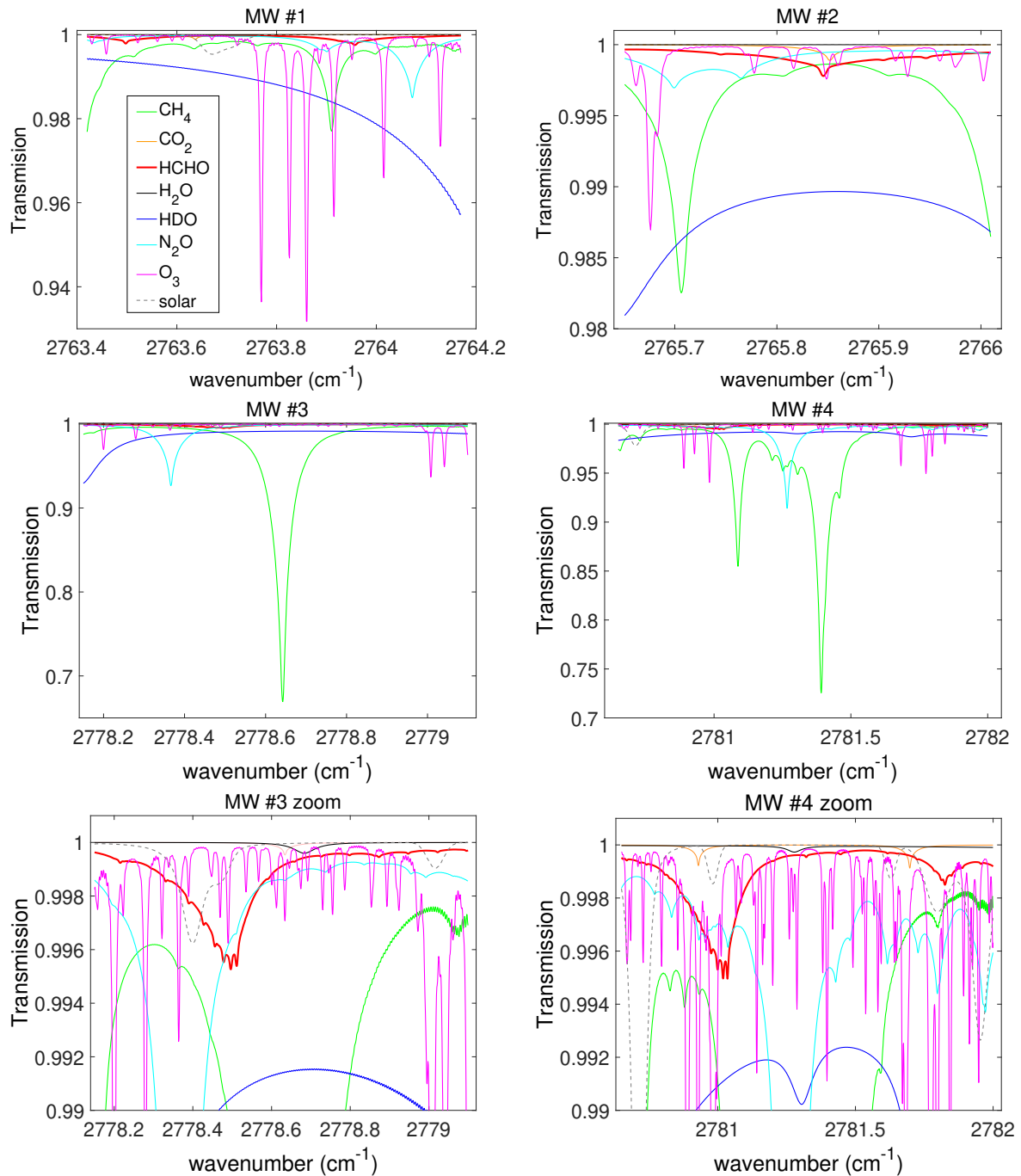


Figure 2. Retrieved contributions of all fitted species in the four MWs (upper and middle panels) used in the analysis for a spectrum recorded on 12-02-2014 at Maïdo and corresponding to a retrieved HCHO total column of 2.48×10^{15} molec/cm². The lower panels are magnifications of the MWs #3 and #4.

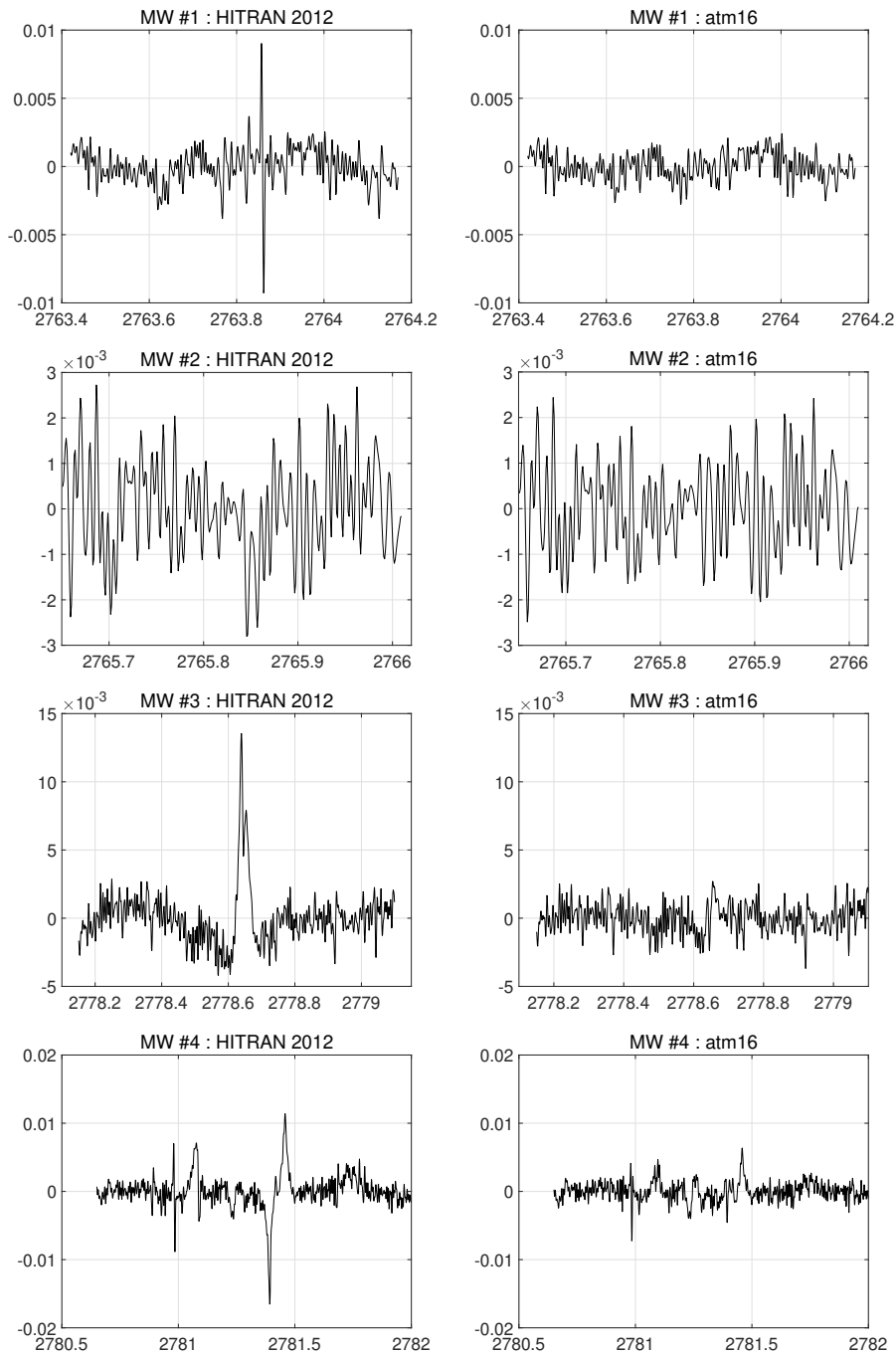


Figure 3. Residuals (calculated - observed spectrum) in each of the four MWs for the retrieval of a spectrum recorded on 12-02-2014 at Maïdo and corresponding to a retrieved HCHO total column of 2.48×10^{15} molec/cm². The x-axis represents the wavenumber in cm⁻¹. The left panels are obtained when the HITRAN 2012 spectroscopy is used, and the right panels show the improvement made by using the atm16 linelist.

from 1980 to 2020. For H₂O and HDO, which have a high atmospheric variability, it is usually preferred (except at the stations Lauder, Mexico City and Altzomoni) not to use a single *a priori* profile: for each individual spectrum, the water vapor *a priori* profiles are taken either from the 6-hourly vertical profiles provided by NCEP, or from independent preliminary profile retrievals. The H₂O absorption being very weak in the chosen MWs, and the HDO profile being retrieved simultaneously with HCHO, the impact of using a single *a priori* profile at the three cited stations is assumed to be small. For the regularization matrix \mathbf{R} , we followed Vigouroux et al. (2009) and Sussmann et al. (2011) and used *ad hoc* Tikhonov (Tikhonov, 1963) L1 regularization as described e.g. in Steck (2002), for the reason that we do not have realistic *a priori* covariance matrix \mathbf{S}_{var} from other measurements sources, especially with a good vertical resolution. The regularization matrix $\mathbf{R} = \alpha \mathbf{L}_1^T \mathbf{L}_1$ is used in most cases for the determination of HCHO low vertical resolution profiles, but also for profile retrievals of the interfering species when improvement is observed compared to the fit of a single scaling factor (which is applied to the *a priori* profiles). This is the case for HDO and CH₄ for which profile retrievals are made, and at some stations for O₃. For the stations Kiruna, Izaña, Zugspitze, and Paris, a scaling of HCHO *a priori* profiles is preferred to a Tikhonov regularization, but due to the low DOFS available for this species (see Sect. 2.3), this has little influence on the retrieved total columns (below 2% when tested at Maïdo). For the other stations, the α values are site dependent, since it can depend e.g. on the HCHO amounts or the SNR of the spectra. Note that the SNR value can be the "real" one coming from the inherent noise in each spectrum, but also can be chosen as an "effective" SNR, that is used as well as a regularization parameter. This effective SNR is smaller than the real one, since the residuals in a spectral fit are not only coming from pure measurement noise but also from uncertainties in the forward model parameters. The regularization choice (α and SNR if an effective one is used) is made at each station in order to obtain stable retrievals (no "overfitting") with significant decrease of the residuals (no "underfitting"), as in the well-known L-curve method (Hansen , 1992).

It is worth noting that another important forward model parameter is the instrumental line shape (ILS) since it impacts the gases absorption line shapes. The treatment of ILS in the retrievals has not been harmonized yet among the stations because the stability and quality of the alignment is site dependent and/or the instrument's PIs have their own preferences. This is however another step toward full harmonization that should be done in the future within NDACC. At present, there are three options for considering the ILS and we refer to Vigouroux et al. (2015) for more details. In the present work, the NIWA, NCAR and University of Bremen stations use a constant and ideal ILS (both modulation efficiency and phase error), i.e. the spectrometers are perfectly aligned. This is a valid approximation based upon LINEFIT ILS analysis of HBr cell spectra measurements (Sect. 2.1). The IMK-ASF, LERMA and UNAM stations use fixed ILS parameters that are previously retrieved using the cell measurements and the LINEFIT code (Hase et al., 1999). For the other stations, the effective apodization parameter is retrieved simultaneously with the target species profiles, while the phase error parameter is assumed to be ideal.

2.3 Characterization: averaging kernels and uncertainty budget

The vertical resolution and sensitivity of the retrieved HCHO products can be characterized by the averaging kernel matrix \mathbf{A} (Rodgers, 2000):

$$\mathbf{A} = (\mathbf{K}^T \mathbf{S}_\epsilon^{-1} \mathbf{K} + \mathbf{R})^{-1} \mathbf{K}^T \mathbf{S}_\epsilon^{-1} \mathbf{K}, \quad (1)$$

5 where \mathbf{K} is the weighting function matrix that links the measurement vector \mathbf{y} to the state vector \mathbf{x} : $\mathbf{y} = \mathbf{K}\mathbf{x} + \epsilon$, with ϵ representing the measurement error. In our retrievals, we assume \mathbf{S}_ϵ to be diagonal, with the diagonal elements being the inverse square of the SNR. \mathbf{R} is the regularization matrix which, in this work, has been chosen as the Tikhonov L1 matrix (see Sect. 2.2).

We give the trace of this averaging kernel matrix \mathbf{A} for the elements corresponding to the HCHO profiles, the so-called DOFS, in Table 3, for each station. The DOFS are ranging from 1.0 to 1.5, meaning that we can not provide more than one
 10 piece of information on the vertical profile. This is the reason why only total columns of HCHO are discussed in this paper, and not vertical profiles. We show in Fig. 4 (upper panels) the averaging kernels (AK, rows of \mathbf{A}), for four different stations, having DOFS ranging from 1 (only scaling) to 1.5. Similar averaging kernels are obtained for the other stations with similar DOFS (not shown). We can observe that, in each case, the AK peak at about the same altitude (8 km) with full-width-at-half-maximum of about 16-18 km, showing that we have limited vertical resolution, and that we are sensitive to the whole troposphere mainly,
 15 and to a lesser extent to the lowermost stratosphere. The total column averaging kernel (TotAK), to be associated with the FTIR retrieved total columns, is plotted as well. The associated *a priori* profiles are also shown in Fig. 4 (lower panels) for completeness, together with the mean and standard deviation of the retrieved profiles. As expected by the low DOFS, the shape of the retrieved profiles is very similar to the shape of the *a priori* profiles.

The uncertainty budget is calculated following the formalism of Rodgers (2000), and can be divided into three different
 20 sources: the measurement noise uncertainty (purely random), the forward model parameter uncertainties (random and systematic), and the smoothing error expressing the uncertainty due to the limited vertical resolution of the retrieval (random and systematic). At each station, the random uncertainty (square root of sum of squares of the measurement noise error and of all the random forward model errors) and the systematic uncertainty (square root sum of the squares of all systematic errors) are calculated for each single measurement. Except for a few cases (NCAR stations and Wollongong) for which a typical
 25 smoothing error is given, and St. Petersburg for which the mean value for 2013 is given, the smoothing uncertainty is also calculated for each individual measurement. We give in Table 3 the mean of the random and systematic uncertainties, of the smoothing uncertainties (both random and systematic parts), and of the total random/systematic uncertainties (square root sum of the squares of the random/systematic error and the smoothing random/systematic error), obtained using the FTIR complete time-series, at each station.

30 The random uncertainty given in Table 3 is dominated at all sites by the measurement noise whose error covariance matrix \mathbf{S}_n is calculated as:

$$\mathbf{S}_n = \mathbf{G}_y \mathbf{S}_\epsilon \mathbf{G}_y^T, \quad (2)$$

Table 3. Mean of the HCHO total columns (TC), in 10^{14} molec/cm², and Degrees of Freedom for Signal (DOFS) obtained at each FTIR station. The stations with strictly 1 DOFS (Kiruna, Izaña, Zugspitze, and Paris) only make a scaling of the HCHO *a priori* profile, i.e. no change in the vertical shape of the *a priori* profile is allowed. We give, in 10^{14} molec/cm², the mean of 1) the random uncertainties (Rand) that were calculated for each individual HCHO total column (excluding the smoothing part); 2) the smoothing random error (Smoo Rand); 3) the total random error (Total Rand= $\sqrt{\text{Rand}^2 + \text{Smoo Rand}^2}$). We also provide the total random error in % for completeness. We give the mean of the systematic uncertainties in %: first without the smoothing part (Syst), then the smoothing systematic error (Smoo Syst), and the total systematic error (Total Syst= $\sqrt{\text{Syst}^2 + \text{Smoo Syst}^2}$). If Rodgers and Connor (2003) methodology is used in model/instrument comparisons, only the Rand and Syst uncertainties need to be taken into account (not the total errors). We provide in addition the mean differences between two subsequent FTIR measurements taken within 30 minutes, in both absolute (10^{14} molec/cm²) and percent units (Diff30), relative to mean TC. The PROFFIT stations are indicated with (***)

Station	DOFS	mean TC	Rand	Smoo Rand	Total Rand	Syst	Smoo Syst	Total Syst	Diff30
Eureka	1.3	12.7	1.0	0.6	1.2 (9.3%)	12.2%	3.5%	12.8%	1.5 (11.7%)
Ny-Alesund	1.6	15.8	1.8	0.5	1.9 (11.7%)	13.3%	3.4%	13.8%	3.9 (24.9%)
Thule	1.1	15.7	1.3	0.9	1.5 (9.8%)	14.3%	3.8%	14.8%	1.8 (11.7%)
Kiruna***	1	17.5	3.5	0.8	3.6 (20.8%)	25.6%	8.6%	27.1%	0.7 (3.8%)
Sodankyla	1.1	25.4	1.5	1.7	2.3 (9.0%)	13.4%	3.8%	14.1%	2.4 (9.3%)
St. Petersburg	1.4	59.4	2.6	2.1	3.3 (5.6%)	13.9%	2.4%	14.2%	2.8 (4.6%)
Bremen	1.2	59.6	2.3	1.7	2.9 (4.8%)	12.9%	2.9%	13.3%	3.1 (5.2%)
Paris***	1	73.0	5.3	1.4	5.5 (7.6%)	16.3%	4.6%	17.0%	3.3 (4.8%)
Zugspitze***	1	12.3	2.2	0.5	2.3 (18.6%)	20.7%	5.8%	21.7%	1.0 (8.0%)
Toronto	1.3	95.1	5.1	4.1	6.7 (7.1%)	12.6%	2.7%	13.0%	19.3 (20.4%)
Boulder	1.1	57.6	2.6	3.9	4.7 (8.2%)	12.7%	2.1%	13.0%	5.3 (9.2%)
Izaña***	1	20.4	3.3	0.2	3.3 (16.0%)	20.9%	4.4%	21.4%	0.8 (4.0%)
Mauna Loa	1.1	10.1	1.4	1.0	1.8 (17.3%)	12.5%	3.8%	13.1%	1.4 (14.0%)
Mexico City***	1.0	220.9	11.1	2.5	11.4 (5.2%)	12.0%	1.2%	12.1%	24.0 (10.9%)
Altzomoni***	1.1	21.8	2.3	1.2	2.6 (11.7%)	16.0%	3.2%	16.3%	2.3 (10.5%)
Paramaribo	1.5	64.3	3.4	1.3	3.6 (5.6%)	12.2%	3.1%	12.7%	11.9 (18.5%)
Porto Velho	1.1	190.0	3.5	8.3	9.1 (4.8%)	12.8%	4.1%	13.5%	5.9 (3.1%)
Saint-Denis	1.2	38.8	2.2	0.8	2.4 (6.1%)	13.4%	4.3%	14.1%	2.8 (7.2%)
Maïdo	1.2	20.0	1.4	0.4	1.4 (7.3%)	12.9%	2.3%	13.1%	1.1 (5.6%)
Wollongong	1.5	78.9	3.0	2.2	3.7 (4.7%)	10.9%	3.0%	11.6%	11.6 (15.0%)
Lauder	1.4	25.6	1.5	0.4	1.6 (6.3%)	12.4%	2.6%	12.8%	3.6 (14.0%)
Median	1.1	25.6	2.3	1.2	2.9 (7.6%)	12.9%	3.4%	13.5%	2.8 (9.3%)

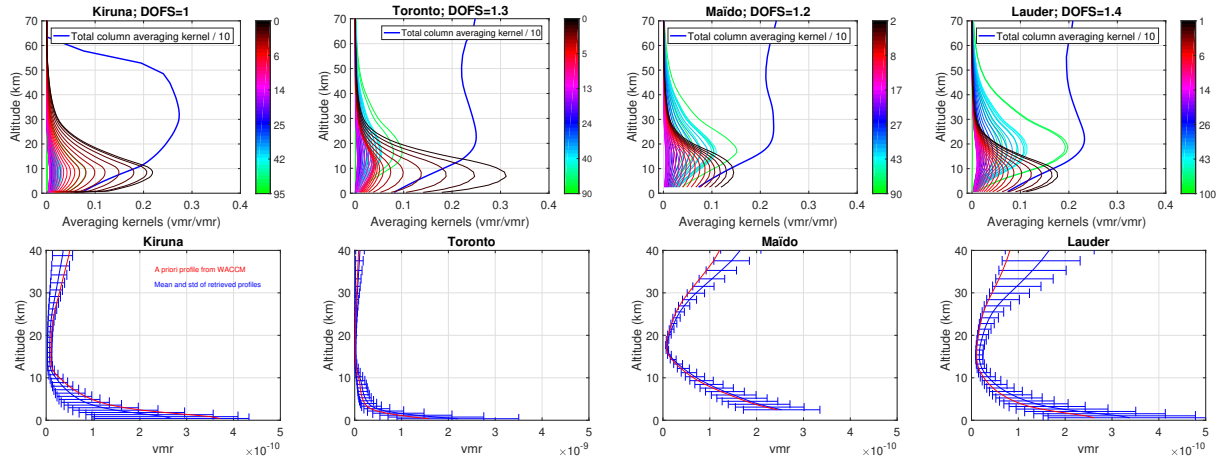


Figure 4. Upper panels: averaging kernels (rows of \mathbf{A}) and total column averaging kernel for four of the FTIR stations, with DOFS ranging from 1.0 to 1.4. The total column averaging kernel is also shown in thick blue line (divided by 10 for visibility). The color code for the different averaging kernels depending on their altitude is given in the color bar in km. Lower panels: *a priori* profiles from the WACCM v4 model (red), and the mean and standard deviation of the retrieved profiles, for the same four stations.

where \mathbf{S}_ϵ is assumed to be diagonal, with the square of the inverse of the SNR as diagonal elements, and \mathbf{G}_y denotes the contribution matrix $\mathbf{A} = \mathbf{G}_y \mathbf{K}$. In this calculation of the measurement noise error, the SNR must be the real one coming from the noise in the spectra, and not a regularization one as can be chosen in the retrieval process (as in Eq. 1; see also Sect. 2.2). For the HCHO spectra used in this study, this SNR can vary between 100 for the worst cases and 3000, with a mean of about 700-1000 for the Bruker 120/5 HR instruments, and 500 for the Bomem DA8.

The forward model parameters error covariance matrices \mathbf{S}_f are calculated according to:

$$\mathbf{S}_f = (\mathbf{G}_y \mathbf{K}_b) \mathbf{S}_b (\mathbf{G}_y \mathbf{K}_b)^T, \quad (3)$$

where \mathbf{S}_b is the covariance matrix of \mathbf{b} , the vector of forward model parameters. For each individual forward model parameter, the \mathbf{K}_b sensitivity matrix is mostly calculated by using analytic derivatives, while the covariance matrix \mathbf{S}_b is an estimate of the uncertainty on the model parameter itself.

Effort has been made in this study to harmonize the uncertainty budget at all sites. This is done by calculating across the network the errors from the same forward model parameters (solar zenith angle, temperature, spectroscopic line parameters, baseline, ...) and by choosing the same \mathbf{S}_b matrix for relevant parameters (i.e. when they are not site or instrument dependent like e.g. for the spectroscopic line parameters). However, some differences remain between the SFIT4 and PROFFIT codes that result in small differences still occurring between the two groups of users, despite the use of harmonized parameters. For the SFIT4 users, the random uncertainty given in Table 3 is dominated by the measurement noise (Eq. 2). We see from Table 3 that the random error is between 1.0 and 3.6×10^{14} molec/cm² for the SFIT4 stations equipped with the high-resolution Bruker spectrometers 120/5 HR or M (the higher values coming from the 120/125 M instruments at Paramaribo and PortoVelho),

while it can reach 5.1×10^{14} molec/cm² with the Bomem DA8 in Toronto. For the PROFFIT users, the random uncertainty is calculated a little bit larger (from 3.5 to 5.3×10^{14} molec/cm²) for the sites with high-resolution spectrometers, and 11.1×10^{14} molec/cm² with the low-resolution spectrometer Bruker Vertex 80 at Mexico City. The main difference between SFIT4 and PROFFIT is the additional error calculated at the PROFFIT stations due to the channeling of the spectra. However, we give also in Table 3 the mean differences between two subsequent FTIR measurements taken within 30 minutes (Diff30), as an upper limit for the total random uncertainty: this difference can be larger than the error budget if HCHO has faster variability than 30 minutes, but with enough statistics, the mean differences should not be lower than the total random errors. We see that this empirical upper estimation of total random uncertainty has a median value (2.8×10^{14} molec/cm²) very close to the median total random uncertainty obtained by error propagation theory (2.6×10^{14} molec/cm²), which gives confidence in the overall FTIR error estimation. At all the PROFFIT sites, except the highly polluted one (Mexico city), the total random uncertainty is larger than the Diff30, which could be an indicator that the uncertainty calculated in PROFFIT is slightly too conservative, probably due to this additional channeling error that would be estimated too large. For SFIT4 users, the Diff30 values are usually close, within 0.5×10^{14} molec/cm², to the calculated total random uncertainty, with the exception of Ny-Alesund and Lauder, where the small calculated errors of 1.9 and 1.6×10^{14} molec/cm², respectively, might be a little bit optimistic, and with the exception of Toronto, Wollongong and Paramaribo where 7 to 13×10^{14} molec/cm² differences are observed between the Diff30 values and the total random errors.

After the measurement noise error (and the channeling for PROFFIT users), the largest contributions to the forward model parameters random uncertainty are coming from the temperature, the interfering species, and the off-set baseline. For temperature, the S_b matrix has been estimated using the differences between an ensemble of NCEP and sonde temperature profiles at Reunion Island, leading to 2 to 4 K in the troposphere and 3 to 6 K in the stratosphere. This matrix is currently used by all SFIT4 users, while for the PROFITT users, these value are chosen smaller (1 K in the troposphere, 2 K up to the middle/upper stratosphere, and 5 K for the highest levels). For each interfering species, the associated S_b matrix is the covariance matrix obtained with the WACCM v4 climatology. At some stations, the ILS is also a high contribution to the random error budget.

If one uses the FTIR HCHO measurements to validate a model or a satellite with fine-vertical resolution, considering the random and systematic uncertainties (without smoothing) in Table 3 (4th column) is sufficient for making correct comparisons, because the smoothing error due to the low-vertical resolution of the FTIR measurements is vanishing if one takes into account the FTIR averaging kernels and *a priori* profile in the comparisons (Rodgers and Connor, 2003). However, if one wants to have a better knowledge of the real precision of the FTIR data themselves, this smoothing uncertainty can be estimated, for the random part, using the smoothing error covariance S_s (Rodgers, 2000):

$$S_s^{\text{rand}} = (\mathbf{I} - \mathbf{A})\mathbf{S}_{\text{var}}(\mathbf{I} - \mathbf{A})^T, \quad (4)$$

where S_{var} should represent the natural variability of the target molecule. For HCHO, this S_{var} variability matrix is unfortunately not well known due to the poor number of vertically resolved measurements. In Table 3, the smoothing errors have been calculated taking the covariance matrices obtained using the WACCM v4 profiles at each station as an approximation of the S_{var} matrices. However, models usually underestimate the variability, and we expect that the smoothing errors provided

here may be underestimated, especially in locations where HCHO is expected to have stronger vertical gradient variability than in the model. As an example, in the study of Vigouroux et al. (2009), the S_{var} was taken from aircraft measurements PEM-Tropics-B, and led to a smoothing error estimation of 14% at St-Denis, while the present estimation based on the WACCM model gives about only 2% for this station. However, the S_{var} matrix constructed from PEM-Tropics-B showed from 33% to 70% of HCHO variability which seems too much compared to what is observed at Reunion Island from the FTIR measurements (about 20%). This illustrates that, ideally, the S_{var} matrix should be re-evaluated at the sites, whenever better model data or correlative measurements become available. The FTIR data sets always including their associated averaging kernel matrices, this can be done *a posteriori* by future users, using Eq. 4.

The smoothing systematic uncertainty, reflecting the bias that would occur on the retrieved profile if the *a priori* x_a is biased compare to the real expected profile $\langle x \rangle$, is calculated following von Clarmann (2014):

$$S_s^{\text{sys}} = (\mathbf{I} - \mathbf{A})(x_a - \langle x \rangle)(x_a - \langle x \rangle)^T (\mathbf{I} - \mathbf{A})^T. \quad (5)$$

The $x_a - \langle x \rangle$ is obviously not known (otherwise, $\langle x \rangle$ would be chosen as the correct *a priori* in the retrievals). Therefore, we have chosen to use $x_a - \langle x \rangle = -50\%$, -20% , -10% , $+10\%$, $+8\%$, $+5\%$ for the ground-4km; 4-8km; 8-13km; 13-25km; 25-40km; 40-120km layers, respectively. The values have to vary with altitude to induce a different *a priori* profile shape: if 50% is used at all altitudes, the *a priori* profile is then different from $\langle x \rangle$ by a simple scaling factor, and the systematic smoothing error is close to zero. Using the above values, we obtain smoothing systematic errors from 1 to 9% (median value of 3.4%), which is small compared to the other systematic error sources (Table 3). These values assume that the model WACCM profile shapes are not too far from the reality, which should be the case: due to the known short lifetime of HCHO and its production at or near the surface, we expect that the mean profile peaks at the ground. This is, as for the random smoothing part, only an estimate of the smoothing systematic error. As discussed in von Clarmann (2014), one would prefer even to not give these smoothing errors at all. We prefer to give them to provide to the reader as least an idea of the impact of the smoothing in the precision and accuracy of our FTIR HCHO measurements. But when making model or instrument comparisons, the appropriate use of the averaging kernel and *a priori* profile information, following Rodgers and Connor (2003), allows the user to take implicitly into account the smoothing uncertainty. This means that, for satellite or model comparison, if the methodology of Rodgers and Connor (2003) is used, there cannot be some different systematic biases at different stations due to different $x_a - \langle x \rangle$.

The dominating systematic uncertainty sources are the spectroscopic parameters: the line intensities and the pressure broadening coefficients of the absorption lines present in our micro-windows. For the HCHO spectroscopic parameters, the linelist in atm16 is following HITRAN 2012 (Rothman et al., 2013), which used the work of Jacquemart et al. (2010), and we use 10% for the three parameters (line intensity, air- and self- broadening coefficients). The larger error source is then the HCHO line intensity parameter, and to a lesser extent the HCHO air-broadening coefficient. In addition, the uncertainties on HCHO columns due to the interfering species spectroscopic parameters are calculated. The dominant ones were found to be due to the pressure broadening coefficients of CH_4 , HDO , and N_2O , with an order of magnitude of about 20% of the error due to the HCHO line intensity.

The other systematic error sources due to forward model parameters are lower or within a few percent (ILS, temperature), except for the PROFFIT channeling source (from 7 to 17%), which also has a systematic component. We see from the Table 3 that the total systematic uncertainty is between 10 and 17% at the SFIT4 stations. For the PROFFIT stations, it lies between 12% and 31%.

5 3 Complete FTIR individual HCHO columns data sets

3.1 A network sampling very low to highly polluted levels of HCHO

We show in Fig. 5 the individual HCHO total columns obtained at each station, for a single year only (2016, except for St-Denis: 2011), in order to better see the day-to-day variability. The complete time-series at each station are shown in Supplementary material (Fig. S1). The error bars in Figs. 5 and S1 are the total random uncertainty, i.e. we do not include the systematic errors in order to better visualize the precision of the FTIR measurements compared to the observed day-to-day variability. The FTIR network samples a wide range of concentrations. Indeed we can distinguish first the “clean” sites (shown with the same vertical axis with maximum 15×10^{15} molec/cm²) such as the Arctic stations (Eureka, Ny-Alesund, Thule, Kiruna, Sodankyla), the marine stations (Izaña, Mauna Loa, Maïdo, St-Denis, and Lauder; the three former being in addition high altitude stations), and the high-mountain sites (Zugspitze and Altzomoni). These clean sites can have HCHO concentrations at the limit of detection (few 10^{13} molec/cm²) with mean values of $10\text{-}25 \times 10^{14}$ molec/cm² (Table 3), except St-Denis which reaches a mean of 39×10^{14} molec/cm².

Second, we show the intermediate concentration sites (with the same vertical axis with maximum 30×10^{15} molec/cm²) such as the tropical coastal site Paramaribo and the mid-latitudes polluted sites in or close to cities and/or vegetation (Peterhof close to St. Petersburg, Bremen, Paris, Boulder). These intermediate sites have mean HCHO total columns of $58\text{-}73 \times 10^{14}$ molec/cm². The sites with the highest levels of HCHO (vertical axis 50 or 70×10^{15} molec/cm²) are Toronto and Mexico City where large anthropogenic emissions are indeed expected (mean of 95 and 221×10^{14} molec/cm², respectively), and places which are also affected by large biogenic emissions such as at Wollongong (mean of 79×10^{14} molec/cm²) and the new site of Porto Velho, located at the edge of the Amazonian forest (mean of 188×10^{14} molec/cm²).

3.2 HCHO diurnal cycle

As explained in the introduction, to reconcile the different results obtained using satellites observing at different time (e.g. SCIAMACHY and GOME-2 measuring in the morning and OMI in the afternoon), it is crucial to have ground-based observations of the HCHO diurnal cycles (Barkley et al., 2013; De Smedt et al., 2015; Stavrakou et al., 2015). The diurnal cycle is also important for model validation, since emissions, chemistry and other processes depend on the time of the day. Our FTIR data set is able now to provide the diurnal cycles at 21 different locations. To separate the effect of the strong seasonal cycle (that will be shown in the next section), we give the diurnal cycle at four different seasons in Fig. 6 for a selection of the sites, while the other ones are provided in Supplemental material (Fig. S2). As seen from Figs. 6 and S2, the diurnal cycles are often

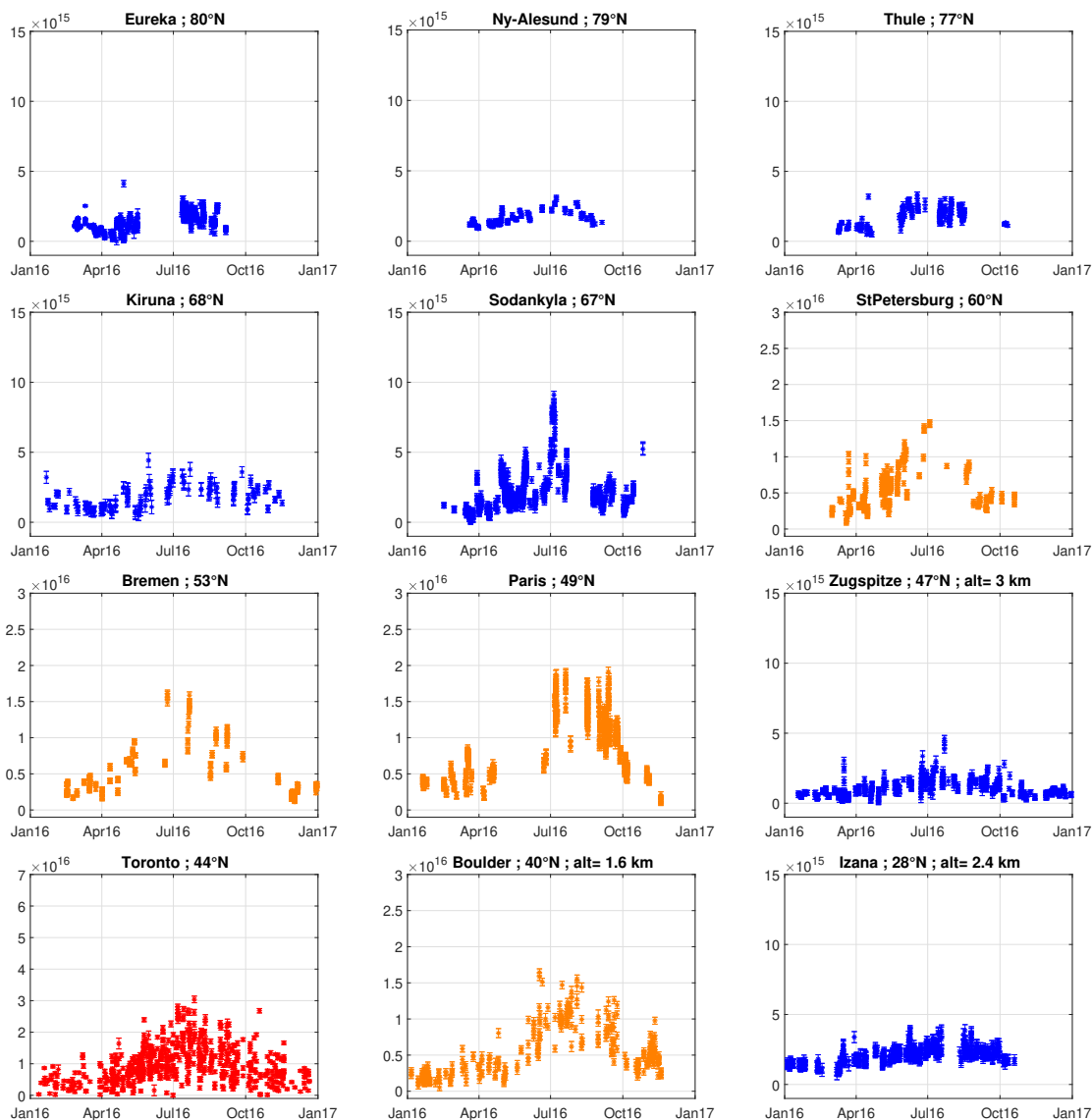


Figure 5. Overview of the individual HCHO total columns (molec/cm^2) at each station, for a single year (2016, except for St-Denis: 2011). The complete time-series at each station are shown in Supplementary material (Fig. S1) The clean, intermediate, high levels HCHO sites are shown using blue, orange, and red colors, respectively. The error bars are the total random uncertainty. When the altitude of the station is higher than 1.5 km, we explicitly give it.

site and season dependent. While there is no clear diurnal cycle at the Arctic sites and at some of the mid-latitude cities during winter (St. Petersburg, Bremen, Toronto), we usually see an increase from the morning, often more pronounced in June-July-August (and Dec-Jan-Feb in Southern Hemisphere), at most of the stations (in the cities, but also at marine sites). A maximum is often found around midday (St. Petersburg, Mexico City, Izaña, St-Denis, Wollongong), or much later in the afternoon (4-

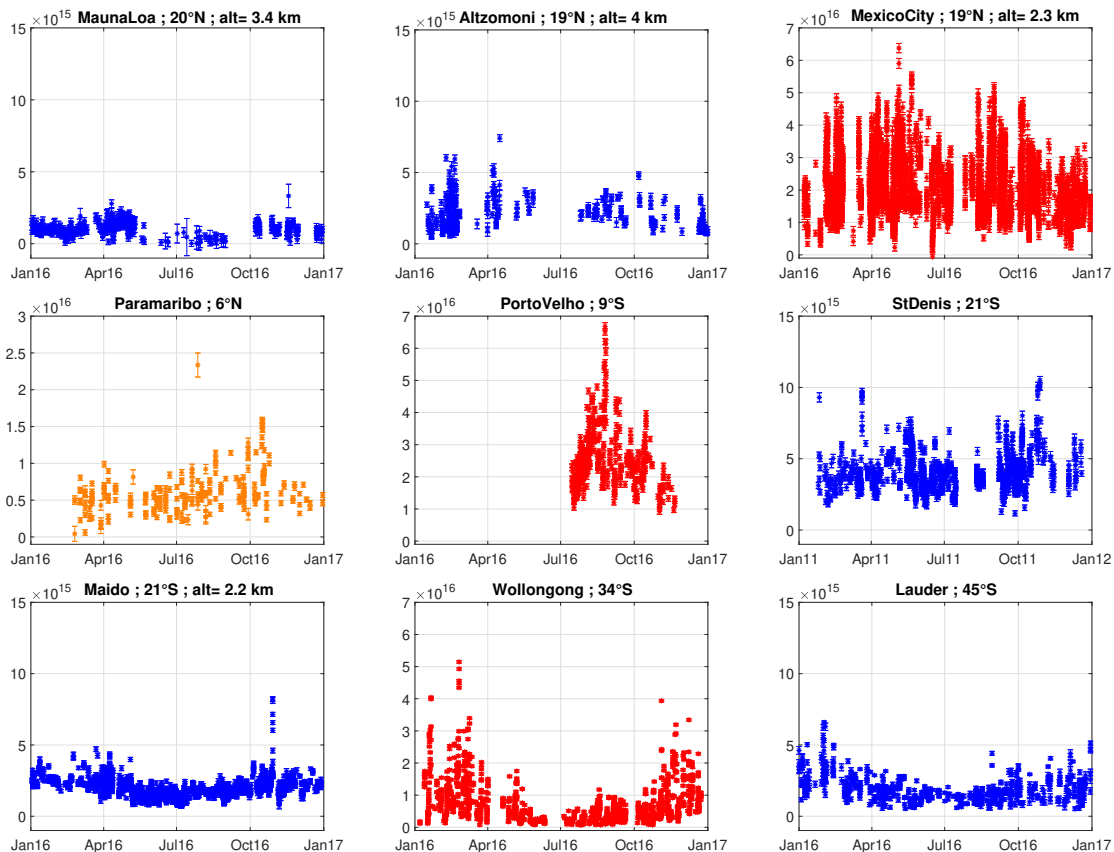


Figure 5. Continued.

6pm), as in Bremen, Paris, Toronto, Lauder, Altzomoni. Only in a few cases, a minimum is found at midday (St. Petersburg in SON, Zugspitze in MAM-SON, Sodankyla in MAM). The marine sites at high altitudes (free of local pollution) have a small minimum at about 8 am (Izaña, Maïdo). This diversity in the FTIR diurnal cycles is also observed with MAX-DOAS at other sites (De Smedt et al., 2015): very weak diurnal cycle at OHP (Southern France) in Winter and Spring; wide minimum around midday at Beijing and Xianghe in Spring and Autumn, and constant increase in Summer (as observed with FTIR for Bremen, Toronto and Paris). The diurnal cycles observed at the Jungfrauoch station by FTIR measurements (Franco et al., 2015) are showing, for all months of the year, a midday maximum, which is very different from the ones observed at our close station Zugspitze. The IMAGES model shows diurnal cycles in phase agreement with our FTIR measurements at Zugspitze except for the Summer for which the model diurnal cycle is very weak (not shown). However, two very close sites can indeed observe different diurnal cycles (as seen for St-Denis and Maïdo). More investigation is needed to understand the different diurnal cycles at these two close mountain sites.

We see from Fig. 6 that the FTIR measurements at Porto Velho do not show a clear pattern, in particular if one is interested in the 9:30 and 13:30 differences between the overpass of two different satellites (De Smedt et al., 2015). From the one year of

data available at present at this new site, it seems that the diurnal cycle cannot help to reconcile the differences observed over Rondonia between GOME-2 and OMI (De Smedt et al., 2015). In contrast, the diurnal cycles observed over cities confirm that the observation of a positive bias between OMI (13:30) and GOME-2 (9:30) over urban areas can be indeed explained, at least partly, by the diurnal cycle.

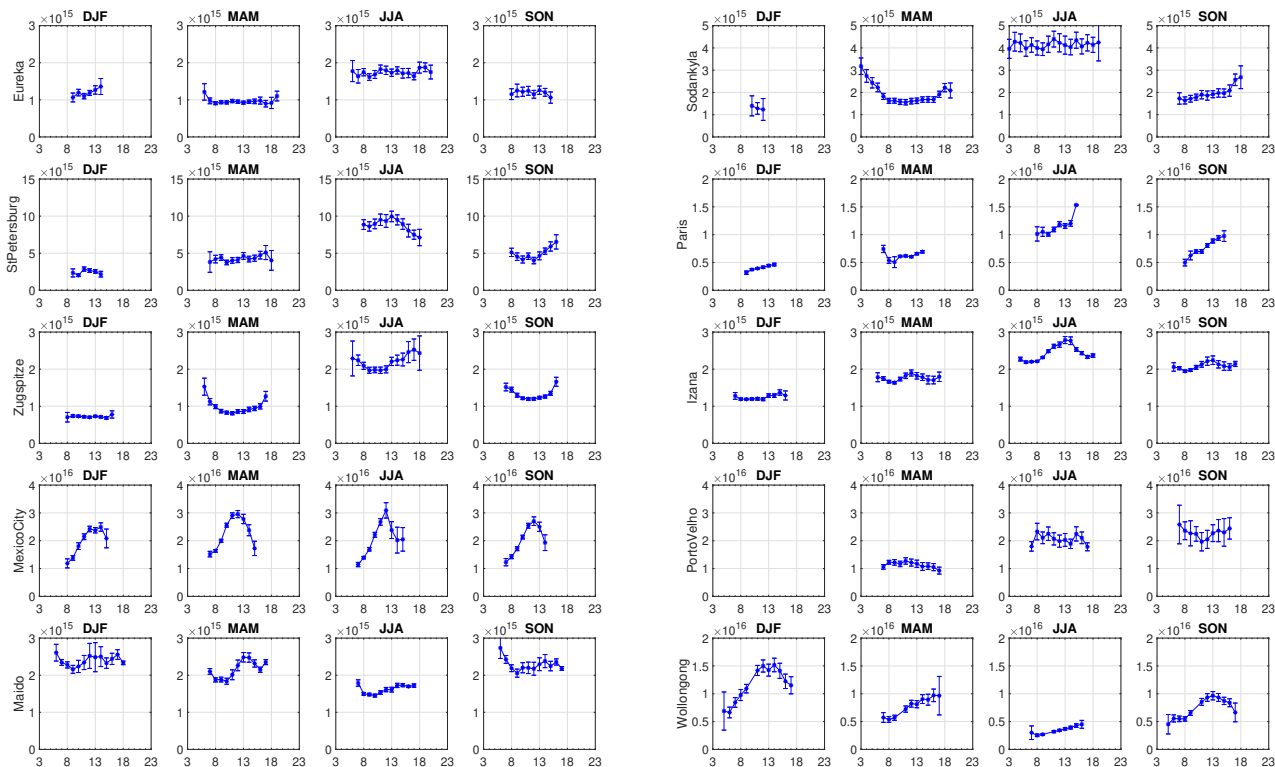


Figure 6. Diurnal cycles of HCHO total columns (molec/cm²) at selected stations for the four seasons. The diurnal cycles for the other stations are shown in Supplementary material (Fig. S2). The error bars are the standard errors on the mean: $2 \times \sigma / \sqrt{n}$, with σ the standard deviation and n the number of measurements at a given time. If $n < 8$, the hourly value is not shown. The time is the Local Standard Time Meridian (LSTM).

5 3.3 Long-term HCHO trends

The length of the HCHO time-series allows trends to be derived for some stations. We have calculated the trends at each station using the monthly mean time series $Y_m(t)$ with a simple model including a fit of the seasonal cycles:

$$\begin{aligned}
 Y_m(t) = & A_0 + A_1 \cdot \cos(2\pi t/12) + A_2 \cdot \sin(2\pi t/12) \\
 & + A_3 \cdot \cos(4\pi t/12) + A_4 \cdot \sin(4\pi t/12) \\
 & + A_5 \cdot t,
 \end{aligned}$$

10

with A_5 the annual trend.

It turned out that, due to the very high variability of HCHO, the uncertainties on the trends are often too large to obtain significant values. A more sophisticated multi-regression model might be able to reduce the uncertainties, but this is beyond the scope of this paper. However, for a few stations, significant trends are found. They are mainly negative: at St. Petersburg (-3.9 ± 3.3 %/dec), Mexico City (-9.6±5.1 %/dec), Wollongong (-18.8±10.8 %/dec), and close to significance at Zugspitze (-7.7±7.7%/dec). Only the marine sites Izaña and St-Denis show positive significant trends (+17.3 ± 15.2 and +15.8 ± 5.2 %/dec, respectively). Note that at Maïdo, the trend is not significant. A careful combination of the measurements at both Reunion Island sites (St-Denis + Maïdo) could be carried out in the future.

For the longest time-series, we observe in general a very good agreement with previous studies. The negative trends observed over the European stations St. Petersburg and Zugspitze are in agreement with the negative trends observed by OMI (2004-2014) over St. Petersburg and Germany (De Smedt et al., 2015). At the Jungfraujoch station, a negative trend (-6.1±2.6%/dec) was also observed for the 1996-2015 period (Franco et al., 2016). Note that the calculation of the uncertainties on the trends in our study takes into account the auto-correlation in the residuals, following Santer et al. (2000), which increases the uncertainties. For e.g. Zugspitze the uncertainty without correcting for this auto-correlation, as in Franco et al. (2016) or De Smedt et al. (2015), would be 4.9% (instead of 7.7%), showing then a more significant trend, in agreement with these studies. The non significant trends observed at the Northern European station (Kiruna), and the mid-latitude American stations (Toronto, Boulder) are in agreement with De Smedt et al. (2015). In the Southern Hemisphere, the negative trend observed at Wollongong was also found in De Smedt et al. (2015), as well as a positive trend at Madagascar, close to Reunion Island, in agreement with the high positive trend observed at StDenis. At Lauder, OMI also shows non significant trend (De Smedt et al., 2015).

4 HCHO FTIR and IMAGES model comparisons

In this study, we do not aim to validate the model input parameters or to attribute different emission sources at the different stations. We use the model to assess the internal consistency of the network using harmonized retrieval settings. This means that we expect that for the same latitude regions and/or type of sites (polluted; clean), the comparisons with the model will give consistent biases.

4.1 IMAGES model description

The IMAGESv2 global model calculates the distribution of 170 chemical compounds gases with a time step of 6 hours at $2^\circ \times 2.5^\circ$ resolution, with 40 a hybrid (σ -pressure) levels in the verticals between the surface and the lower stratosphere (44 hPa level). The model calculates daily averaged concentrations of chemical compounds. The effect of diurnal variations is accounted for through correction factors on the photolysis and kinetic rates obtained from a full diurnal cycle simulation using a time step of 20 minutes. The same model simulation also stores on files the diurnal shapes of formaldehyde columns required for the comparison with FTIR data. Meteorological fields (winds, temperature, humidity, 3-dimensional cloud cover, solid and liquid cloud water content, large-scale and convective precipitation rates, visible downward radiation, convective updraft fluxes,

boundary layer diffusivities, snow depth, sea ice fraction, surface roughness lengths, surface sensible heat flux, friction velocity, etc.) are obtained from ERA-Interim analyses of the European Centre for Medium-range Weather Forecasts (ECMWF).

Anthropogenic emissions of NO_x, CO, SO₂, and NMVOC are provided by the Hemispheric Transport of Air Pollution dataset version 2 (HTAPv2) (Janssens-Maenhout et al., 2015), with the NMVOC speciation provided by the emission inventory of the Atmospheric Chemistry and Climate Model Intercomparison Project (ACCMIP) (Lamarque et al., 2010). Emissions from open vegetation fires are taken from the last version of the Global Fire Emissions Database, GFED4s, which includes the contribution of small fires (Randerson et al., 2012; Giglio et al., 2013). The GFED data are available at daily frequency at 0.25° × 0.25° from 1997 through the present (<http://www.globalfiredata.org>). The vertical distribution of these emissions follows Sofiev et al. (2013). Isoprene and monoterpenes emissions are obtained from the MEGAN-MOHYCAN model (Müller et al., 2008; Stavrou et al., 2014; Guenther et al., 2012) for all years of the study period at a resolution of 0.5° × 0.5° (<http://tropo.aeronomie.be/models/isoprene.htm>). Methanol biogenic emissions are obtained from the inverse modeling study of Stavrou et al. (2011). Besides the dependence on temperature, visible radiation, leaf area and leaf age, the model accounts for the inhibition of isoprene emissions under drought conditions through a dimensionless soil moisture activity factor (γ_{SM}). However, the parameterization of γ_{SM} is very uncertain, as discussed in Bauwens et al. (2016), and we assume $\gamma_{SM} = 1$ in this study. The average global annual emissions are 419 Tg/yr isoprene, 100 Tg/yr methanol and 103 Tg/yr monoterpenes.

The chemical degradation mechanism of pyrogenic NMVOCs is described in Bauwens et al. (2016). The oxidation mechanism for isoprene is also based on Bauwens et al. (2016), with a few updates. It accounts for the revised kinetics of isoprene peroxy radicals according to the Leuven Isoprene Mechanism version 1 (LIM1) (Peeters et al., 2014) and further modified to account for laboratory findings (Teng et al., 2017; Bates et al., 2016). The formaldehyde yield in isoprene oxidation by OH is close to 2.4 mol/mol in high NO_x (1 ppbv NO₂, after 2 months of simulation) and 1.9 mol/mol at low NO_x (0.1 ppbv NO₂). The chemical mechanism for monoterpenes is simplified, with product yields of formaldehyde, acetone, methylglyoxal and glyoxal based on box model calculations using the α - and β -pinene oxidation mechanism from the Master Chemical Mechanism (MCM) (Saunders et al., 2003). The overall formaldehyde yield is 4.2 HCHO per monoterpene oxidized, coming down to 2.3 after subtracting the contributions of acetone and methylglyoxal oxidation. This yield is further reduced by 45% to account for wet/deposition of intermediates and secondary organic aerosol formation. This fraction of 45% is higher, but of the same order, as the estimated overall impact of deposition on the average HCHO yield from isoprene oxidation (28%), based on IMAGES model calculations. The higher fraction for monoterpenes is intended to account for the impact of the more complex chemistry and larger number of oxygenated intermediates involved in their oxidation, compared to isoprene. The large deposited fraction is uncertain, but appears justified by the larger number and lower volatility of intermediates involved in formaldehyde formation from monoterpene oxidation.

The calculation of the model columns at the FTIR stations accounts for its location in the horizontal (nearest model pixel), for the FTIR *a priori* profiles and averaging kernels as prescribed in Rodgers and Connor (2003), as well as for the station altitude above sea level. The model column is calculated from the calculated formaldehyde profile, between the altitude of the station and the model uppermost level (approximately 20 km), and from the *a priori* FTIR profile, above that level. When the

model surface lies higher than the station, the model column is increased by a partial column assuming a constant mixing ratio between the two altitudes, taken equal to the value at the lowermost model level. The monthly averaged formaldehyde columns are calculated by accounting for the temporal sampling of the observations at each site and month. Also, the local time of each observation is taken into account by re-scaling the daily averaged concentration using the formaldehyde diurnal shape factors calculated by the model with a time step of 20 minutes.

4.2 HCHO monthly means and seasonal cycle comparisons

We compare the monthly means of FTIR HCHO total columns at each station with the IMAGES columns calculated for the 2003-2016 period. The time-series of both products are shown in Fig. 7. Since the random uncertainty of the FTIR monthly means is divided by the square root of the number of measurements within each month, the dominant contribution to the FTIR error bars in Fig. 7 is the systematic uncertainty (estimated at 12-31%). Except for very few cases (Mexico City and Paramaribo), the model is in overall good agreement in terms of absolute levels (Fig. 7) and seasonal cycle (Fig. 8) with the FTIR measurements.

For each station the correlation, the bias and the standard deviation (std) of the statistical comparisons between the monthly means, $\text{mean}(\text{IMAGES (smoothed)} - \text{FTIR}) / \text{mean}(\text{FTIR})$, are summarized in Table 4. The median correlation between FTIR and IMAGES for the 21 stations is very high (0.81), with weaker values at the Mexican stations (0.4/0.5) and at Mauna Loa (0.10). The median standard deviation for all comparisons is 25% (ranging from 11% to 41%). This agreement is good considering the FTIR variability (i.e. the std) of HCHO monthly means (median of 35%). The standard deviation of the comparisons can be explained partly by the lower variability of the model monthly means (31%) compared to FTIR, as seen in Fig. 7. In addition, the variability of the model data within a month is also much smaller (median of about 11%; this STD within a month is shown as magenta error bars in Fig. 7) than the FTIR one (mean of about 28%).

The median of IMAGES and FTIR differences is small (-15%) and within the FTIR systematic uncertainty estimated at 12-31%. However, the biases range from -64% to +51%, which requires an investigation of their possible reasons. The main source of systematic uncertainty is the spectroscopic parameters, which have been harmonized in this work, each station using the same line parameters database, and the same spectral micro-windows. Therefore, it is expected that all FTIR stations should provide consistent HCHO total columns within 5-17% (systematic errors due to other sources than spectroscopic ones). To check this, we divide the FTIR stations according to their concentrations levels and latitudes, and use the model for comparisons.

4.2.1 Clean Arctic sites

We distinguish two groups of Arctic sites: Eureka, Ny-Alesund and Thule which are very remote (77-80°N), and two European sites, Kiruna and Sodankyla (67-68°N). As seen in Table 4, the former group shows similar negative biases of the model compared to the data (-20 / -17 / -28%), while the latter group shows positive biases (+32 / +11%). Except at Kiruna, the biases are not constant through the year, the model showing less pronounced seasonal cycles (see also Fig. 8). The model underestimates the summer HCHO levels at the three 77-80°N stations (-26 / -20 / -28%), while the winter levels are in close agreement (+6 / -3%). At the 67-68°N stations, the model is positively biased in winter (+27 / +56%), as well as in summer

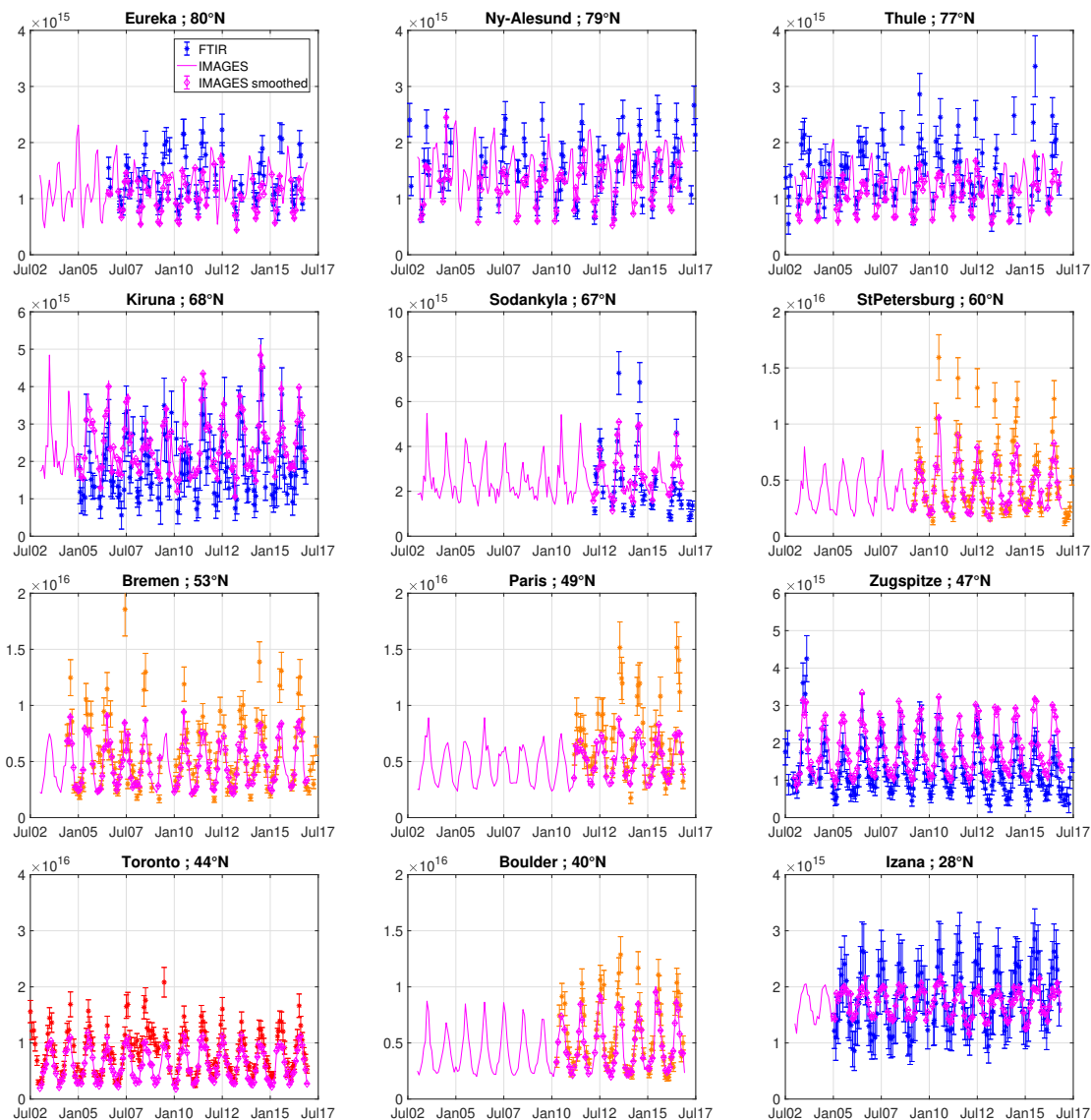


Figure 7. Monthly means of HCHO total columns (molec/cm^2) at each station for FTIR measurements are shown with stars (clean, intermediate, and high levels HCHO sites are shown using blue, orange, and red colors, respectively) and model data (magenta line for “raw” model data; magenta diamonds for model data smoothed by FTIR AK). The FTIR error bars represent the total uncertainties on monthly means which, due to monthly averaging, are mainly the systematic uncertainties. The model error bars represent the standard deviation of the model for each month.

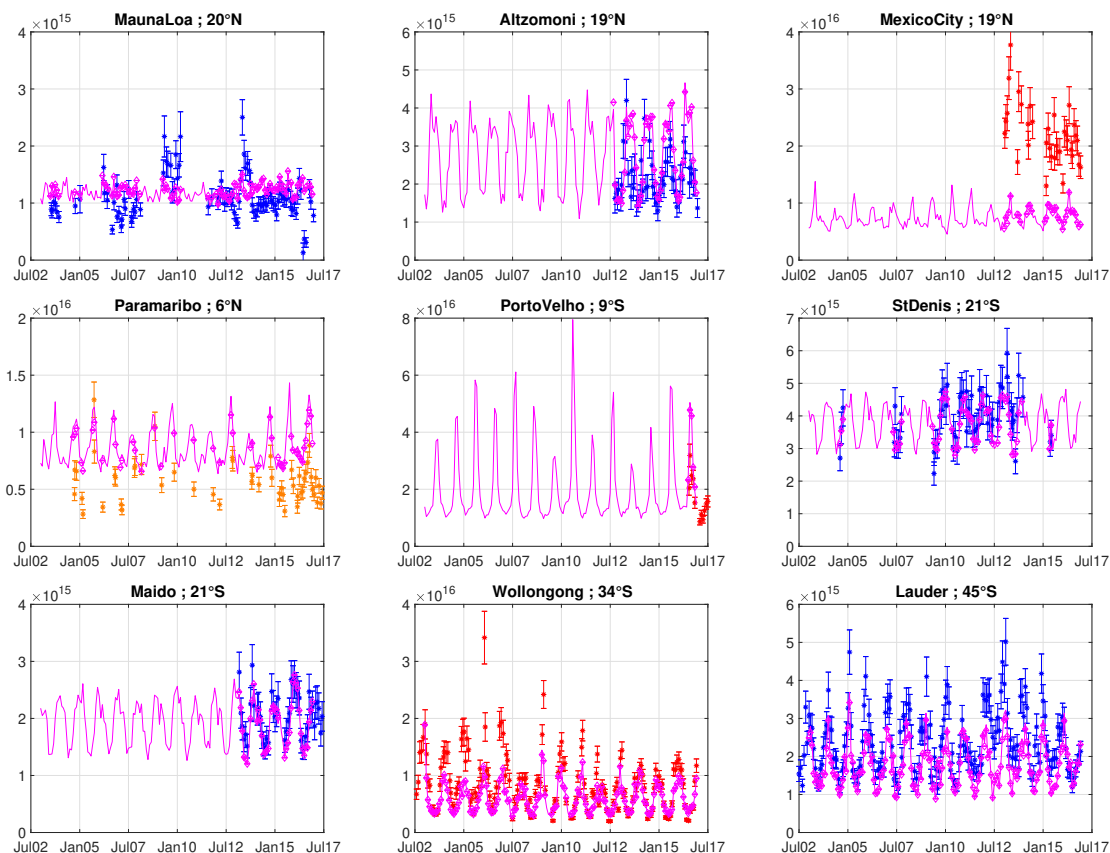


Figure 7. Continued.

at Kiruna (+22%). Note that the Arctic sites do not have measurement during polar night, so the winter months correspond basically to February (Fig. 8).

4.2.2 Mid-latitude cities

Very similar biases (-16 / -15 / -22%) between IMAGES and FTIR are obtained at the three European cities, St. Petersburg (the site is actually at Peterhof, a small coastal city at about 30 km west of St. Petersburg), Bremen, and Paris. As for the Arctic sites, the model underestimates the amplitude of the seasonal cycle (Fig. 8), leading to smaller biases in winter (-14 to -17 %) compared to summer (-19 to -30 %).

The Northern American sites Toronto and Boulder give similar biases (-26%/-17%), especially in summer (-25%/-17%). Toronto is the only mid-latitude urban site where the model shows a higher underestimation of the HCHO levels in winter (-39%).

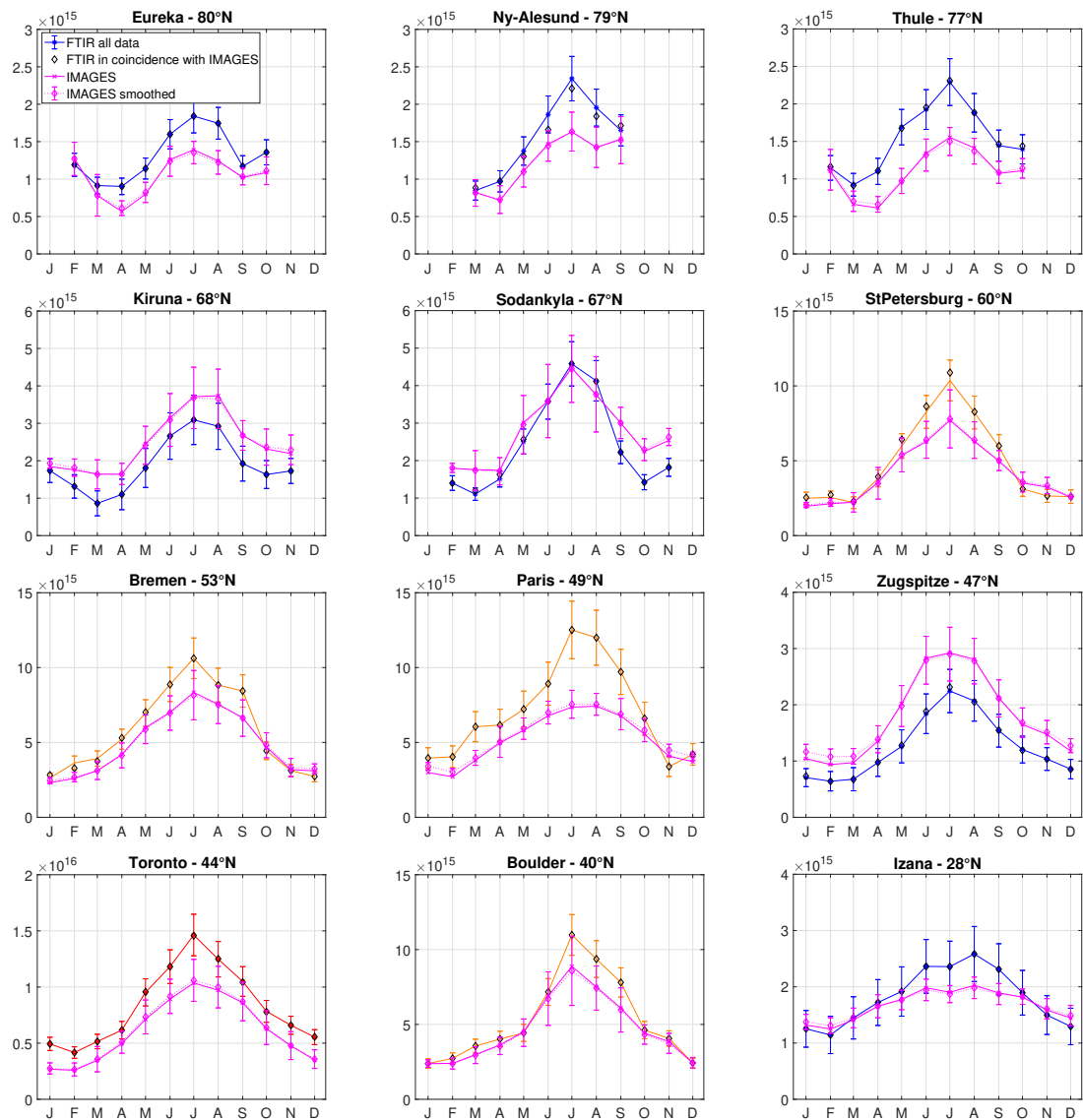


Figure 8. Seasonal cycle of HCHO total columns (molec/cm²) at each station for FTIR measurements (clean, intermediate, and high levels HCHO sites are shown using blue, orange, and red stars, respectively, when only data in coincidence with the model are used; black diamonds correspond to the seasonal cycles when all FTIR data are used) and model data (magenta line for “raw” model data; magenta diamonds for model data smoothed by FTIR AK). The FTIR error bars represent mainly the systematic uncertainties. The model error bars represent the standard deviation of the model for each month. Only the model data in coincidence with FTIR measurements are taken into account in these seasonal cycles.

Table 4. Correlation (Corr), bias \pm standard deviation (STD_{stat}) of the statistical comparisons between the monthly means, $\text{mean(IMGES (smoothed) -FTIR)} / \text{mean(FTIR)}$. Also given: the mean of the standard deviations in the IMGES and FTIR monthly means, i.e. the variability within a month (STD_m), and the standard deviation of the whole FTIR and IMGES monthly mean time-series (STD_{all}). All numbers, except the correlations, are given in %.

Station	Corr	bias \pm STD_{stat}	bias \pm STD_{stat}	bias \pm STD_{stat}	STD_m IMGES/FTIR	STD_{all} IMGES/FTIR
	All	All	JJA	DJF	Within a month	All
Eureka	0.77	-20 \pm 21	-26 \pm 15	+6 \pm 22	10 / 28	28 / 32
Ny-Alesund	0.72	-17 \pm 23	-20 \pm 18	-	9 / 25	30 / 33
Thule	0.74	-28 \pm 24	-28 \pm 23	-3 \pm 18	9 / 31	28 / 35
Kiruna	0.80	+32 \pm 27	+22 \pm 20	+27 \pm 37	10 / 28	31 / 44
Sodankyla	0.85	+11 \pm 33	-4 \pm 27	+56 \pm 35	12 / 34	37 / 60
St. Petersburg	0.94	-16 \pm 29	-25 \pm 20	-14 \pm 23	12 / 32	43 / 60
Bremen	0.87	-15 \pm 30	-19 \pm 27	-16 \pm 40	8 / 20	42 / 56
Paris	0.84	-22 \pm 29	-30 \pm 25	-17 \pm 40	6 / 19	30 / 45
Zugspitze	0.87	+41 \pm 26	+32 \pm 24	+59 \pm 24	10 / 31	37 / 51
Toronto	0.88	-26 \pm 23	-25 \pm 16	-39 \pm 44	15 / 40	46 / 47
Boulder	0.93	-17 \pm 22	-17 \pm 15	-13 \pm 32	12 / 24	47 / 52
Izaña	0.81	-3 \pm 20	-19 \pm 9	+22 \pm 15	8 / 18	14 / 29
Mauna Loa	0.10	+13 \pm 35	+14 \pm 45	+24 \pm 35	9 / 28	9 / 34
Mexico City	0.45	-64 \pm 21	-59 \pm 17	-66 \pm 26	17 / 37	18 / 23
Altzomoni	0.43	+26 \pm 41	+49 \pm 22	-6 \pm 22	16 / 42	35 / 29
Paramaribo	0.67	+51 \pm 25	+59 \pm 15	+85 \pm 17	12 / 35	17 / 33
			DJF	JJA		
Porto Velho	0.87	+41 \pm 35	-	+35 \pm 35	24 / 27	39 / 26
St-Denis	0.71	-7 \pm 13	-3 \pm 12	-9 \pm 15	9 / 27	16 / 18
Maído	0.87	-7 \pm 11	+3 \pm 7	-14 \pm 7	13 / 20	23 / 20
Wollongong	0.83	-26 \pm 37	-29 \pm 34	-3 \pm 35	18 / 50	43 / 59
Lauder	0.77	-25 \pm 22	-24 \pm 17	-26 \pm 26	11 / 31	31 / 35
			“Summer”	“Winter”		
Median	0.81	-15 \pm 25	-19 \pm 19	-5 \pm 26	11 / 28	31 / 35

4.2.3 High-mountain sites

The mountain sites are more difficult to model especially when they are close to cities. They are often very clean sites, but the model cannot reproduce this at the current resolution ($2^\circ \times 2.5^\circ$) when they are surrounded by emission sources in the same

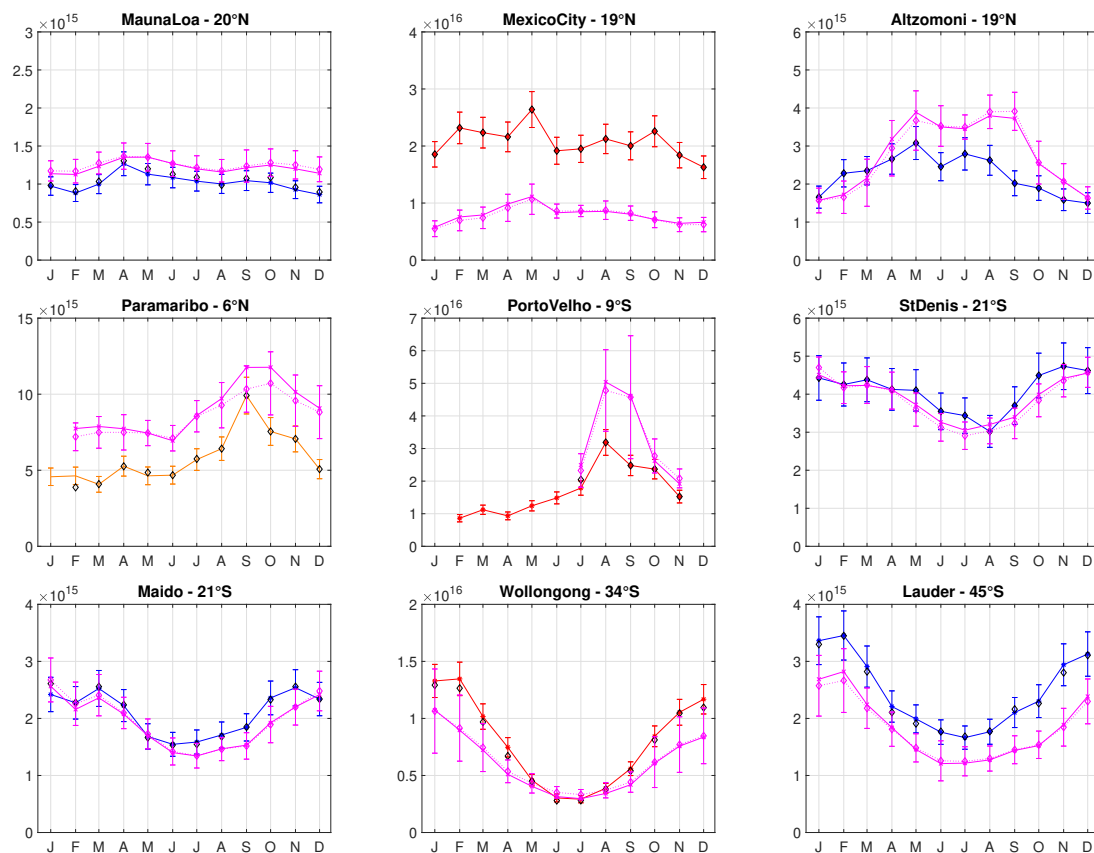


Figure 8. Continued.

pixel. This seems to be the case at Altzomoni, which lies in the same model pixel as Mexico City, leading to an overestimation of 26%, much larger in summer (+49%), and at the European station Zugspitze where the model overestimates the HCHO levels by +41%. Note that in the study of Franco et al. (2015), a negative bias (-13%) was observed between FTIR at Jungfraujoch (47°N, 8°E) and IMAGES, but the retrieval settings used were different than in the present study. Only a change in the spectroscopic database, from HITRAN 2008 to HITRAN 2012, led to lower HCHO columns by 49% at Jungfraujoch (Franco et al., 2015). It is therefore not possible at present to compare the biases obtained at these two close stations.

At the mountain site of Izaña, located at a clean marine area, the model and FTIR are in overall good agreement (-3%), with a negative bias in summer (-19%) and a positive one in winter (+22%), as a result of the weak seasonal amplitude in the model. A moderate positive model bias is calculated at Mauna Loa (+13%), more pronounced in winter (+24%), and a good agreement is seen between the model and FTIR mean seasonal cycle (Fig. 8). The observed variability (34%) is however important at this site, and similar to e.g. the clean Arctic sites (Fig. 7), with values ranging from 0.5 to 2.5×10^{15} molec/cm². This is not reproduced by the model values lying within $1-1.5 \times 10^{15}$ molec/cm². The reasons of the pronounced observed variability are unclear at present.”

4.2.4 Central and South American sites

The model falls short in reproducing the enhanced HCHO levels observed at Mexico City ($ca\ 2 \times 10^{16}$ molec/cm²), mainly due to the coarse model resolution ($2^\circ \times 2.5^\circ$), as suggested by the strong negative bias (-64%), which is almost constant across the year.

5 Comparison at two sites in South America, the coastal site of Paramaribo and the Porto Velho site at the edge of the Amazonian forest, indicates a consistent model overestimation (+51 / +43%). At Porto Velho, this overestimation is more significant during the dry season (August-September, Fig. 8), which corresponds to the maximum of fire intensity in Amazonia. An overestimation of biogenic (isoprene) and biomass burning emissions in Amazonia was already found in IMAGES in the study of Bauwens et al. (2016).

10 4.2.5 Southern Hemisphere 21-45°S sites

The two marine sites at Reunion Island (St-Denis at sea level, and Maïdo at 2.2 km altitude) show a small model bias (-7%) and standard deviation, especially at Maïdo (11%). At these sites, HCHO shows the lowest variability in the monthly means (18-20%), and the model reproduces quite well the seasonal cycle. As shown in Fig. 8, the largest seasonal bias is not found in austral summer (DJF) as seen in the Northern Hemisphere sites, but during September-November months, which correspond
15 to the maximum of the biomass burning period in Southern Africa and Madagascar, close to Reunion Island. The biomass burning source at this location might be underestimated, while it was overestimated in South America.

The Wollongong site shows the same behavior as most of the Northern Hemisphere sites: an overall underestimation of the model (-26%), larger in austral summer (-29%). A first look on the Lauder comparison gives a similar annual bias (-25%), which remains constant over the year, as seen in Table 4 and Fig. 8. However, Fig. 7 shows that during the austral winters
20 (JJA) 2012 to 2015, the FTIR time-series presents unusually high columns. By limiting the comparison to the first years of the period, a better agreement with the model in winter is obtained at Lauder as often observed at other sites.

Since the time-series at St-Denis, Wollongong and Lauder have been published in the past using different retrieval strategies (Vigouroux et al., 2009; Jones et al., 2009; Zeng et al., 2015), we report here the bias observed at these stations between the previous and present data sets. The bias at St-Denis between the previous data set using the strategy in Vigouroux et al. (2009),
25 in which the *a priori* profile and the spectroscopy were different (mostly for interfering species, the HCHO spectroscopic intensity parameters being from the same work of Perrin et al. (2010)), and the mws were smaller than the present work, is only of 1.4% (the new HCHO columns being smaller). Therefore, the comparisons with MAX-DOAS shown in Vigouroux et al. (2009) would still provide a good agreement between the two techniques. Concerning Lauder and Wollongong, where the previous retrieval strategy was from Jones et al. (2009), the present HCHO columns are 49% smaller than the previous data
30 sets. Therefore, the new data set is in much closer agreement with the simulation of four different models that were all of them found 50% lower than the old Lauder and Wollongong data sets (Zeng et al., 2015). From performed sensitivity tests, this high bias between the two strategies is very likely mostly due to the $2869.65\text{-}2870.0\text{ cm}^{-1}$ window used in Jones et al. (2009).

5 Conclusions

Only five NDACC FTIR sites delivered HCHO time-series until now (Paton-Walsh et al., 2005; Jones et al., 2009; Vigouroux et al., 2009; Viatte et al., 2014; Franco et al., 2015), using different retrieval settings. The small number of stations and the bias differences associated with the different retrieval strategies made it difficult to use the FTIR network as a coherent tool for satellite or
5 model validation. In this study, we have designed a harmonized HCHO retrieval strategy to derive total columns at 21 stations, at locations characterized by very different concentrations, from very clean Arctic sites where HCHO is at the limit of detection (a few 10^{13} molec/cm²) to highly polluted sites such Mexico City or Porto Velho, near the Amazonian forest, where columns up to 7×10^{16} molec/cm² have been observed. This network includes well-established NDACC stations, as well as several new sites (Sodankyla, Boulder, Paris, Porto Velho) that aim to be affiliated with NDACC. The FTIR network is also growing, with
10 new sites such as Hefei in China, which will again expand its spatial coverage.

We have presented the retrieval settings that have been optimized for this challenging species, and the FTIR HCHO products have been characterized by their averaging kernels, and their uncertainty budget. The systematic uncertainty of an individual HCHO total column measurement lies between 12 and 27%, with still some differences between the SFIT4 code users (12-15%) and the PROFFIT users (12-27%), which needs to be investigated in the future within the NDACC InfraRed Working
15 Group. The random uncertainty lies between 1 and 11×10^{14} molec/cm², with a median value of 2.9×10^{14} molec/cm², the high maximum value being due to the lower quality of the Bruker Vertex compared to the high resolution ones (Bruker 120/5M or 120/5HR).

In addition to the well-defined seasonal cycles, the diurnal cycles were presented at each site. These observations are crucial to interpret the differences observed between satellites measuring at different local times. For example, the diurnal cycle at
20 Porto Velho which shows insignificant variations suggests that the negative bias observed over Rondônia between OMI (13:30) and GOME-2 (9:30) (De Smedt et al., 2015) is unlikely due to the diurnal cycle. In contrast, the FTIR diurnal cycles in the cities confirm that the positive bias between OMI and GOME-2 over urban areas is likely due, at least partly, to the diurnal cycle.

The monthly mean time-series as well as the seasonal cycles have been compared to the IMAGES model. We did not
25 aim at evaluating the model, but at showing that the FTIR network provides coherent absolute values and seasonal cycles. We observed an overall good agreement with IMAGES, the model usually (but not always) underestimating the HCHO total columns (median bias \pm standard deviation of $-15\% \pm 25\%$), with a more pronounced bias during summer ($-19\% \pm 19\%$). The similar biases obtained at stations under similar conditions (clean Arctic sites, urban sites, marine sites) strengthen our confidence in the harmonization of the HCHO products within the network. When the model showed different behavior for
30 some of the stations, we could explain it by either the too large size of the model pixel ($2.0^\circ \times 2.5^\circ$), especially for high-altitude sites, as in Zugspitze, Alzomoni, Mexico City; or an overestimation of the biogenic and biomass burning sources in South America (Paramaribo, Porto Velho), which was already pointed out in Bauwens et al. (2016). However for a few sites, the behavior of the model remained unexplained (positive biases at Kiruna and Sodankyla, the too low model variability at Mauna Loa).

These HCHO time-series, harmonized and well-characterized, provide an important data set for past and present satellite, and model validation. They are continuously extended by new measurements and will be used in the coming years for the validation of new satellites, such as Sentinel 5P, and Sentinel 4.

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